CSC 311: Introduction to Machine Learning Tutorial - Matrix Decomposition & Probabilistic Models

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Based on slides by Haonan Duan

- We can decompose an integer into its prime factors, e.g., $12 = 2 \times 2 \times 3$.
- Similarly, matrices can be decomposed into product of other matrices.
- Examples are Eigendecomposition, SVD, Schur decomposition, LU decomposition, . . .
- Here, we focus on Eigendecomposition and SVD

• An eigenvector of a square matrix A is a nonzero vector v such that multiplication by A only changes the scale of v:

$$Av = \lambda v$$

- The scalar λ is known as the eigenvalue.
- If v is an eigenvector of A, so is any rescaled vector αv .
- αv has the same eigenvalue as v. Thus, we constrain the eigenvector to be of unit length.

Compute eigenvalues - characteristic polynomial

• Eigenvalue equation of matrix A:

$$Av = \lambda v$$
$$\lambda v - Av = 0$$
$$\lambda I - A)v = 0$$

• If nonzero solution for v exists, then it must be the case that:

$$det(\lambda I - A) = 0$$

• Unpacking the determinant as a function of λ , we get a polynomial, called the characteristic polynomial:

$$P_A(\lambda) = det(\lambda I - A) = \lambda^n + c_{n-1}\lambda^{n-1} + \ldots + c_1\lambda + c_0$$

• Compute eigenvalues of $A \to \text{solve } P_A(\lambda) = 0$

Consider the matrix:

$$A = \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix}$$

• What are the eigenvalues and eigenvectors of A?

Solution

We first need to calculate the eigenvalues,

$$det(A - \lambda I) = 0 \implies det \begin{bmatrix} 2 - \lambda & 1 \\ 1 & 2 - \lambda \end{bmatrix} = 0$$
$$\implies (2 - \lambda)^2 - 1 = 0 \implies \lambda_1 = 3, \lambda_2 = 1$$

Then, we solve $(A - \lambda_i I)v_i = 0$ to find eigenvectors:

$$(A - \lambda_1 I)v_1 = 0 \implies \begin{bmatrix} -1 & 1\\ 1 & -1 \end{bmatrix} v_1 = 0$$
$$\implies v_1 = \begin{bmatrix} 1\\ 1 \end{bmatrix} \stackrel{\text{normalize}}{\implies} v_1 = \frac{1}{\sqrt{2}} \begin{bmatrix} 1\\ 1 \end{bmatrix}$$

Similarly,

$$(A - \lambda_2 I)v_2 = 0 \implies v_1 = \frac{1}{\sqrt{2}} \begin{bmatrix} -1\\1 \end{bmatrix}$$

Eigendecomposition

- Spectral Theorem Every symmetric matrix $A \in \mathbb{R}^{n \times n}$ has a set of n orthonormal eigenvectors forming a basis. Furthermore, all eigenvalues are real.
- Therefore, A can be decomposed to the following form

$$A = PDP^{-1}$$

• *P* is an orthogonal matrix of the eigenvectors of *A*, and *D* is a diagonal matrix of eigenvalues.

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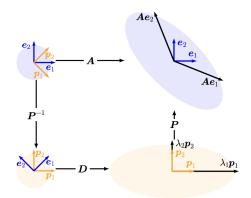
$$A\underbrace{[v_1,\ldots,v_n]}_P = [Av_1,\ldots,Av_n]$$
$$= [\lambda_1v_1,\ldots,\lambda_nv_n]$$
$$= \underbrace{[v_1,\ldots,v_n]}_P \underbrace{\begin{bmatrix}\lambda_1 & \ldots & 0\\ \vdots & \ddots & \vdots\\ 0 & \ldots & \lambda_n\end{bmatrix}}_D$$

Intuitions of Eigendecomposition

- Diagonal matrix allows fast computations of their determinants, powers and inverses.
- Eigendecomposition transforms a matrix into a diagonal form by changing the basis.

 $det(A) = det(PDP^{-1}) = det(P) det(D) det(P)^{-1}$ = det(D) $= \prod_{i=1}^{n} \lambda_i$ $A^{-1} = PD^{-1}P^{-1}$

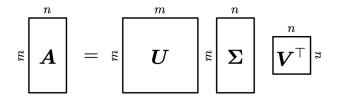
Geometric intuitions of eigendecomposition



- Top-left to bottom-left: P^{-1} performs a basis change.
- $\bullet\,$ Bottom-left to bottom-right: D performs a scaling.
- Bottom-right to top-right: P undoes the basis change.

Singular Value Decomposition (SVD)

- If $A \in \mathbb{R}^{m \times n}$ is not square, eigendecomposition is undefined.
- SVD is a decomposition of the form $A = U\Sigma V^T$.
- SVD is more general than eigendecomposition. **Every** real matrix has a SVD.



- U and V are orthogonal matrices, and Σ is a diagonal matrix (not necessarily square).
- Diagonal entries of Σ are called singular values of A.
- Columns of U are the left singular vectors, and columns of V are the right singular vectors.

- SVD can be interpreted in terms of eigendecomposition.
- Left singular vectors of A are the eigenvectors of AA^T .
- Right singular vectors of A are the eigenvectors of $A^T A$
- Nonzero singular values of A are square roots of eigenvalues of $A^T A$ and AA^T . $A^T A$ and AA^T are positive semi-definite (PSD), thus their eigenvalues are positive.

Informal Proof

Since $B = AA^{\top} \in \mathbb{R}^{m \times m}$ is symmetric, eigendecomposition holds $B = PDP^{-1}$

Now, assume SVD exists, i.e., $A = U\Sigma V^{\top}$. Therefore,

$$B = AA^{\top} = (U\Sigma V^{\top})(V\Sigma^{\top}U^{\top}) = U\Sigma\Sigma^{\top}U^{\top}$$

Matching those two:

 $PDP^{-1} = U\Sigma\Sigma^{\top}U^{\top}$

Therefore, U = P and $\Sigma \equiv D^{\frac{1}{2}}$ or $\sigma_i = \sqrt{d_i}$.

A similar approach on $C = A^{\top}A \in \mathbb{R}^{n \times n}$ leads to V.

Compute SVD of the matrix:

$$A = \begin{bmatrix} 3 & 2 & 2 \\ 2 & 3 & -2 \end{bmatrix}$$

Solution

Here, we calculate U and Σ . First, define $B = AA^{\top}$

$$B = \begin{bmatrix} 3 & 2 & 2 \\ 2 & 3 & -2 \end{bmatrix} \begin{bmatrix} 3 & 2 \\ 2 & 3 \\ 2 & -2 \end{bmatrix} = \begin{bmatrix} 17 & 8 \\ 8 & 17 \end{bmatrix}$$

Then, we can calculate the eigenvalues and eigenvectors (using characteristic polynomial): $\lambda_1 = 25, \lambda_2 = 9$ and $v_1 = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 \\ -1 \end{bmatrix}, v_2 = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 \\ 1 \end{bmatrix}$. Therefore, $B = PDP^{-1}$ where

$$P = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 & 1\\ -1 & 1 \end{bmatrix}, \quad D = \begin{bmatrix} 25 & 0\\ 0 & 9 \end{bmatrix}$$

We had U = P and $\Sigma \equiv D^{\frac{1}{2}}$:

$$\Sigma = \begin{bmatrix} 5 & 0 & 0 \\ 0 & 3 & 0 \end{bmatrix}$$

Find V for exercise.

Intro ML (UofT)

- Given a matrix A, SVD allows us to find its "best" rank-r approximation A_r (r < n).
- Why? store less parameters
- We can write $A = U\Sigma V^{\top}$ as $A = \sum_{i=1}^{n} \sigma_i u_i v_i^{\top}$, where σ_i are sorted from the largest to the smallest.

• The rank-r approximation A_r is defined as:

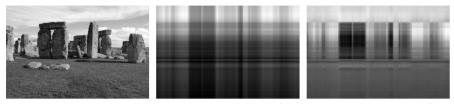
$$A = \sum_{i=1}^{r} \sigma_i u_i v_i^T$$

• A_r is the best approximation of rank r by many norms, such as spectral norm.

$$||A||_2 := \sup_x \frac{||Ax||_2}{||x||_2}$$

• It means that $||A - A_r||_2 \le ||A - B||_2$ for any rank r matrix B.

Rank-r approximation



(a) Original image A. (b) Rank-1 approximation $\widehat{A}(1)$.(c) Rank-2 approximation $\widehat{A}(2)$.



(d) Rank-3 approximation $\widehat{A}(3)$.(e) Rank-4 approximation $\widehat{A}(4)$.(f) Rank-5 approximation $\widehat{A}(5)$.

- Goal: estimate parameters θ from observed data $\{x_1, \cdots, x_N\}$
- Main idea: We should choose parameters that assign high probability to the observed data:

$$\hat{\theta} = \operatorname{argmax} L(\theta; x_1, \cdots, x_N)$$

Three steps for computing MLE

• Write down the likelihood objective:

$$L(\theta; x_1, \cdots, x_N) = \prod_{i=1}^N L(\theta; x_i)$$

2 Transform to log likelihood:

$$l(\theta; x_1, \cdots, x_N) = \sum_{i=1}^N \log L(\theta; x_i)$$

• Compute the critical point:

$$\frac{\partial l}{\partial \theta} = 0$$

X is a discrete random variable with the following probability mass function $(0 \le \theta \le 1$ is an unknown parameter):

Χ	0	1	2	3
$P(\mathbf{X})$	$2\theta/3$	$\theta/3$	$2(1-\theta)/3$	$(1-\theta)/3$

- The following 10 independent observations were taken from \mathbf{X} : $\{3, 0, 2, 1, 3, 2, 1, 0, 2, 1\}$.
- What is the MLE for θ ?

$$L(\theta) = P(X = 3)P(X = 0)P(X = 2)P(X = 1)P(X = 3)$$

× $P(X = 2)P(X = 1)P(X = 0)P(X = 2)P(X = 1)$
= $(\frac{2\theta}{3})^2(\frac{\theta}{3})^3(\frac{2(1-\theta)}{3})^3(\frac{(1-\theta)}{3})^2$

$$l(\theta) = \log L(\theta)$$

= $2(\log \frac{2}{3} + \log \theta) + 3(\log \frac{1}{3} + \log \theta)$
+ $3(\log \frac{2}{3} + \log(1 - \theta)) + 2(\log \frac{2}{3} + \log(1 - \theta))$
= $C + 5(\log \theta + \log(1 - \theta))$

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Step 3: critical points

$$\begin{aligned} \frac{\partial l}{\partial \theta} &= 0\\ \rightarrow 5(\frac{1}{\theta} - \frac{1}{1 - \theta}) &= 0\\ \rightarrow \hat{\theta} &= 0.5 \end{aligned}$$

Suppose that X_1, \dots, X_n form a random sample from a uniform distribution on the interval $(0, \theta)$, where of the parameter $\theta > 0$ but is unknown. Find MLE of θ .

Solution

- Calculate the likelihood:

$$L(X_1, \dots, X_n; \theta) = \prod_i P_{\theta}(X_i) = \prod_i \frac{\mathbb{I}(X_i \in (0, \theta))}{\theta}$$

- Calculate the log-likelihood:

$$l(\theta) = \log \prod_{i} P_{\theta}(X_i) = \sum_{i} \log \frac{\mathbb{I}(X_i \in (0, \theta))}{\theta}$$

If $X_i \notin (0, \theta)$, then log 0 will be undefined. Therefore, $\theta \in [\max_i \{X_i\}, \infty)$

- What value of θ maximizes $l(\theta)$? Given that $\theta \in [\max_i \{X_i\}, \infty)$, we have

$$l(\theta) = \sum_{i} \log \frac{1}{\theta} = -\sum_{i} \log \theta = -n \log \theta$$

Since log is a monotonic function, increasing θ will increase log θ and decrease $l(\theta)$. Therefore, to maximize $l(\theta)$, we choose the smallest feasible value of θ , i.e., $\hat{\theta} = \max_i \{X_i\}.$ Intro ML (UofT)

- Bayesian interprets probability as degrees of beliefs.
- Bayesian treats parameters as random variables.
- Bayesian learning is updating our beliefs (probability distribution) based on observations.

- MLE is the standard frequentist inference method.
- Bayesian and frequentist are the two main approaches in statistical machine learning. Some of their ideological differences can be summarized as:

	Frequentist	Bayesian
Probability is	relative frequency	degree of beliefs
Parameter θ is	unknown constant	random variable

Han Liu and Larry Wasserman, Statistical Machine Learning, 2014

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The Bayesian approach to machine learning

- We define a model that expresses qualitative aspects of our knowledge (eg, forms of *distributions*, independence assumptions). The model will have some unknown *parameters*.
- We specify a *prior* probability distribution for these unknown parameters that expresses our beliefs about which values are more or less likely, before seeing the data.
- We gather data.
- We compute the *posterior* probability distribution for the parameters, given the observed data.
- We use this posterior distribution to draw scientific conclusions and make predictions

Radford M. Neal, Bayesian Methods for Machine Learning, NIPS 2004 tutorial Intro ML (UofT) CSC311-Tut7 29/33 • The posterior distribution is computed by the Bayes' rule:

 $P(parameter|data) = \frac{P(parameter)P(data|parameter)}{P(data)}$

• The denominator is just the required normalizing constant. So as a proportionality, we can write:

posterior \propto prior \times likelihood

- Suppose you have a Beta(4, 4) prior distribution on the probability θ that a coin will yield a 'head' when spun in a specified manner.
- The coin is independently spun ten times, and 'heads' appear fewer than 3 times. You are not told how many heads were seen, only that the number is less than 3.
- Calculate your exact posterior density (up to a proportionality constant) for θ and sketch it.

Solution

• Prior:

$$\theta \sim \text{Beta}(4,4) \implies \text{prior}(\theta) = \frac{1}{B(4,4)} \theta^3 (1-\theta)^3$$

where B(4,4) is a normalization constant.

• Number of heads in n trials follows a binomial distribution $Binomial(k; n, \theta)$. For n = 10, i.e.,

$$P(\text{Heads} = k; n = 10, \theta) = {\binom{10}{k}} \theta^k (1-\theta)^{10-k}$$

Therefore, the Likelihood:

 $L(\text{less than 3 heads out of 10 samples}; \theta)$ $= \sum_{k=0}^{2} P(\text{Heads} = k; n = 10, \theta)$ $= (1 - \theta)^{10} + 10 \cdot \theta (1 - \theta)^9 + 45 \cdot \theta^2 (1 - \theta)^8$

• Posterior:

posterior(θ) $\propto L$ (less than 3 heads out of 10 samples; θ) \cdot prior(θ) $\propto \left[(1-\theta)^{10} + 10 \cdot \theta (1-\theta)^9 + 45 \cdot \theta^2 (1-\theta)^8 \right] \cdot \theta^3 (1-\theta)^3$ $= \theta^3 (1-\theta)^{13} + 10 \cdot \theta^4 (1-\theta)^{12} + 45 \cdot \theta^5 (1-\theta)^{11}$