CSC 411: Introduction to Machine Learning CSC 411 Lecture 19: Bayesian Linear Regression

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Overview

- We've covered both parametric and nonparametric models for regression and classification.
 - Parametric models summarize the data into a finite-sized model. E.g., linear regression, logistic regression, neural nets, (linear) SVM, Naïve Bayes, GDA
 - Nonparametric models refer back to the data to make predictions.
 E.g., KNN
- The next two lectures are about Bayesian approaches to regression.
 - This lecture: Bayesian linear regression, a parametric model
 - Next lecture: Gaussian processes, a nonparametric model

Overview

- We're going to be Bayesian about the parameters of the model.
 - This is in contrast with naïve Bayes and GDA: in those cases, we used Bayes' rule to infer the class, but used point estimates of the parameters.
 - By inferring a posterior distribution over the *parameters*, the model can know what it doesn't know.
- How can uncertainty in the predictions help us?
 - Smooth out the predictions by averaging over lots of plausible explanations (like ensembles)
 - Assign confidences to predictions
 - Make more robust decisions
 - Guide exploration (focus on areas you're uncertain about)
 - E.g., Bayesian optimization

Recap: Linear Regression

- Given a training set of inputs and targets $\{(\mathbf{x}^{(i)}, t^{(i)})\}_{i=1}^{N}$
- Linear model:

$$y = \mathbf{w}^ op \psi(\mathbf{x})$$

Squared error loss:

$$\mathcal{L}(y,t) = \frac{1}{2}(t-y)^2$$

• L₂ regularization:

$$\mathcal{R}(\mathbf{w}) = \frac{\lambda}{2} \|\mathbf{w}\|^2$$

• Solution 1: solve analytically by setting the gradient to 0

$$\mathbf{w} = (\mathbf{\Psi}^{\top}\mathbf{\Psi} + \lambda \mathbf{I})^{-1}\mathbf{\Psi}^{\top}\mathbf{t}$$

• Solution 2: solve approximately using gradient descent

$$\mathbf{w} \leftarrow (1 - \alpha \lambda) \mathbf{w} - \alpha \mathbf{\Psi}^{\top} (\mathbf{y} - \mathbf{t})$$

Recap: Linear Regression

 We can give linear regression a probabilistic interpretation by assuming a Gaussian noise model:

$$t \mid \mathbf{x} \sim \mathcal{N}(\mathbf{w}^{\top} \boldsymbol{\psi}(\mathbf{x}), \ \sigma^2)$$

• Linear regression is just maximum likelihood under this model:

$$\begin{split} &\frac{1}{N} \sum_{i=1}^{N} \log p(t^{(i)} | \mathbf{x}^{(i)}; \mathbf{w}, b) \\ &= \frac{1}{N} \sum_{i=1}^{N} \log \mathcal{N}(t^{(i)}; \mathbf{w}^{\top} \psi(\mathbf{x}), \sigma^{2}) \\ &= \frac{1}{N} \sum_{i=1}^{N} \log \left[\frac{1}{\sqrt{2\pi}\sigma} \exp \left(-\frac{(t^{(i)} - \mathbf{w}^{\top} \psi(\mathbf{x}))^{2}}{2\sigma^{2}} \right) \right] \\ &= \operatorname{const} - \frac{1}{2N\sigma^{2}} \sum_{i=1}^{N} (t^{(i)} - \mathbf{w}^{\top} \psi(\mathbf{x}))^{2} \end{split}$$

Recap: Linear Regression

- We can view an L_2 regularizer as MAP inference.
- $arg \max_{\mathbf{w}} \log p(\mathbf{w} \mid \mathcal{D}) = arg \max_{\mathbf{w}} [\log p(\mathbf{w}) + \log p(\mathcal{D} \mid \mathbf{w})]$
- We just derived the likelihood term $\log p(\mathcal{D} \mid \mathbf{w})$:

$$\log p(\mathcal{D} \mid \mathbf{w}) = -\frac{1}{2N\sigma^2} \sum_{i=1}^{N} (t^{(i)} - \mathbf{w}^{\top} \mathbf{x} - b)^2 + \text{const}$$

• Assume a Gaussian prior, $\mathbf{w} \sim \mathcal{N}(\mathbf{m}, \mathbf{S})$:

$$\begin{split} \log p(\mathbf{w}) &= \log \mathcal{N}(\mathbf{w}; \mathbf{m}, \mathbf{S}) \\ &= -\frac{1}{2} (\mathbf{w} - \mathbf{m})^{\top} \mathbf{S}^{-1} (\mathbf{w} - \mathbf{m}) + \text{const} \end{split}$$

ullet Set $\mathbf{m}=\mathbf{0}$ and $\mathbf{S}=\eta\mathbf{I}$, then

$$\log p(\mathbf{w}) = -\frac{1}{2\eta} \|\mathbf{w}\|^2 + \text{const.}$$

Recap: Full Bayesian Inference

- Recall: full Bayesian inference makes predictions by averaging over all likely explanations under the posterior distribution.
- Compute posterior using Bayes' Rule:

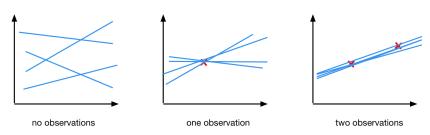
$$p(\mathbf{w} \mid \mathcal{D}) \propto p(\mathbf{w})p(\mathcal{D} \mid \mathbf{w})$$

• Make predictions using the posterior predictive distribution:

$$p(t | \mathbf{x}, \mathcal{D}) = \int p(\mathbf{w} | \mathcal{D}) p(t | \mathbf{x}, \mathbf{w}) d\mathbf{w}$$

• Doing this lets us quantify our uncertainty.

- Bayesian linear regression considers various plausible explanations for how the data were generated.
- It makes predictions using all possible regression weights, weighted by their posterior probability.



- Prior distribution: $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \mathbf{S})$
- Likelihood: $t \mid \mathbf{x}, \mathbf{w} \sim \mathcal{N}(\mathbf{w}^{\top} \psi(\mathbf{x}), \sigma^2)$

Bayesian Linear Regression: Posterior

Deriving the posterior distribution:

$$\begin{split} &\log p(\mathbf{w} \,|\, \mathcal{D}) \\ &= \log p(\mathbf{w}) + \log p(\mathcal{D} \,|\, \mathbf{w}) + \mathrm{const} \\ &= -\, \tfrac{1}{2} \mathbf{w}^\top \mathbf{S}^{-1} \mathbf{w} - \frac{1}{2\sigma^2} \| \mathbf{\Psi} \mathbf{w} - \mathbf{t} \|^2 + \mathrm{const} \\ &= -\, \tfrac{1}{2} \mathbf{w}^\top \mathbf{S}^{-1} \mathbf{w} - \frac{1}{2\sigma^2} \left(\mathbf{w}^\top \mathbf{\Psi}^\top \mathbf{\Psi} \mathbf{w} - 2 \mathbf{t}^\top \mathbf{\Psi} \mathbf{w} + \mathbf{t}^\top \mathbf{t} \right) + \mathrm{const} \\ &= -\, \tfrac{1}{2} (\mathbf{w} - \boldsymbol{\mu})^\top \mathbf{\Sigma}^{-1} (\mathbf{w} - \boldsymbol{\mu}) + \mathrm{const} \end{split}$$

where

$$oldsymbol{\mu} = \sigma^{-2} oldsymbol{\Sigma} oldsymbol{\Psi}^{ op} \mathbf{t}$$
 $oldsymbol{\Sigma}^{-1} = \sigma^{-2} oldsymbol{\Psi}^{ op} oldsymbol{\Psi} + oldsymbol{S}^{-1}$

ullet This is a multivariate Gaussian distribution, i.e. $oldsymbol{w} \mid \mathcal{D} \sim \mathcal{N}(oldsymbol{\mu}, oldsymbol{\Sigma})$

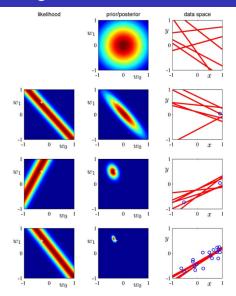
Bayesian Linear Regression: Posterior

Just showed:

$$\begin{split} \mathbf{w} \, | \, \mathcal{D} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \\ \boldsymbol{\mu} &= \sigma^{-2} \boldsymbol{\Sigma} \boldsymbol{\Psi}^{\top} \mathbf{t} \\ \boldsymbol{\Sigma}^{-1} &= \sigma^{-2} \boldsymbol{\Psi}^{\top} \boldsymbol{\Psi} + \mathbf{S}^{-1} \end{split}$$

- Since a Gaussian prior leads to a Gaussian posterior, this means the Gaussian distribution is the conjugate prior for linear regression!
- ullet Compare μ the closed-form solution for linear regression:

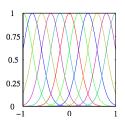
$$\mathbf{w} = (\mathbf{\Psi}^{\top}\mathbf{\Psi} + \lambda \mathbf{I})^{-1}\mathbf{\Psi}^{\top}\mathbf{t}$$



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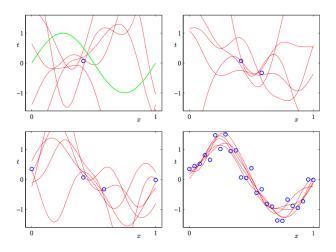
• Example with radial basis function (RBF) features

$$\psi_j(x) = \exp\left(-\frac{(x-\mu_j)^2}{2s^2}\right)$$



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Functions sampled from the posterior:



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Posterior predictive distribution:

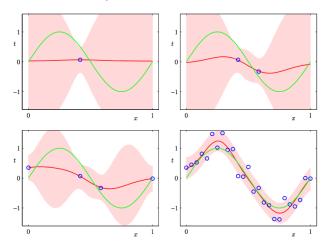
$$p(t \mid \mathbf{x}, \mathcal{D}) = \int \underbrace{p(t \mid \mathbf{x}, \mathbf{w})}_{\mathcal{N}(t; \mathbf{w}^{\top} \psi(\mathbf{x}), \sigma)} \underbrace{p(\mathbf{w} \mid \mathcal{D})}_{\mathcal{N}(\mathbf{w}; \boldsymbol{\mu}, \boldsymbol{\Sigma})} d\mathbf{w}$$

- Another interpretation: $t = \mathbf{w}^{\top} \psi(\mathbf{x}) + \varepsilon$, where $\varepsilon \sim \mathcal{N}(0, \sigma)$ is independent of \mathbf{w} .
- By the linear combination rules for Gaussian random variables, t is a Gaussian distribution with parameters

$$\begin{split} & \mu_{\mathrm{pred}} = \boldsymbol{\mu}^{\top} \boldsymbol{\psi}(\mathbf{x}) \\ & \sigma_{\mathrm{pred}}^2 = \boldsymbol{\psi}(\mathbf{x})^{\top} \boldsymbol{\Sigma} \boldsymbol{\psi}(\mathbf{x}) + \sigma^2 \end{split}$$

• Hence, the posterior predictive distribution is $\mathcal{N}(t; \mu_{\mathrm{pred}}, \sigma_{\mathrm{pred}}^2)$.

Here we visualize confidence intervals based on the posterior predictive mean and variance at each point:



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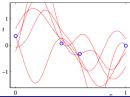
Bayesian Decision Theory

- What do we actually do with the posterior predictive distribution $p(t \mid \mathbf{x}, \mathcal{D})$?
- Often, we want to make a decision. We can formulate this as minimizing the expected loss under the posterior distribution. This is known as **decision theory**.
- Simple example: want to choose a single prediction *y* to minimize the expected squared error loss.

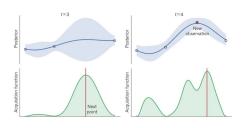
$$\arg\min_{y} \mathbb{E}_{p(t \mid \mathbf{x}, \mathcal{D})}[(y - t)^{2}] = \mathbb{E}_{p(t \mid \mathbf{x}, \mathcal{D})}[t]$$

• Similarly, you can show that under absolute value loss, you should pick the median.

- Black-box optimization: we want to minimize a function, but we only get to query function values (i.e. no gradients!)
 - Each query is expensive, so we want to do as few as possible
 - Canonical example: minimize the validation error of an ML algorithm with respect to its hyperparameters
- Bayesian Optimization: approximate the function with a simpler function called the surrogate function.
- After we've queried a certian number of points, we can condition on these to infer the posterior over the surrogate function using Bayesian linear regression.



- To choose the next point to query, we must define an acquisition function, which tells us how promising a candidate it is.
- Desiderata:
 - \checkmark points we expect to be good $-\mathbb{E}[f(\theta)]$
 - \checkmark points we're uncertain about $Var(f(\theta))$
 - X points we've already tried



Candidate 1: probability of improvement (PI)

$$PI = Pr(f(\theta) < \gamma - \epsilon),$$

where γ is the best value so far, and ϵ is small.

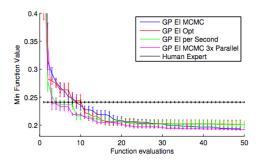
- The problem with Probability of Improvement (PI): it queries points it is highly confident will have a small imporvement
 - Usually these are right next to ones we've already evaluated
- Candidate 2: Expected Improvement (EI)

$$EI = \mathbb{E}[\max(\gamma - f(\theta), 0)]$$

- The idea: if the new value is much better, we win by a lot; if it's much worse, we haven't lost anything.
- There is an explicit formula for this if the posterior predictive distribution is Gaussian.

- Higher-dimensional case is conceptually no different.
 - Maximize the acquisition function using gradient descent
 - Use lots of random restarts, since it is riddled with local maxima
 - BayesOpt can be used to optimize tens of hyperparameters.
- BayesOpt in terms of Bayesian linear regression with basis functions learned by a neural net.
 - In practice, it's typically done with Gaussian processes
 - Bayesian linear regression scales better to large numbers of queries
- One variation: some configurations can be much more expensive than others
 - Another Bayesian regression model to estimate the computational cost, and query the point that maximizes expected improvement per second

- BayesOpt can often beat hand-tuned configurations in a relatively small number of steps.
- Results on optimizing hyperparameters (layer-specific learning rates, weight decay, and a few other parameters) for a CIFAR-10 conv net:



- Each function evaluation takes about an hour
- Human expert = Alex Krizhevsky, the creator of AlexNet

- Occam's Razor: "Entities should not be multiplied beyond necessity."
 - Named after the 14th century British theologian William of Occam
- Huge number of attempts to formalize mathematically
 - See Domingos, 1999, "The role of Occam's Razor in knowledge discovery" for a skeptical overview.

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https://homes.cs.washington.edu/~pedrod/papers/dmkd99.pdf
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Common misinterpretation: your prior should favor simple explanations

- Suppose you have a finite set of models, or **hypotheses** $\{\mathcal{H}_i\}_{i=1}^M$ (e.g. polynomials of different degrees)
- Posterior inference over models (Bayes' Rule):

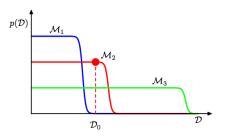
$$p(\mathcal{H}_i \mid \mathcal{D}) \propto \underbrace{p(\mathcal{H}_i)}_{\text{prior}} \underbrace{p(\mathcal{D} \mid \mathcal{H}_i)}_{\text{evidence}}$$

- Which of these terms do you think is more important?
- The evidence is also called **marginal likelihood** since it requires marginalizing out the parameters:

$$p(\mathcal{D} \mid \mathcal{H}_i) = \int p(\mathbf{w} \mid \mathcal{H}_i) \, p(\mathcal{D} \mid \mathbf{w}, \mathcal{H}_i) \, \mathrm{d}\mathbf{w}$$

• If we're comparing a handful of hypotheses, $p(\mathcal{H}_i)$ isn't very important, so we can compare them based on marginal likelihood.

- Suppose M_1 , M_2 , and M_3 denote a linear, quadratic, and cubic model.
- M_3 is capable of explaning more datasets than M_1 .
- But its distribution over \mathcal{D} must integrate to 1, so it must assign lower probability to ones it can explain.



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Approximating the integral:

$$p(\mathcal{D} \mid \mathcal{H}_i) = \int p(\mathcal{D} \mid \mathbf{w}, \mathcal{H}_i) \, p(\mathbf{w} \mid \mathcal{H}_i)$$

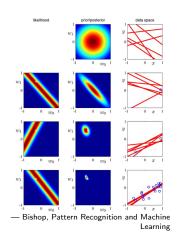
$$\simeq \underbrace{p(\mathcal{D} \mid \mathbf{w}_{\text{MAP}}, \mathcal{H}_i)}_{\text{best-fit likelihood}} \underbrace{p(\mathbf{w}_{\text{MAP}} \mid \mathcal{H}_i) \, \Delta \mathbf{w}}_{\text{Occam factor}}$$

• Multivariate case:

$$p(\mathcal{D} \mid \mathcal{H}_i) \simeq \underbrace{p(\mathcal{D} \mid \mathbf{w}_{\mathsf{MAP}}, \mathcal{H}_i)}_{\mathsf{best-fit likelihood}} \underbrace{p(\mathbf{w}_{\mathsf{MAP}} \mid \mathcal{H}_i) \left| \mathbf{A} \right|^{-1/2}}_{\mathsf{Occam factor}}$$

where
$$\mathbf{A} = \nabla_{\mathbf{w}}^2 \log p(\mathcal{D} \mid \mathbf{w}, \mathcal{H}_i)$$

- The determinant appears because we're taking the volume.
- The more parameters in the model, the higher dimensional the parameter space, and the faster the volume decays.



Analyzing the asymptotic behavior:

$$\begin{split} \mathbf{A} &= \nabla_{\mathbf{w}}^2 \log p(\mathcal{D} \,|\, \mathbf{w}, \mathcal{H}_i) \\ &= \sum_{j=1}^N \underbrace{\nabla_{\mathbf{w}}^2 \log p(y_i \,|\, \mathbf{x}_i, \mathbf{w}, \mathcal{H}_i)}_{\triangleq A_i} \\ &\approx N \, \mathbb{E}[A_i] \\ \log \text{Occam factor} &= \log p(\mathbf{w}_{\text{MAP}} \,|\, \mathcal{H}_i) + \log |\mathbf{A}|^{-1/2} \\ &\approx \log p(\mathbf{w}_{\text{MAP}} \,|\, \mathcal{H}_i) + \log |N \, \mathbb{E}[A_i]|^{-1/2} \\ &= \log p(\mathbf{w}_{\text{MAP}} \,|\, \mathcal{H}_i) - \frac{1}{2} \log |\mathbb{E}[A_i]| - \frac{D \log N}{2} \\ &= \text{const} - \frac{D \log N}{2} \end{split}$$

• Bayesian Information Criterion (BIC): penalize the complexity of your model by $\frac{1}{2}D \log N$.

Summary

$$p(\mathcal{H}_i \mid \mathcal{D}) \propto p(\mathcal{H}_i) p(\mathcal{D} \mid \mathcal{H}_i)$$

$$p(\mathcal{D} \mid \mathcal{H}_i) \simeq p(\mathcal{D} \mid \mathbf{w}_{\text{MAP}}, \mathcal{H}_i) p(\mathbf{w}_{\text{MAP}} \mid \mathcal{H}_i) |\mathbf{A}|^{-1/2}$$

Asymptotically, with lots of data, this behaves like

$$\log p(\mathcal{D} \mid \mathcal{H}_i) = \log p(\mathcal{D} \mid \mathbf{w}_{MAP}, \mathcal{H}_i) - \frac{1}{2}D \log N.$$