Today's lecture

Approximate inference in graphical models.

- Forward and Backward KL divergence
- Variational Inference
- Mean Field: Naive and Structured
- Marginal Polytope
- Local Polytope
- Relaxation methods
- Loopy BP
- LP relaxations for MAP inference

Figures from D. Sontag, Murphy's book

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- Almost all approximate inference algorithms in practice are
 - Variational algorithms (e.g., mean-field, loopy belief propagation)
 - Sampling methods (e.g., Gibbs sampling, MCMC)

Variational Methods

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- The KL-divergence is asymmetric

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- What can we do?

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• Let's look at the unnormalized distribution

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- The KL is always non-negative, so we see that J(q) is an upper bound on the negative log likelihood (NLL)

$$J(q) = KL(q||p) - \log Z \ge -\log Z = -\log p(\mathcal{D})$$

Alternative Interpretations

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which is the expected energy minus the entropy.

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• This is the expected NLL plus a penalty term that measures how far apart the two distributions are

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• They differ only when q is minimized over a restricted set of probability distribution $Q = \{q_1, \dots\}$, and $p \neq q$. Why?

Forward or Reverse KL

• Minimizing KL(p||q) or KL(q||p) will give different results
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This is infinite if q(x) = 0 and p(x) > 0. This is zero avoiding, and the forward KL over-estimates the support of p

KL divergence - M projection

$$q^* = \arg\min_{q \in Q} \mathsf{KL}(p||q) = \sum_{\mathbf{x}} p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})}$$

 $p(\mathbf{x})$ is a 2D Gaussian and Q is the set of all Gaussian distributions with diagonal covariance matrices



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KL Divergence (single Gaussian)

• In this example, both the M-projection and I-projection find an approximate $q(\mathbf{x})$ that has the correct mean (i.e., $\mathbb{E}_p(\mathbf{z}) = \mathbb{E}_q(\mathbf{x})$)



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What if $p(\mathbf{x})$ is multimodal?

M projection (Mixture of Gaussians)

$$q^* = \arg\min_{q \in Q} KL(p||q) = \sum_{\mathbf{x}} p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})}$$

 $p(\mathbf{x})$ is a mixture of two 2D Gaussians and Q is the set of all 2D Gaussian distributions (with arbitrary covariance matrices)



M-projection yields a distribution $q(\mathbf{x})$ with the correct mean and covariance.

I projection (Mixture of Gaussians)

$$q^* = \arg \min_{q \in Q} KL(q||p) = \sum_{\mathbf{x}} q(\mathbf{x}) \log \frac{q(\mathbf{x})}{p(\mathbf{x})}$$

$$\prod_{q \in Q} KL(q||p) = \sum_{\mathbf{x}} q(\mathbf{x}) \log \frac{q(\mathbf{x})}{p(\mathbf{x})}$$

The I-projection does not necessarily yield the correct moments

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• We can do the maximization one node at a time, in an iterative fashion

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• Focus on q_i (holding all other terms constant)

$$\begin{split} \mathcal{L}(q_j) &= \sum_{\mathbf{x}} \prod_{i} q_i(\mathbf{x}) \left[\log \tilde{p}(\mathbf{x}) - \sum_{k} \log q_k(\mathbf{x}_k) \right] \\ &= \sum_{\mathbf{x}_j} \sum_{\mathbf{x}_{-j}} q_j(\mathbf{x}_j) \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\log \tilde{p}(\mathbf{x}) - \sum_{k} \log q_k(\mathbf{x}_k) \right] \\ &= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) - \\ &\sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\sum_{k \neq j} \log q_k(\mathbf{x}_k) + \log q_j(\mathbf{x}_j) \right] \\ &= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log f_j(\mathbf{x}_j) - \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log q_j(\mathbf{x}_j) + \text{const} \\ &\log f_j(\mathbf{x}_j) = \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) = \mathbb{E}_{-q_j}[\log \tilde{p}(\mathbf{x})] \end{split}$$

where

• Focus on q_i (holding all other terms constant)

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$$= \sum_{\mathbf{x}_j} \sum_{\mathbf{x}_{-j}} q_j(\mathbf{x}_j) \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\log \tilde{p}(\mathbf{x}) - \sum_k \log q_k(\mathbf{x}_k) \right]$$

$$= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) - \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\sum_{k \neq j} \log q_k(\mathbf{x}_k) + \log q_j(\mathbf{x}_j) \right]$$

$$= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log f_j(\mathbf{x}_j) - \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log q_j(\mathbf{x}_j) + \text{const}$$

$$\log f_j(\mathbf{x}_j) = \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) = \mathbb{E}_{-q_j}[\log \tilde{p}(\mathbf{x})]$$

• So we average out all the variables except x_j , and can rewrite $L(q_j)$ as $L(q_j) = - \mathcal{K}L(q_j || f_j)$

where

• Suppose that we have an arbitrary graphical model

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Mean Field for Variational Inference

 $\max_{q \in Q} \sum_{c \in C} \sum_{\mathbf{x}_c} q(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + H(q(\mathbf{x}))$

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- Although this function is concave and thus in theory should be easy to optimize, we need some compact way of representing q(x)
- Mean field: assume a factored representation of the joint distribution



$$q(\mathbf{x}) = \prod_{i \in V} q_i(x_i)$$

This is called "naive" mean field

Naive Mean Field

• Suppose that *Q* consists of all fully factorized distributions, then we can simplify

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subject to the constraints

$$egin{aligned} q_i(x_i) \geq 0 & orall i \in V, x_i \ & \sum_{x_i} q_i(x_i) = 1 & orall i \in V \end{aligned}$$

• For pairwise MRFs we have

$$\max_{q} \sum_{ij \in E} \sum_{x_i, x_j} \theta_{ij}(x_i, x_j) q_i(x_i) q_j(x_j) - \sum_{i \in V} \sum_{x_i} q_i(x_i) \ln q_i(x_i)$$
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• See Mean field example for the Ising Model, Murphy 21.3.2

Structured mean-field approximations

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Structured mean-field approximations

- Rather than assuming a fully-factored distribution for *q*, we can use a structured approximation, such as a spanning tree
- For example, for a factorial HMM, a good approximation may be a product of chain-structured models (see Murphy 21.4.1)



- Mean field inference approximates posterior as product of marginal distributions
- Allows use of different forms for each variable: useful when inferring statistical parameters of models, or regression weights
- An alternative approximate inference algorithm is **loopy belief propagation**
- Same algorithm shown to do exact inference in trees last class
- In loopy graphs, BP not guaranteed to give correct results, may not converge, but often works well in practice

Algorithm 22.1: Loopy belief propagation for a pairwise MRF

- 1 Input: node potentials $\psi_s(x_s)$, edge potentials $\psi_{st}(x_s, x_t)$;
- 2 Initialize messages $m_{s \to t}(x_t) = 1$ for all edges s t;
- 3 Initialize beliefs $bel_s(x_s) = 1$ for all nodes s;
- 4 repeat
- 5 Send message on each edge

 $m_{s \to t}(x_t) = \sum_{x_s} \left(\psi_s(x_s) \psi_{st}(x_s, x_t) \prod_{u \in \operatorname{nbr}_s \setminus t} m_{u \to s}(x_s) \right);$

- 6 Update belief of each node $bel_s(x_s) \propto \psi_s(x_s) \prod_{t \in nbr_s} m_{t \to s}(x_s);$
- 7 **until** beliefs don't change significantly;
- 8 Return marginal beliefs $bel_s(x_s)$;

Loopy BP for Factor Graph



$$m_{i \to f}(x_i) = \prod_{h \in M(i) \setminus f} m_{h \to i}(x_i)$$
$$m_{f \to i}(x_i) = \sum_{\mathbf{x}_c \setminus x_i} f(\mathbf{x}_c) \prod_{j \in N(f) \setminus i} m_{j \to f}(x_j)$$
$$\mu_i(x_i) \propto \prod_{f \in M(i)} m_{f \to i}(x_i)$$

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- Change from synchronous to asynchronous updates
 - Update sets of nodes at a time, e.g., spanning trees (*tree reparameterization*)

• More theoretical analysis of LBP from variational point of view: (Wainwright & Jordan, 2008)

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- Simplify by considering pairwise UGMs, discrete variables

Variational Inference for Graphical Models

• Suppose that we have an arbitrary graphical model

$$p(\mathbf{x};\theta) = \frac{1}{Z(\theta)} \prod_{c \in C} \phi_c(\mathbf{x}_c) = \exp\left(\sum_{c \in C} \theta_c(\mathbf{x}_c) - \ln Z(\theta)\right)$$

We can compute the KL

$$\begin{aligned} \mathsf{KL}(q||p) &= \sum_{\mathbf{x}} q(\mathbf{x}) \ln \frac{q(\mathbf{x})}{p(\mathbf{x})} \\ &= -\sum_{\mathbf{x}} q(\mathbf{x}) \ln p(\mathbf{x}) - \sum_{\mathbf{x}} q(\mathbf{x}) \ln \frac{1}{q(\mathbf{x})} \\ &= -\sum_{\mathbf{x}} q(\mathbf{x}) \left(\sum_{c \in C} \theta_c(\mathbf{x}_c) - \ln Z(\theta) \right) - H(q(\mathbf{x})) \\ &= -\sum_{c \in C} \sum_{\mathbf{x}} q(\mathbf{x}) \theta_c(\mathbf{x}_c) + \sum_{\mathbf{x}} q(\mathbf{x}) \ln Z(\theta) - H(q(\mathbf{x})) \\ &= -\sum_{c \in C} \mathbb{E}_q[\theta_c(\mathbf{x}_c)] + \ln Z(\theta) - H(q(\mathbf{x})) \end{aligned}$$

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Zemel & Urtasun (UofT)

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- Thus, any approximating distribution q(x) gives a lower bound on the log-partition function
- Recall that KL(p||q) = 0 if an only if p = q. Thus, if we optimize over all distributions we have

$$\ln Z(\theta) = \max_{q} \sum_{c \in C} \mathbb{E}_{q}[\theta_{c}(\mathbf{x}_{c})] + H(q(\mathbf{x}))$$

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• This casts exact inference as a variational optimization problem

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• Assume that $p(\mathbf{x})$ is in the exponential family, and let $\mathbf{f}(\mathbf{x})$ be its sufficient statistic vector

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where M is the marginal polytope, having all valid marginal vectors
• We next push the max inside

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- For a discrete-variable MRF, the sufficient statistic vector f(x) is simply the concatenation of indicator functions for each clique of variables that appear together in a potential function
- For example, if we have a pairwise MRF on binary variables with m = |V| variables and |E| edges, d = 2m + 4|E|

Marginal Polytope for Discrete MRFs



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- The local consistency polytope, M_L is defined by these constraints
- The μ_i and μ_{ij} are called pseudo-marginals

polytope for a tree-structured MRF, and the pseudomarginals are the marginals. marginal polytope, i.e., $M \subseteq M_L$

Mean-field vs relaxation

 $\max_{q \in Q} \sum_{c \in C} \sum_{\mathbf{x}_c} q(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + H(q(\mathbf{x}))$

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$$q(\mathbf{x}) = \prod_{i \in V} q_i(x_i)$$

Naive Mean-Field

• Using the same notation naive mean-field is:

$$(*) \max_{\mu} \sum_{c \in C} \sum_{\mathbf{x}_{c}} \mu_{c}(\mathbf{x}_{c}) \theta_{c}(\mathbf{x}_{c}) + \sum_{i \in V} H(\mu_{i}) \quad \text{subject to}$$
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• Recall the MAP inference task

$$\arg\max_{\mathbf{x}} p(\mathbf{x}), \qquad p(\mathbf{x}) = \frac{1}{Z} \prod_{c \in C} \phi_c(\mathbf{x}_c)$$

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• Since the log is monotonic, let $\theta_c(\mathbf{x}_c) = \log \phi_c(\mathbf{x}_c)$

$$\arg\max_{\mathbf{x}}\sum_{c\in C}\theta_c(\mathbf{x}_c)$$

This is called the max-sum

Application: protein side-chain placement

• Find "minimum energy" configuration of amino acid side-chains along fixed carbon backbone:



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• Find "minimum energy" configuration of amino acid side-chains along fixed carbon backbone:



- Orientations of the side-chains are represented by discretized angles called rotamers
- Rotamer choices for nearby amino acids are energetically coupled (attractive and repulsive forces)

• Given a sentence, predict the dependency tree that relates the words



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- We represent the problem as

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with $\mathbf{x}_{|i} = \{x_{ij}\}_{j \neq i}$ (all outgoing edges)

Application: Semantic Segmentation

• Use Potts to encode that neighboring pixels are likely to have the same discrete label and hence belong to the same segment



$$p(\mathbf{x}, \theta) = \max_{\mathbf{x}} \sum_{i} \theta_i(x_i) + \sum_{i,j} \theta_{i,j}(x_i, x_j)$$





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- What is the dimension of μ , if binary variables?
- Are these two problems equivalent?

Constraints

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- But it might be too slow...

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• Relax integrality constraints, allowing the variables to be between 0 and 1 $\mu_i(x_i) \in [0,1] \ \forall i \in V, x_i \qquad \mu_{ij}(x_i, x_j) \in [0,1] \ \forall ij \in E, x_i, x_j$

$$\begin{split} LP(\theta) &= \max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} (x_i, x_j) \mu_{ij}(x_i, x_j) \\ &\mu_i(x_i) \in [0, 1] \quad \forall i \in V, x_i \\ &\mu_{ij}(x_i, x_j) \in [0, 1] \quad \forall i, j \in E, x_i, x_j \\ &\sum_{x_i} \mu_i(x_i) = 1 \quad \forall i \in V \\ &\mu_i(x_i) = \sum_{x_j} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_i \\ &\mu_j(x_j) = \sum_{x_i} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_j \end{split}$$

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- Zero limit temperature of the variational inference for Marginals

• Introducing Lagrange multipliers and solving we get (see Murphy 22.3.5.4)

$$M_{i
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- We then compute the maximal value of $\mu_s(x_s)$
- What if two solutions that have the same score?

• Tsukuba images from Middlebury stereo database

Left







• Tsukuba images from Middlebury stereo database





• MRF for each pixel, with states the disparity

• Tsukuba images from Middlebury stereo database





- MRF for each pixel, with states the disparity
- Our unary is the matching term

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where pixel $p_i = (x, y)$

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where pixel $p_i = (x, y)$

• The pairwise factor θ_{ij} between neighboring pixels favor smoothness

• If we only use the unary terms. How would you do inference in this case?



• If we only use the unary terms. How would you do inference in this case?



• If full graphical model



[Credit: Coughlan BP Tutorial]

Zemel & Urtasun (UofT)

Subsequent iterations:







Note:

Little change after first few iterations.

Model can be improved to give better results

-- this is just a simple example to illustrate BP.

[Credit: Coughlan BP Tutorial]