

# Tutorial 4 – Uncertainty & Bayesian Parameter Estimation

## CSC2541 Neural Net Training Dynamics – Winter 2022

Slides adapted from CSC2541: Scalable and Flexible Models of Uncertainty – Fall 2017

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## Why model uncertainty?

- **Confidence calibration:** know how reliable a prediction is (e.g. so it can ask a human for clarification)
- **Regularization:** prevent your model from overfitting
- **Ensembling:** smooth your predictions by averaging them over multiple possible models
- **Model selection:** decide which of multiple plausible models best describes the data
- **Sparsification:** drop connections, encode them with fewer bits
- **Exploration:** decide which training examples are worth labeling (active learning), optimize an expensive black-box function (Bayesian optimization), estimating rewards from multi-armed bandits (reinforcement learning)
- **Robustness:** make good predictions when the data is either naturally perturbed or explicitly modified by an adversary

# Bayesian Parameter Estimation: A Toy Example

- Motivating example: estimating the parameter of a biased coin
  - You flip a coin 100 times. It lands heads  $N_H = 55$  times and tails  $N_T = 45$  times.
  - What is the probability it will come up heads if we flip again?
- Model: observations  $x_i$  are **independent and identically distributed (i.i.d.)** Bernoulli random variables with parameter  $\theta$ .
- The **likelihood function** is the probability of the observed data (the entire sequence of H's and T's) as a function of  $\theta$ :

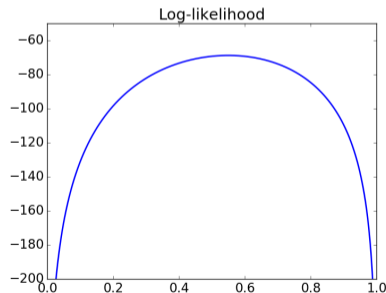
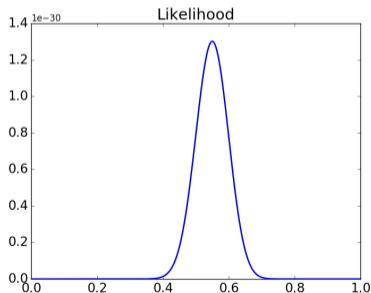
$$\begin{aligned}L(\theta) = p(\mathcal{D}) &= \prod_{i=1}^N \theta^{x_i} (1 - \theta)^{1-x_i} \\ &= \theta^{N_H} (1 - \theta)^{N_T}\end{aligned}$$

- $N_H$  and  $N_T$  are **sufficient statistics**.

# Bayesian Parameter Estimation: A Toy Example

- The likelihood is generally very small, so it's often convenient to work with log-likelihoods.

$$L(\theta) = \theta^{N_H} (1 - \theta)^{N_T} \approx 7.9 \times 10^{-31}$$
$$\ell(\theta) = \log L(\theta) = N_H \log \theta + N_T \log(1 - \theta) \approx -69.31$$



## A Toy Example

- Good values of  $\theta$  should assign high probability to the observed data. This motivates the **maximum likelihood criterion**.
- Solve by setting derivatives to zero:

$$\begin{aligned}\frac{d\ell}{d\theta} &= \frac{d}{d\theta} (N_H \log \theta + N_T \log(1 - \theta)) \\ &= \frac{N_H}{\theta} - \frac{N_T}{1 - \theta}\end{aligned}$$

- Setting this to zero gives the maximum likelihood estimate:

$$\hat{\theta}_{\text{ML}} = \frac{N_H}{N_H + N_T},$$

- Normally there's no analytic solution, and we need to solve an optimization problem (e.g. using gradient descent).

# Bayesian Parameter Estimation: A Toy Example

- Maximum likelihood has a pitfall: if you have too little data, it can overfit.
- E.g., what if you flip the coin twice and get H both times?

$$\theta_{\text{ML}} = \frac{N_H}{N_H + N_T} = \frac{2}{2 + 0} = 1$$

- But even a fair coin has 25% chance of showing this result.
- Because it never observed T, it assigns this outcome probability 0. This problem is known as **data sparsity**.
- If you observe a single T in the test set, the likelihood is  $-\infty$ .

# A Toy Example

- In maximum likelihood, the observations are treated as random variables, but the parameters are not.
- The **Bayesian** approach treats the parameters as random variables as well.
- To define a Bayesian model, we need to specify two distributions:
  - The **prior distribution**  $p(\theta)$ , which encodes our beliefs about the parameters *before* we observe the data
  - The **likelihood**  $p(\mathcal{D} | \theta)$ , same as in maximum likelihood
- When we **update** our beliefs based on the observations, we compute the **posterior distribution** using Bayes' Rule:

$$p(\theta | \mathcal{D}) = \frac{p(\theta)p(\mathcal{D} | \theta)}{\int p(\theta')p(\mathcal{D} | \theta') d\theta'}.$$

- We rarely ever compute the denominator explicitly due to intractability.

# Bayesian Parameter Estimation: A Toy Example

- Let's revisit the coin example. We already know the likelihood:

$$L(\theta) = p(\mathcal{D}) = \theta^{N_H}(1 - \theta)^{N_T}$$

- It remains to specify the prior  $p(\theta)$ .
  - We can choose an **uninformative prior**, which assumes as little as possible. A reasonable choice is the uniform prior.
  - But our experience tells us 0.5 is more likely than 0.99. One particularly useful prior that lets us specify this is the **beta distribution**:

$$p(\theta; a, b) = \frac{\Gamma(a + b)}{\Gamma(a)\Gamma(b)} \theta^{a-1}(1 - \theta)^{b-1}.$$

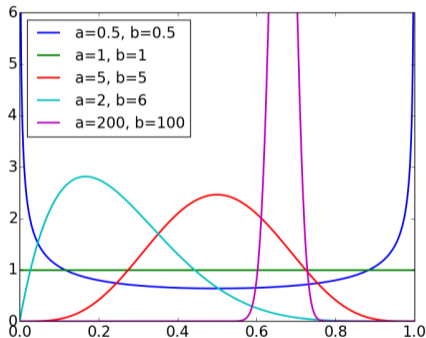
- This notation for proportionality lets us ignore the normalization constant:

$$p(\theta; a, b) \propto \theta^{a-1}(1 - \theta)^{b-1}.$$



# Bayesian Parameter Estimation: A Toy Example

Beta distribution for various values of  $a$ ,  $b$ :



- Some observations:
  - The expectation  $\mathbb{E}[\theta] = a/(a + b)$ .
  - The distribution gets more peaked when  $a$  and  $b$  are large.
  - The uniform distribution is the special case where  $a = b = 1$ .
- The main thing the beta distribution is used for is as a prior for the Bernoulli distribution.

# Bayesian Parameter Estimation: A Toy Example

- Computing the posterior distribution:

$$\begin{aligned} p(\boldsymbol{\theta} | \mathcal{D}) &\propto p(\boldsymbol{\theta})p(\mathcal{D} | \boldsymbol{\theta}) \\ &\propto \left[ \theta^{a-1}(1-\theta)^{b-1} \right] \left[ \theta^{N_H}(1-\theta)^{N_T} \right] \\ &= \theta^{a-1+N_H}(1-\theta)^{b-1+N_T}. \end{aligned}$$

- This is just a beta distribution with parameters  $N_H + a$  and  $N_T + b$ .
- The posterior expectation of  $\theta$  is:

$$\mathbb{E}[\theta | \mathcal{D}] = \frac{N_H + a}{N_H + N_T + a + b}$$

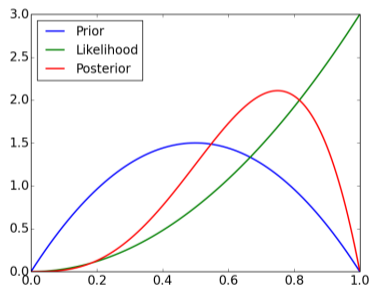
- The parameters  $a$  and  $b$  of the prior can be thought of as **pseudo-counts**.
  - The reason this works is that the prior and likelihood have the same functional form. This phenomenon is known as **conjugacy**, and it's very useful.

# Bayesian Parameter Estimation: A Toy Example

Bayesian inference for the coin flip example:

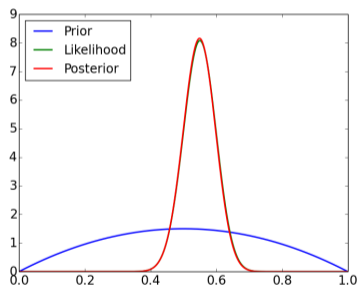
Small data setting

$$N_H = 2, N_T = 0$$



Large data setting

$$N_H = 55, N_T = 45$$



When you have enough observations, the data overwhelm the prior.

# Bayesian Parameter Estimation: A Toy Example

- What do we actually do with the posterior?
- The **posterior predictive distribution** is the distribution over future observables given the past observations. We compute this by marginalizing out the parameter(s):

$$p(\mathcal{D}' | \mathcal{D}) = \int p(\boldsymbol{\theta} | \mathcal{D})p(\mathcal{D}' | \boldsymbol{\theta}) d\boldsymbol{\theta}. \quad (1)$$

- For the coin flip example:

$$\begin{aligned} \theta_{\text{pred}} &= \Pr(\mathbf{x}' = H | \mathcal{D}) \\ &= \int p(\theta | \mathcal{D})\Pr(\mathbf{x}' = H | \theta) d\theta \\ &= \int \text{Beta}(\theta; N_H + a, N_T + b) \cdot \theta d\theta \\ &= \mathbb{E}_{\text{Beta}(\theta; N_H + a, N_T + b)}[\theta] \\ &= \frac{N_H + a}{N_H + N_T + a + b}, \end{aligned} \quad (2)$$

# Bayesian Parameter Estimation: A Toy Example

- **Maximum a-posteriori (MAP) estimation:** find the most likely parameter settings under the posterior
- This converts the Bayesian parameter estimation problem into a maximization problem

$$\begin{aligned}\hat{\theta}_{\text{MAP}} &= \arg \max_{\theta} p(\theta | \mathcal{D}) \\ &= \arg \max_{\theta} p(\theta, \mathcal{D}) \\ &= \arg \max_{\theta} p(\theta) p(\mathcal{D} | \theta) \\ &= \arg \max_{\theta} \log p(\theta) + \log p(\mathcal{D} | \theta)\end{aligned}$$

# Bayesian Parameter Estimation: A Toy Example

- Joint probability in the coin flip example:

$$\begin{aligned}\log p(\theta, \mathcal{D}) &= \log p(\theta) + \log p(\mathcal{D} | \theta) \\ &= \text{const} + (a - 1) \log \theta + (b - 1) \log(1 - \theta) + N_H \log \theta + N_T \log(1 - \theta) \\ &= \text{const} + (N_H + a - 1) \log \theta + (N_T + b - 1) \log(1 - \theta)\end{aligned}$$

- Maximize by finding a critical point

$$0 = \frac{d}{d\theta} \log p(\theta, \mathcal{D}) = \frac{N_H + a - 1}{\theta} - \frac{N_T + b - 1}{1 - \theta}$$

- Solving for  $\theta$ ,

$$\hat{\theta}_{\text{MAP}} = \frac{N_H + a - 1}{N_H + N_T + a + b - 2}$$

# Bayesian Parameter Estimation: A Toy Example

Comparison of estimates in the coin flip example:

	Formula	$N_H = 2, N_T = 0$	$N_H = 55, N_T = 45$
$\hat{\theta}_{\text{ML}}$	$\frac{N_H}{N_H + N_T}$	1	$\frac{55}{100} = 0.55$
$\theta_{\text{pred}}$	$\frac{N_H + a}{N_H + N_T + a + b}$	$\frac{4}{6} \approx 0.67$	$\frac{57}{104} \approx 0.548$
$\hat{\theta}_{\text{MAP}}$	$\frac{N_H + a - 1}{N_H + N_T + a + b - 2}$	$\frac{3}{4} = 0.75$	$\frac{56}{102} \approx 0.549$

How many samples do we need for  $\hat{\theta}_{\text{ML}}$  to be a good estimate of  $\theta$ ? Use **Hoeffding's Inequality** for sampling complexity bound

$$p(|\hat{\theta}_{\text{ML}} - \theta| \geq \varepsilon) \leq 2e^{-2N\varepsilon^2}$$

where  $N = N_H + N_T$ .

## Lessons learned

- Bayesian parameter estimation is more robust to data sparsity.
- Maximum likelihood is about optimization, while Bayesian parameter estimation is about integration.
- The Bayesian solution converges to the maximum likelihood solution as we observe more data.



# Linear Regression as Maximum Likelihood

- We can give linear regression a probabilistic interpretation by assuming a Gaussian noise model:

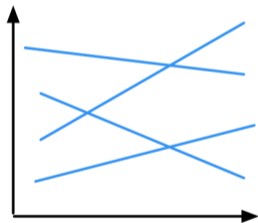
$$t | \mathbf{x} \sim \mathcal{N}(\mathbf{w}^\top \mathbf{x} + b, \sigma^2)$$

- Linear regression is just maximum likelihood under this model:

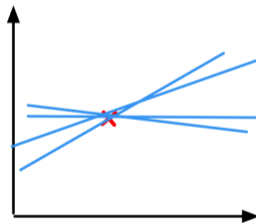
$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N \log p(t^{(i)} | \mathbf{x}^{(i)}; \mathbf{w}, b) &= \frac{1}{N} \sum_{i=1}^N \log \mathcal{N}(t^{(i)}; \mathbf{w}^\top \mathbf{x} + b, \sigma^2) \\ &= \frac{1}{N} \sum_{i=1}^N \log \left[ \frac{1}{\sqrt{2\pi}\sigma} \exp \left( -\frac{(t^{(i)} - \mathbf{w}^\top \mathbf{x} - b)^2}{2\sigma^2} \right) \right] \\ &= \text{const} - \frac{1}{2N\sigma^2} \sum_{i=1}^N (t^{(i)} - \mathbf{w}^\top \mathbf{x} - b)^2 \end{aligned}$$

# Bayesian Linear Regression

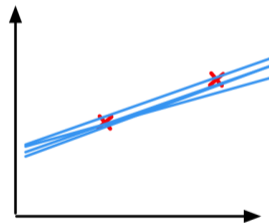
- **Bayesian linear regression** considers various plausible explanations for how the data points were generated.
- It makes predictions using all possible regression weights, weighted by their posterior probability.



no observations



one observation



two observations

# Bayesian Linear Regression

- Leave out the bias for simplicity
- **Prior distribution:** a broad, spherical (multivariate) Gaussian centered at zero:

$$\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \nu^2 \mathbf{I})$$

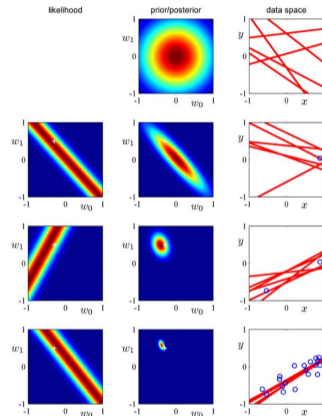
- **Likelihood:** same as in the maximum likelihood formulation:

$$t | \mathbf{x}, \mathbf{w} \sim \mathcal{N}(\mathbf{w}^\top \mathbf{x}, \sigma^2)$$

- **Posterior:**

$$\mathbf{w} | \mathcal{D} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$\boldsymbol{\mu} = \sigma^{-2} \boldsymbol{\Sigma} \mathbf{X}^\top \mathbf{t} \quad \boldsymbol{\Sigma}^{-1} = \nu^{-2} \mathbf{I} + \sigma^{-2} \mathbf{X}^\top \mathbf{X}$$



— Bishop, Pattern Recognition and Machine Learning

Posterior predictive distribution:

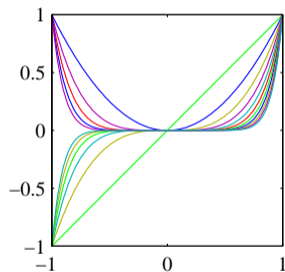
$$\begin{aligned} p(t | \mathbf{x}, \mathcal{D}) &= \int p(t | \mathbf{x}, \mathbf{w}) p(\mathbf{w} | \mathcal{D}) d\mathbf{w} \\ &= \mathcal{N}(t | \boldsymbol{\mu}^\top \mathbf{x}, \sigma_{\text{pred}}^2(\mathbf{x})) \\ \sigma_{\text{pred}}^2(\mathbf{x}) &= \sigma^2 + \mathbf{x}^\top \boldsymbol{\Sigma} \mathbf{x}, \end{aligned}$$

where  $\boldsymbol{\mu}$  and  $\boldsymbol{\Sigma}$  are the posterior mean and covariance of  $\boldsymbol{\Sigma}$ .

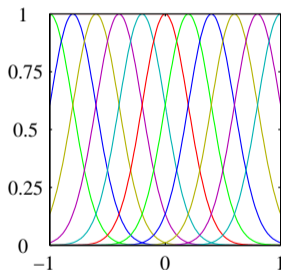
# Bayesian Linear Regression

- We can turn this into nonlinear regression using basis functions.

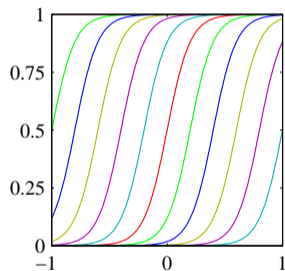
$$\phi_j(x) = x^j$$



$$\phi_j(x) = \exp\left(-\frac{(x - \mu_j)^2}{2s^2}\right)$$

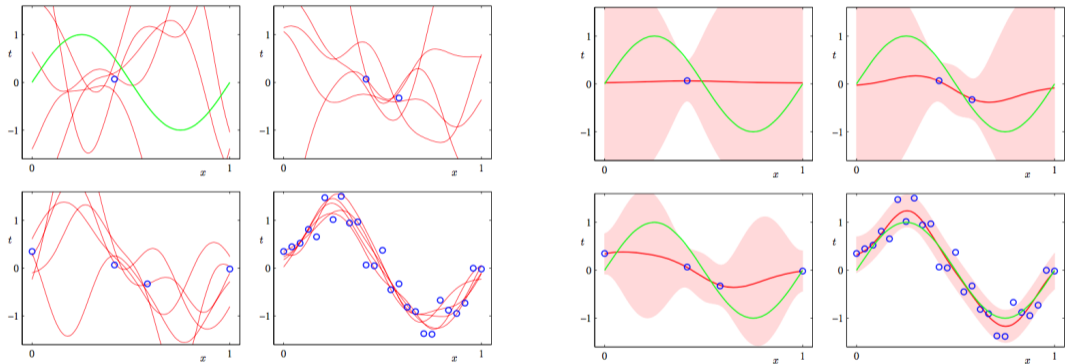


$$\phi_j(x) = \sigma\left(\frac{x - \mu_j}{s}\right)$$



— Bishop, Pattern Recognition and Machine Learning

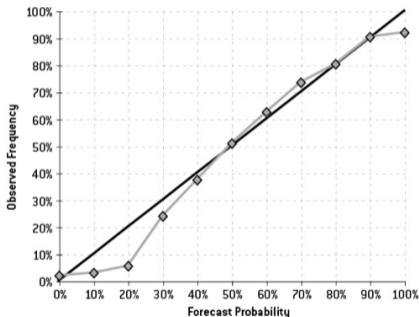
# Bayesian Linear Regression



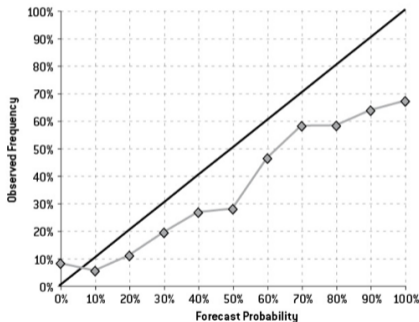
— Bishop, Pattern Recognition and Machine Learning

# Calibration

- **Calibration**: of the times your model predicts something with 90% confidence, is it right 90% of the time?
- Example: calibration of weather forecasts



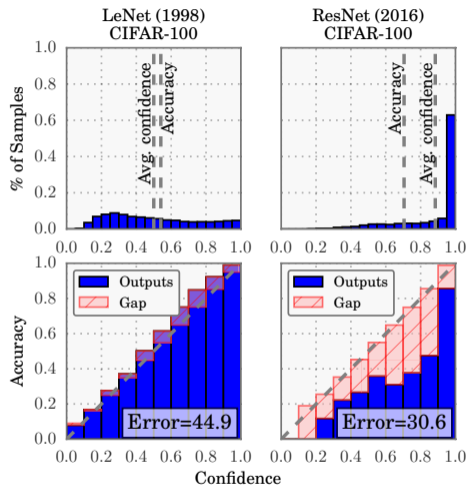
The Weather Channel



Local Weather Station

# Calibration

- Most of our neural nets output probability distributions, e.g. over object categories. Are these calibrated?
- While more accurate, modern neural networks are overconfident in their decisions.



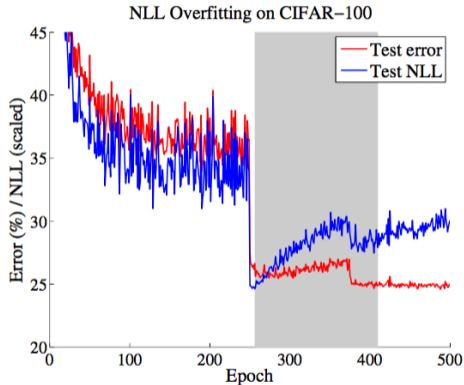
— Guo et al., 2017, On calibration of modern neural networks



- Suppose an algorithm outputs a probability distribution over targets, and gets a loss based on this distribution and the true target.
- A **scoring rule** is a numerical quantization of the calibration of a predictive distribution  $p(y|x)$ . If the underlying true distribution over data points is denoted  $q(x, y)$ , then the expected scoring rule is defined as  $S(p, q) = \mathbb{E}_q[S(p, (x, y))]$  for a scoring function  $S(p, (x, y))$ .
- A **proper scoring rule** is a rule which ensures that  $S(p, q) \leq S(q, q)$  with equality iff  $p(y|x) = q(y|x)$ .
- The canonical example is **negative log-likelihood (NLL)**. If  $k$  is the category label,  $\mathbf{t}$  is the indicator vector for the label, and  $\mathbf{y}$  are the predicted probabilities,

$$L(\mathbf{y}, \mathbf{t}) = -\log y_k = -\mathbf{t}^\top (\log \mathbf{y})$$

- Calibration failures show up in the test NLL scores:



— Guo et al., 2017, On calibration of modern neural networks

# Calibration

- Guo et al. explored 7 different calibration methods, but the one that worked the best was also the simplest: **temperature scaling**.
- A classification network typically predicts  $\sigma(\mathbf{z})$ , where  $\sigma$  is the softmax function

$$\sigma(\mathbf{z})_k = \frac{\exp(z_k)}{\sum_{k'} \exp(z_{k'})}$$

and  $\mathbf{z}$  are called the **logits**.

- They replace this with

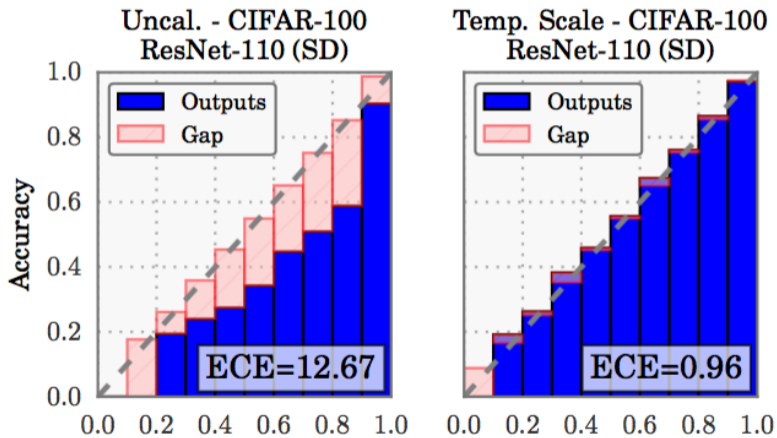
$$\sigma(\mathbf{z}/T),$$

where  $T$  is a scalar called the **temperature**.

- $T$  is tuned to minimize the NLL on a validation set.
- Intuitively, because NLL is a proper scoring rule, the algorithm is incentivized to match the true probabilities as closely as possible.




# Calibration

- Before and after temperature scaling:



# References I

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-  Christopher M Bishop and Nasser M Nasrabadi, *Pattern recognition and machine learning*, vol. 4, Springer, 2006.
-  Chuan Guo, Geoff Pleiss, Yu Sun, and Kilian Q Weinberger, *On calibration of modern neural networks*, International Conference on Machine Learning, PMLR, 2017, pp. 1321–1330.
-  Nate Silver, *The signal and the noise: Why so many predictions fail-but some don't*, Penguin, 2012.