

An asymptotically tight bound on the adaptable chromatic number

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Abstract

The *adaptable chromatic number* of a multigraph G , denoted $\chi_a(G)$, is the smallest integer k such that every edge labeling of G from $[k] = \{1, 2, \dots, k\}$ permits a vertex coloring of G from $[k]$ such that no edge $e = uv$ has $c(e) = c(u) = c(v)$. Hell and Zhu proved that for any multigraph G with maximum degree Δ , the adaptable chromatic number is at most $\left\lceil \sqrt{e(2\Delta - 1)} \right\rceil$. We strengthen this to the asymptotically best possible bound of $(1 + o(1))\sqrt{\Delta}$.

1 Introduction

Hell and Zhu introduced the adaptable chromatic number [15]. Given a (not necessarily proper) edge coloring of a multigraph G , we say that a vertex coloring of G is a *proper adaptable coloring* if there is no edge $e = uv$ in G such that $c(u) = c(v) = c(e)$. The *adaptable chromatic number* of G , denoted $\chi_a(G)$, is the minimum number k such that every edge coloring of G using k colors permits an adaptable vertex coloring with the same k colors. Note that $\chi_a(G) \leq \chi(G)$, as every proper vertex coloring is a proper adaptable vertex coloring for any edge coloring. Motivations for this problem arise in the contexts of matrix partitions of graphs and full constraint satisfaction problems [9, 10, 11], job scheduling [16], coloring problems in metric spaces [2], and split colorings [6].

Hell and Zhu prove that $\chi_a(G) \leq \left\lceil \sqrt{e(2\Delta - 1)} \right\rceil$ where Δ is the maximum degree of G . In this paper we improve this bound to the asymptotically best possible bound of $(1 + o(1))\sqrt{\Delta}$.

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We do this by proving a more general result for adaptable list coloring, which is defined naturally. If each vertex v of a multigraph G is given a list of colors, $L(v)$, then a *proper adaptable list coloring* is a proper adaptable coloring where each vertex's color comes from its list, i.e., $c(v) \in L(v)$. The *adaptable choosability number*, denoted $\text{ch}_a(G)$, is the smallest integer k such that every edge-coloring and assignment of lists of size k permits a proper adaptable coloring from these lists.

Theorem 1.1. *For every $\epsilon > 0$, there is a Δ_0 such that if G is an edge-colored multigraph of maximum degree $\Delta > \Delta_0$ with a list of colors $L(v)$ for each vertex $v \in G$ such that $|L(v)| \geq \sqrt{(1 + \epsilon)\Delta}$, then there exists a proper adaptable coloring of G from these lists.*

We can see that this bound is asymptotically best possible. For a multigraph example, take the $(k + 1)$ -clique with k copies of each edge, each colored a different color from $1, 2, \dots, k$. This graph has maximum degree k^2 and is not k -colorable. To see a simple graph example, we refer the reader to [14], where the authors prove that $\chi_a(K_n^2) = n$ if there is a projective plane of order n or $n + 1$. Note that the maximum degree of K_n^2 is $(n - 1)^2$.

An analogous problem for list coloring was introduced by Reed [21], where he proved that for any d , if a graph G has lists $L(v)$ for each vertex v such that (i) for each color $c \in L(v)$, the number of neighbors u of v such that $c \in L(u)$ is at most d , and (ii) $|L(v)| \geq 2ed$, then G can be colored. Haxell [13] improved this result by showing G can be colored if $|L(v)| = 2d$. Reed and Sudakov [22] later showed that for any $\epsilon > 0$, if d is sufficiently large and $|L(v)| = (1 + \epsilon)d$, then G can be colored.

The result of [22] is a special case of Theorem 1.1, since we can convert an instance of the list coloring problem, G , to an instance of the adaptable list coloring problem, G' . Let G' have the same vertex set as G and give each vertex the same list that it had in G . For each edge uv in G and each color $c \in L(u) \cap L(v)$, let G' have a copy of the edge uv and give it the color c . Now G' has maximum degree at most $\Delta \leq dl$, where l is the maximum size of the lists, and can be adaptively colored if and only if G has a coloring from the given lists. Note that $l \geq \sqrt{(1 + \epsilon)\Delta}$ is equivalent to $l \geq (1 + \epsilon)d$.

Our proof of Theorem 1.1 follows the approach taken in [22].

Lower bounds for the adaptable chromatic number for the general multigraph have also been studied; the order of $\chi_a(G)$ is at least $\chi(G)/\sqrt{n \log \chi(G)}$ [12]. A lot of attention has been given to the study of complete graphs [2, 6, 15]; the best upper and lower bounds are both of order \sqrt{n} [6]. Planar graphs were first looked at in [15], and later [8] proved that 4 colors were sufficient without using the Four Color Theorem. The adaptable chromatic number of a graph can be arbitrarily high for any girth [15], and can be as large as the chromatic number [14]. In [15, 16], graphs with $\chi_a(G) = 2$ and graphs with $\text{ch}_a(G) = 2$ were characterized; recognizing graphs with $\chi_a(G) = t$ for $t > 2$ is NP-complete [15].

Remark: Our proof makes use of the Lovász Local Lemma. It is straightforward to apply the results of [19, 20] to obtain a randomized algorithm for

finding an adaptable coloring of a graph which satisfies the conditions of Theorem 1.1; the expected running time is $O(|V|\Delta^2 \log^2 \Delta + |E|)$. We can similarly apply the results of [4] to get a polytime deterministic algorithm.

2 Probabilistic Tools

In this section, we describe some important probabilistic tools that will be used extensively in the proof of Theorem 1.1. One of the most important probabilistic tools is the Lovász Local Lemma, which was first proved in [7] and is presented in many places including [1, 18].

Lovász Local Lemma. *Let A_1, A_2, \dots, A_n be a set of “bad” events in a probability space. Suppose that each event A_i is mutually independent of all the other events A_j but at most d , and that $\Pr(A_i) \leq p$ for all $1 \leq i \leq n$. If $ep(d+1) \leq 1$, then with positive probability, no event A_i holds.*

We use two important concentration bounds. The first is the commonly used Chernoff Bound [5]. Instead of using the original statement we use the (weaker) version found in [18]. Here, $BIN(n, p)$ is the sum of n independent variables, each equal to 1 with probability p and 0 otherwise.

Chernoff Bound. *For any $0 \leq t \leq np$:*

$$\Pr(|BIN(n, p) - np| > t) < 2e^{-\frac{t^2}{3np}}$$

Our next tool is Talagrand’s Inequality [23]. Instead of using the original statement we use the (weaker) version found in [17] (see Equation (2) of Section 3.1).

Talagrand’s Inequality. *Let X be a non-negative random variable determined by the independent trials T_1, T_2, \dots, T_n . Suppose that for every set of possible outcomes of the trials, we have:*

1. *changing the outcome of any one trial can affect X by at most c ; and*
2. *for each $s > 0$, if $X \geq s$, then there is a set of at most rs trials whose outcomes certify that $X \geq s$.*

Then for any $t \geq 0$ where $\frac{t}{2} \geq 20c + \sqrt{rE(X)} + 64c^2r$, we have

$$\Pr(|X - E(X)| > t) \leq 4e^{-\frac{t^2}{32c^2r(E(X)+t)}}.$$

3 A Simpler Bound

Reed’s initial result [21] can be generalized to the following lemma. Note that while this result is weaker than Theorem 1.1 above, this result holds for all Δ rather than just sufficiently large Δ . This theorem is also an essential part of proving Theorem 1.1.

Lemma 3.1. *If G is an edge-colored multigraph of maximum degree $\Delta > 0$ with a list of colors $L(v)$ for each vertex v in G such that $|L(v)| \geq \sqrt{e(2\Delta - 1)}$, then there exists a proper adaptable coloring of G from these lists.*

The proof of this theorem is nearly identical to that of the analogous theorem for regular coloring which can be seen in [21]. In [15, 2], the authors prove the result for regular adaptable coloring; the proof is the same.

Proof. This is a straightforward application of the Lovász Local Lemma. Let G be an edge-colored graph with lists for each vertex satisfying the conditions of the theorem. Truncate the lists to size $l = \left\lceil \sqrt{2e(\Delta - 1)} \right\rceil$. Assign each vertex v in G a color uniformly at random from $\{1, 2, \dots, l\}$.

For each edge $e = uv$, let A_e be the bad event where $c(uv) = c(u) = c(v)$. Let \mathcal{E} be the set of all such events. We must determine the probabilities and dependencies of each event A_e corresponding to the edge $e = uv$. Let $c = c(e)$. If $c \notin L_u$ or $c \notin L_v$, then $\Pr(A_e) = 0$, and this is dependent on nothing. If $c \in L_u$ and $c \in L_v$, then $\Pr(A_e) = \frac{1}{l^2}$. To see the dependency of A_e , define the following:

$$\begin{aligned} D_u &= \{A_f : f = uu' \neq e \text{ is an edge in } G \text{ and } c(f) \in L_u \cap L_{u'}\} \\ D_v &= \{A_f : f = vv' \neq e \text{ is an edge in } G \text{ and } c(f) \in L_v \cap L_{v'}\} \end{aligned}$$

A_e is mutually independent of all events in $\mathcal{E} - |D_u| - |D_v|$. Note that $|D_u| \leq \Delta - 1$ and $|D_v| \leq \Delta - 1$. So A_e is dependent on at most $2\Delta - 2$ events.

If p is the probability of each bad event occurring and d is the number of events each bad event is dependent on, we get

$$ep(d + 1) \leq e \left(\frac{1}{l^2} \right) (2\Delta - 1) \leq e \left(\frac{1}{e(2\Delta - 1)} \right) (2\Delta - 1) = 1 ,$$

so we can apply the Lovász Local Lemma. \square

4 Proof of Main Theorem

Instead of specifying Δ_0 , we, several times throughout the proof, use the fact that Δ is a sufficiently large constant.

To prove Theorem 1.1, we use a random coloring approach like the one used in [22] and many other places (see e.g. [18]). We iteratively color a small portion of the vertices, uncoloring any vertices which create collisions. For each vertex v , we maintain a list $L_i(v)$ of available colors at iteration i , each of which have size L_i which is specified below.

Wasteful Coloring Procedure (ith iteration).

1. For each uncolored vertex v , activate v with probability $\frac{1}{\ln \Delta}$.
2. For each activated vertex v , assign v a color uniformly at random from $L_i(v)$.

3. For each activated vertex v : If v is assigned the color c , for each uncolored neighbor u of v such that $c(uv) = c$, remove c from $L(u)$.
4. Uncolor every vertex v which receives the same color c as a neighbor u where $c(uv) = c$.
5. Truncate the lists so they are all of size L_i . We do this in a very specific way, by removing the colors c for which the parameter $t_i(v, c)$ (defined below) is greatest.

Note that this procedure is wasteful since we remove v 's color from all of v 's neighbors' lists in Step 3 even if we uncolor v in Step 4. We could keep these colors, as they are still usable, but removing them makes the calculations simpler. Below we will prove that with high probability after Step 4 every list has size at least L_i and thus Step 5 is possible.

Now we define several useful parameters for each vertex v and color c . All parameters refer to the value at the end of iteration i .

$$\begin{aligned}
L_i(v) &= v\text{'s list of colors} \\
N_i(v, c) &= \{\text{uncolored neighbors } u \text{ of } v \text{ such that } c = c(uv) \in L_i(u)\} \\
t_i(v, c) &= |N_i(v, c)| \\
x_i(v) &= \# \text{ of uncolored neighbors } u \text{ of } v \text{ such that } c(uv) \in L_i(u) \cap L_i(v) \\
&= \sum_{c \in L_i(v)} t_i(v, c)
\end{aligned}$$

Note that initially we have $x_0(v) \leq \Delta$ and $l_0(v) \geq \sqrt{(1 + \epsilon)\Delta}$ for each vertex v .

Proving the concentration of x_i is made more complicated if the value of a t_i is too large, as will be seen in the proof of lemma 5.3. One trial can affect x_i by as much as the largest t_i , and we need this value to not be too large. The following proposition demonstrates that we will be able to modify the lists so that there are no large t_i 's. Since the t_i 's only decrease with each iteration, we only have to do this once at the beginning. We choose $\Delta^{3/5}$ as our threshold for convenience, but Δ to any power strictly between $\frac{1}{2}$ and 1 would suffice.

Proposition 4.1. *For each vertex v , the number of colors c for which $t_0(v, c) < \Delta^{3/5}$ is at least $\sqrt{(1 + \frac{\epsilon}{2})\Delta}$.*

Proof. Note that for each vertex v there are at most $\Delta^{2/5}$ c 's such that $t_0(v, c) \geq \Delta^{3/5}$. So the number of c 's where $t_0(v, c) < \Delta^{3/5}$ is at least $\sqrt{(1 + \epsilon)\Delta} - \Delta^{2/5} \geq \sqrt{(1 + \frac{\epsilon}{2})\Delta}$ for sufficiently large Δ . \square

Proposition 4.1 permits the following modification to our graph G .

Modification 1. Remove from each list $L(v)$ any color c for which $t_0(v, c) \geq \Delta^{3/5}$. Arbitrarily truncate all the lists to have exactly $\lceil \sqrt{(1 + \frac{\epsilon}{2})\Delta} \rceil$ elements.

Because of this truncation, it will be convenient to do the following reparameterization:

Change of Variables: $\alpha = \frac{\epsilon}{2}$.

The reader may notice that small t_i 's also pose a problem. They will be dealt with in Section 5 by the introduction of $L'_i(v)$.

L_i will be the size of all the lists at the end of iteration i , and X_i will be an upper bound for $x_i(v)$. They are defined by the following recurrences.

$$\begin{aligned} L_0 &= \left\lceil \sqrt{(1+\alpha)\Delta} \right\rceil \\ X_0 &= \Delta \\ L_{i+1} &= \left(1 - \frac{1}{\left(1 + \frac{3\alpha}{4}\right) \ln \Delta} \right) L_i \\ X_{i+1} &= \left(1 - \frac{1}{\left(1 + \frac{\alpha}{2}\right) \ln \Delta} \right)^2 X_i \end{aligned}$$

We now introduce a new parameter.

$l_i(v)$ = size of v 's list right before step 5 in iteration i

We would like to maintain the following property.

Property $P(i)$. For each uncolored vertex v of the graph G , the following is true:

$$l_i(v) \geq L_i \text{ and } x_i(v) \leq X_i$$

Lemma 4.2. *With positive probability, at the end of each iteration i for $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, Property $P(i)$ holds.*

We now prove the main theorem using this lemma.

Proof of Theorem 1.1. Carry out the wasteful coloring procedure up through iteration $i' = \lceil \frac{10}{\alpha} \ln \Delta \rceil$. Property $P(i)$ ensures that we can carry out Step 5. At the end of iteration i' , each vertex v has $|L_{i'}(v)| = L_{i'}$ and $x_{i'}(v) \leq X_{i'}$. We will show that the ratio of $x_{i'}(v)$ to $|L_{i'}(v)|^2$ is at most $\frac{1}{2e}$; thus we can complete the coloring using Lemma 3.1.

$$\begin{aligned} \frac{x_{i'}(v)}{|L_{i'}(v)|^2} &\leq \frac{X_{i'}}{L_{i'}^2} = \frac{\left(1 - \frac{1}{\left(1 + \frac{\alpha}{2}\right) \ln \Delta} \right)^{2i'} X_0}{\left(1 - \frac{1}{\left(1 + \frac{3\alpha}{4}\right) \ln \Delta} \right)^{2i'} L_0^2} \\ &= \left(\frac{1 - \frac{1}{\left(1 + \frac{\alpha}{2}\right) \ln \Delta}}{1 - \frac{1}{\left(1 + \frac{3\alpha}{4}\right) \ln \Delta}} \right)^{2i'} \cdot \frac{\Delta}{(1+\alpha)\Delta} \\ &< \left(\frac{\left(1 + \frac{\alpha}{2}\right) \left(1 + \frac{3\alpha}{4}\right) \ln \Delta - \left(1 + \frac{3\alpha}{4}\right)}{\left(1 + \frac{\alpha}{2}\right) \left(1 + \frac{3\alpha}{4}\right) \ln \Delta - \left(1 + \frac{\alpha}{2}\right)} \right)^{2i'} \end{aligned}$$

$$\begin{aligned}
&= \left(1 - \frac{\frac{\alpha}{4}}{\left(1 + \frac{\alpha}{2}\right) \left(1 + \frac{3\alpha}{4}\right) \ln \Delta - \left(1 + \frac{\alpha}{2}\right)} \right)^{2i'} \\
&\leq \left(1 - \frac{\alpha}{10 \ln \Delta} \right)^{2 \cdot \frac{10}{\alpha} \ln \Delta} \\
&\leq e^{-2} < \frac{1}{2e}
\end{aligned}$$

□

To prove Lemma 4.2, we need to show that with high probability x_i decreases quickly enough while l_i does not decrease too quickly. The concentration lemmas, which we prove in Section 5, are as follows.

Lemma 4.3. *For each vertex v and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then*

$$\Pr(l_{i+1}(v) < L_{i+1}) < \Delta^{-\ln \Delta}$$

Lemma 4.4. *For each vertex v and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then*

$$\Pr(x_{i+1}(v) > X_{i+1}) < \Delta^{-\ln \Delta}$$

The proof of Lemma 4.2 follows easily from Lemmas 4.3 and 4.4.

Proof of Lemma 4.2. We prove this by induction on i . The base case $i = 0$ is trivially true. Let A_v be the bad event that $l_{i+1}(v) < L_{i+1}$ and let B_v be the bad event that $x_{i+1}(v) > X_{i+1}$. By Lemmas 4.3 and 4.4 the probability of each of these bad events is at most $\Delta^{-\ln \Delta}$. A_v and B_v are easily seen to be mutually independent from all A_u and B_u where u is a distance greater than 4 from v . Since the maximum degree of G is at most $\sqrt{(1+\epsilon)\Delta} \cdot \Delta < \Delta^2$, A_v and B_u are mutually independent from all but at most Δ^8 events. So, as

$$e \cdot \Delta^{-\ln \Delta} \cdot \Delta^8 < 1 \quad \Delta \text{ for sufficiently large,}$$

we can apply the Lovász Local Lemma. □

5 Proofs of the Concentration Lemmas

The bulk of the proof is in the concentration lemmas.

5.1 Helpful Facts

We start with three facts that are used several times throughout the proofs of our concentration lemmas.

Since $L_0 = \sqrt{(1+\alpha)\Delta}$ and

$$L_i = \left(1 - \frac{1}{\left(1 + \frac{3\alpha}{4}\right) \ln \Delta} \right)^i \cdot L_0$$

$$\geq \left(1 - \frac{1}{(1 + \frac{3\alpha}{4}) \ln \Delta}\right)^{\frac{10}{\alpha} \ln \Delta + 1} \cdot \sqrt{\Delta} = \Theta(\sqrt{\Delta}) ,$$

we get the following.

Fact 1. $L_i = \Theta(\sqrt{\Delta})$ for all i such that $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$.

Similarly, as $X_0 = \Delta$ and

$$\begin{aligned} X_i &= \left(1 - \frac{1}{(1 + \frac{\alpha}{2}) \ln \Delta}\right)^{2i} \cdot X_0 \\ &\geq \left(1 - \frac{1}{(1 + \frac{\alpha}{2}) \ln \Delta}\right)^{\frac{20}{\alpha} \ln \Delta + 1} \cdot \Delta = \Theta(\Delta) , \end{aligned}$$

we get a similar fact for X_i .

Fact 2. $X_i = \Theta(\Delta)$ for all i such that $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$.

Finally, the ratio of X_i to L_i^2 comes up several times and we use the following.

$$\frac{X_i}{L_i^2} = \frac{\left(1 - \frac{1}{(1 + \frac{\alpha}{2}) \ln \Delta}\right)^{2i'} X_0}{\left(1 - \frac{1}{(1 + \frac{3\alpha}{4}) \ln \Delta}\right)^{2i'} L_0^2} \leq \frac{X_0}{L_0^2} = \frac{1}{1 + \alpha}$$

This yields our final useful fact.

Fact 3. $\frac{X_i}{L_i^2} \leq \frac{1}{1 + \alpha}$ for all i such that $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$.

5.2 Proof of Lemma 4.3

We start with bounding the expected value of $l_{i+1}(v)$.

Lemma 5.1. For each vertex v and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then,

$$E(l_{i+1}(v)) \geq \left(1 - \frac{1}{(1 + \alpha) \ln \Delta}\right) L_i.$$

Proof.

$$\begin{aligned} E(l_{i+1}(v)) &= \sum_{c \in L_i(v)} \Pr(\text{no neighbor } u \text{ of } v \text{ is assigned color } c) \\ &= \sum_{c \in L_i(v)} \prod_{u \in N_i(v, c)} (1 - \Pr(u \text{ is activated and assigned } c)) \\ &= \sum_{c \in L_i(v)} \prod_{u \in N_i(v, c)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right) \end{aligned}$$

$$= \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_i(v,c)}$$

In the following claim, we show that we can use $\frac{x_i(v)}{L_i}$ in place of $t_i(v, c)$. Note that this quantity need not be an integer.

Claim 5.1.1. Given $x_i(v) = \sum_{c \in L_i(v)} t_i(v, c)$, the above bound is minimized over the rationals when all the t_i 's are equal.

Proof of claim. Let c_1 and c_2 be two colors in $L_i(v)$ such that the following is true.

$$t_i(v, c_1) > \frac{x_i(v)}{L_i} \text{ and } t_i(v, c_2) < \frac{x_i(v)}{L_i}$$

Let $t_1 = t_i(v, c_1)$, $t_2 = t_i(v, c_2)$ and $\gamma = 1 - \frac{1}{\ln \Delta \cdot L_i}$. Define t'_i as follows, where $\beta = \min(t_1 - \frac{x_i(v)}{L_i}, \frac{x_i(v)}{L_i} - t_2)$.

$$t'_i(v, c) = \begin{cases} t_1 - \beta, & \text{if } c = c_1 \\ t_2 + \beta, & \text{if } c = c_2 \\ t_i(v, c), & \text{o.w.} \end{cases}$$

Let $D = \sum_{c \in L_i(v)} \gamma^{t_i(v,c)} - \sum_{c \in L_i(v)} \gamma^{t'_i(v,c)}$. It suffices to show that $D \geq 0$.

$$\begin{aligned} D &= \gamma^{t_1} + \gamma^{t_2} - \gamma^{t_1 - \beta} - \gamma^{t_2 + \beta} \\ &= -\gamma^{t_1 - \beta} (1 - \gamma^\beta) + \gamma^{t_2} (1 - \gamma^\beta) \\ &= (1 - \gamma^\beta) (\gamma^{t_2} - \gamma^{t_1 - \beta}) \end{aligned}$$

Since $t_1 - \beta \geq t_2$ and $0 < \gamma < 1$, $D \geq 0$. □

Using Claim 5.1.1, we can continue our computation of the bound on the expected value of $l_{i+1}(v)$.

$$\begin{aligned} E(l_{i+1}(v)) &\geq \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{\frac{x_i(v)}{L_i}} \tag{1} \\ &\geq \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{\frac{X_i}{L_i}} \text{ [by Property } P(i)] \\ &= L_i \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{\frac{X_i}{L_i}} \\ &\geq L_i \left(1 - \frac{X_i}{\ln \Delta \cdot L_i^2}\right) \text{ for } \Delta \text{ sufficiently large [by Facts 1 \& 2]} \\ &\geq L_i \left(1 - \frac{1}{(1 + \alpha) \ln \Delta}\right) \text{ [by Fact 3]} \end{aligned}$$

And this completes our proof of Lemma 5.1. □

We now show that the lists stay large enough by proving Lemma 4.3.

Proof of Lemma 4.3. Instead of showing the concentration of l_{i+1} directly, it is more convenient to work with $R = L_i - l_{i+1}(v)$, the number of colors which leave v 's list. By the linearity of expectation and Lemma 5.1, we have $E(R) = L_i - E(l_{i+1}(v)) \leq L_i \cdot \left(\frac{1}{(1+\alpha) \ln \Delta} \right) < \frac{L_i}{\ln \Delta}$.

Changing the assignment of any vertex can affect R by at most 1, as changing the assignment of one of v 's neighbors can at most change R by 1, and changing any other vertex has no effect. Furthermore, for any s , to certify $R \geq s$, we can use s neighbors of v which were assigned different colors. So Talagrand's Inequality with $c = r = 1$ and $t = \frac{L_i}{\ln^2 \Delta}$ yields:

$$\begin{aligned} \Pr \left(|R - E(R)| > \frac{L_i}{\ln^2 \Delta} \right) &< 4e^{-\frac{L_i^2 / \ln^4 \Delta}{32(L_i / \ln \Delta + L_i / \ln^2 \Delta)}} \\ &< 4e^{-\frac{L_i}{64 \ln^3 \Delta}} \\ &< e^{-\ln^2 \Delta} = \Delta^{-\ln \Delta} \text{ [by Fact 1]} \end{aligned}$$

Since $R = L_i - l_{i+1}(v)$ and $E(R) = L_i - E(l_{i+1}(v))$, Lemma 5.1 yields:

$$\begin{aligned} &\Pr \left(l_{i+1}(v) < \left(1 - \frac{1}{(1 + \frac{3\alpha}{4}) \ln \Delta} \right) L_i \right) \\ &\leq \Pr \left(l_{i+1}(v) < \left(1 - \frac{1}{(1 + \alpha) \ln \Delta} \right) L_i - \frac{L_i}{\ln^2 \Delta} \right) \\ &\leq \Pr \left(l_{i+1}(v) < E(l_{i+1}(v)) - \frac{L_i}{\ln^2 \Delta} \right) \\ &\leq \Pr \left(|l_{i+1}(v) - E(l_{i+1}(v))| > \frac{L_i}{\ln^2 \Delta} \right) \\ &= \Pr \left(|R - E(R)| > \frac{L_i}{\ln^2 \Delta} \right) \\ &< \Delta^{-\ln \Delta} \end{aligned}$$

□

5.3 Proof of Lemma 4.4

The concentration of x_i is more involved. It is convenient to split the analysis into steps by introducing the following parameters.

$$\begin{aligned} L_i^*(v) &= v\text{'s list of colors right before Step 5} \\ y_i(v) &= \sum_{c \in L_i^*(v)} t_{i-1}(v, c) \end{aligned}$$

The parameter y_i is like x_i , except we are just reducing the size of the list, and not decreasing the t_i 's, and we are looking at the value before the truncation

step. This allows us to analyze the rate of decrease of the lists and the t_i 's separately. To show that x_i decreases enough, we show that both y_i and t_i decrease enough. We will then factor in the effect of the truncation step. We start our analysis by computing a bound on the expected value of y_{i+1} .

Lemma 5.2. *For each vertex v and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then*

$$E(y_{i+1}(v)) \leq \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^3 \Delta} \right) X_i$$

Proof. Bound the expected value of $y_{i+1}(v)$ with the following calculations. Remember that a color c is removed from v 's list if a neighbor is colored c .

$$\begin{aligned} E(y_{i+1}(v)) &= \sum_{c \in L_i(v)} \Pr(c \in L_{i+1}(v)) \cdot t_i(v, c) \\ &\leq \sum_{c \in L_i(v)} \left[\prod_{u \in N_i(v, c)} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right) \right] \cdot t_i(v, c) \\ &= \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_i(v, c)} \cdot t_i(v, c) \end{aligned}$$

We show that this bound on $E(y_{i+1}(v))$ is maximized over integers when

$$t_i(v, c) = \left\lceil \frac{x_i(v)}{L_i} \right\rceil \text{ or } \left\lfloor \frac{x_i(v)}{L_i} \right\rfloor$$

for all vertices v . Suppose, to the contrary, that there are two colors $c_1, c_2 \in L_i(v)$ such that

$$t_i(v, c_1) > t_i(v, c_2) + 1.$$

Let $t_1 = t_i(v, c_1)$ and $t_2 = t_i(v, c_2)$. Define t'_i as follows.

$$t'_i(v, c) = \begin{cases} t_1 - 1, & \text{if } c = c_1 \\ t_2 + 1, & \text{if } c = c_2 \\ t_i(v, c), & \text{o.w.} \end{cases}$$

Let

$$D = \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t'_i(v, c)} t'_i(v, c) - \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_i(v, c)} t_i(v, c).$$

It suffices to show that $D \geq 0$.

$$\begin{aligned} D &= \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_1-1} (t_1 - 1) + \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_2+1} (t_2 + 1) \\ &\quad - \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_1} t_1 - \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_2} t_2 \end{aligned}$$

$$\begin{aligned}
&= \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_1-1} \left[t_1 - 1 - \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right) t_1 \right] \\
&\quad + \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left[\left(1 - \frac{1}{\ln \Delta \cdot L_i}\right) (t_2 + 1) - t_2 \right] \\
&= \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_1-1} \left(\frac{t_1}{\ln \Delta \cdot L_i} - 1 \right) + \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left(1 - \frac{t_2+1}{\ln \Delta \cdot L_i}\right)
\end{aligned}$$

Property $P(i)$ and Fact 3 yields that the second term here is positive by the following argument.

$$\frac{t_2 + 1}{\ln \Delta \cdot L_i} \leq \frac{\frac{x_i(v)}{L_i}}{\ln \Delta \cdot L_i} \leq \frac{X_i}{L_i^2 \ln \Delta} \leq \frac{1}{(1 + \alpha) \ln \Delta} < 1$$

for sufficiently large Δ .

Thus if $D < 0$ then the first term is negative. So assume $\frac{t_1}{\ln \Delta \cdot L_i} - 1 < 0$. Note that by our definitions of t_1 and t_2 we have that $t_1 - 1 \geq t_2$. So we get the following.

$$\begin{aligned}
D &\geq \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left(\frac{t_1}{\ln \Delta \cdot L_i} - 1 \right) + \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left(1 - \frac{t_2+1}{\ln \Delta \cdot L_i}\right) \\
&= \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left[\left(\frac{t_1}{\ln \Delta \cdot L_i} - 1 \right) + \left(1 - \frac{t_2+1}{\ln \Delta \cdot L_i}\right) \right] \\
&= \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{t_2} \left(\frac{t_1 - 1 - t_2}{\ln \Delta \cdot L_i} \right) > 0
\end{aligned}$$

So the bound on $E(y_{i+1}(v))$ is maximized over the integers when all t_i 's are within one of each other; i.e., $t_i(v, c) = \lceil x_i(v)/L_i \rceil$ or $\lfloor x_i(v)/L_i \rfloor$ for all vertices v . Using Equation 1 from the proof of Lemma 5.1 as a bound on $E(l_{i+1}(v))$, we continue computing the bound on $E(y_{i+1}(v))$ as follows.

$$\begin{aligned}
E(y_{i+1}(v)) &\leq \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{\lfloor \frac{x_i(v)}{L_i} \rfloor} \left\lceil \frac{x_i(v)}{L_i} \right\rceil \\
&\leq \sum_{c \in L_i(v)} \left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)^{\frac{x_i(v)}{L_i} - 1} \left(\frac{x_i(v)}{L_i} + 1 \right) \\
&\leq \frac{E(l_{i+1}(v))}{\left(1 - \frac{1}{\ln \Delta \cdot L_i}\right)} \left(\frac{X_i}{L_i} + 1 \right) \text{ [by the above and } P(i)] \\
&\leq \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^3 \Delta} \right) X_i \text{ [by Facts 1 and 2].}
\end{aligned}$$

□

As mentioned above, proving that with high probability y_i and the t_i 's decrease enough (Theorems 5.3 and 5.4 below) is simplified if we disregard small t_i 's. We introduce the following parameters:

$$L'_i(v) = \{c \in L_i(v) : t_i(v, c) \geq \Delta^{2/5}\}$$

$$\begin{aligned}
x'_i(v) &= \sum_{c \in L'_i(v)} t_i(v, c) \\
L_i^{*'}(v) &= \{c \in L_i^*(v) : t_{i-1}(v, c) \geq \Delta^{2/5}\} \\
y'_i(v) &= \sum_{c \in L_i^{*'}(v)} t_{i-1}(v, c)
\end{aligned}$$

The parameters x'_i and y'_i are the same as x_i and y_i , respectively, except we have disregarded the small t_i 's. Note that x'_i , like x_i , is measured after Step 5. On the other hand, y'_i , like y_i , is measured before Step 5. The choice of $\Delta^{2/5}$ is used for convenience, but anything between $\ln^3 \Delta$ and $o(\sqrt{\Delta})$ would work.

Note that $E(y'_{i+1}(v)) \leq E(y_{i+1}(v))$, so Lemma 5.1 yields:

$$E(y'_{i+1}(v)) \leq \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^3 \Delta} \right) X_i.$$

It is simpler to show that y'_{i+1} , rather than y_{i+1} , decreases quickly enough.

Lemma 5.3. *For each vertex v and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then*

$$\Pr \left(y'_{i+1}(v) > \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^2 \Delta} \right) X_i \right) < \frac{1}{4} \Delta^{-\ln \Delta}$$

Proof. Instead of showing that y'_{i+1} is concentrated directly, it is easier to look at the difference between x'_i and y'_{i+1} . So we define

$$R = x'_i(v) - y'_{i+1}(v) = \sum_{c \in L'_i(v) \setminus L_i^{*'}(v)} t_i(v, c).$$

One trial can have the largest affect on R by adding or removing a color c from $L'_i(v) \setminus L_i^{*'}(v)$. This changes R by $t_i(v, c)$. Remember that in Modification 1, we discarded colors with large t_i 's, so the value of $t_i(v, c)$ is at most $\Delta^{3/5}$.

Furthermore, for any $s > 0$, we can certify that $X \geq s$, by showing at most $s \cdot \Delta^{-2/5}$ colors which are removed from $L'_i(v)$ since $t_i(v, c) \geq \Delta^{2/5}$ for all $c \in L'_i(v)$. We can certify that a color is removed from the list with just one trial: a neighbor of v which is assigned that color. So Talagrand's Inequality with $c = \Delta^{3/5}$, $r = \Delta^{-2/5}$ and $t = \Delta^{19/20}$, yields:

$$\begin{aligned}
\Pr \left(|R - E(R)| > \Delta^{19/20} \right) &\leq 4e^{-\frac{\Delta^{19/10}}{32\Delta^{6/5}\Delta^{-2/5}(\Delta + \Delta^{19/20})}} \\
&\leq 4e^{-\frac{1}{64}\Delta^{1/10}} \\
&< \frac{1}{4}e^{-\ln^2 \Delta} = \frac{1}{4}\Delta^{-\ln \Delta}
\end{aligned}$$

By the linearity of expectation, Fact 2 and our bound on $E(y'_{i+1}(v))$, the above yields:

$$\Pr \left(y'_{i+1}(v) > \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^2 \Delta} \right) x_i(v) \right)$$

$$\begin{aligned}
&\leq \Pr \left(y'_{i+1}(v) > E(y'_{i+1}(v)) + \Delta^{19/20} \right) \\
&\leq \Pr \left(|R - E(R)| > \Delta^{19/20} \right) < \frac{1}{4} \Delta^{-\ln \Delta}
\end{aligned}$$

□

Now we prove that the parameter t_i decreases quickly enough. We again use $L'_i(v)$ to avoid small t_i 's.

Lemma 5.4. *For each vertex v , color $c \in L'_i(v)$, and $0 \leq i \leq \lceil \frac{10}{\alpha} \ln \Delta \rceil$, if $P(i)$ holds, then*

$$\Pr \left(t_{i+1}(v, c) > \left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) t_i(v, c) \right) < \frac{1}{2\Delta} \Delta^{-\ln \Delta}.$$

Proof. Instead of working with $t_{i+1}(v, c)$ directly, we work with K , the number of vertices in $N_i(v, c)$ which keep their color. Note that $t_{i+1}(v, c) \leq t_i(v, c) - K$.

We run into issues when $t_i(v, c)$ is at least $\sqrt{\Delta}$. So, we partition $N_i(v, c)$ into sets of size $o(\sqrt{\Delta})$. We know that $|N_i(v, c)| > \Delta^{2/5}$ since $c \in L'_i(v)$, so we can partition $N_i(v, c)$ into k nearly equal sized sets, P_1, P_2, \dots, P_k , where $\Delta^{1/4} \leq |P_j| \leq 2\Delta^{1/4}$ for all $1 \leq j \leq k$.

Let K_j be the number of vertices in P_j that keep their color. To prove K_j is large enough, we define some new sets:

$$\begin{aligned}
A_j &= \{v \in P_j : v \text{ is activated}\} \\
B_j &= \{v \in A_j : v \text{ is assigned a unique color among } A_j\} \\
C_j &= \{v \in B_j : v \text{ keeps its color}\}
\end{aligned}$$

Note that $K_j \geq |C_j|$.

We can analyze these random sets by carrying out the random choices in a specific sequence. First, we choose which members of P_j are activated. This determines A_j . To get concentration for $|A_j|$, we can use the Chernoff Bound which yields:

$$\begin{aligned}
\Pr \left(\left| |A_j| - \frac{|P_j|}{\ln \Delta} \right| > \frac{|P_j|}{\ln^2 \Delta} \right) &< 2e^{-\frac{|P_j|^2}{\ln^4 \Delta} \cdot \frac{\ln \Delta}{3|P_j|}} \\
&= e^{-\frac{|P_j|}{3 \ln^3 \Delta}} \leq e^{-\frac{\Delta^{1/4}}{3 \ln^3 \Delta}} < \frac{1}{6} \Delta^{-\ln \Delta - 2}
\end{aligned} \tag{2}$$

Next, we choose the colors assigned to A_j . Assuming $\left| |A_j| - \frac{|P_j|}{\ln \Delta} \right| \leq \frac{|P_j|}{\ln^2 \Delta}$, we bound the probability under this second set of random choices that $|B_j|$ is too small. If (at least) $\frac{|P_j|}{\ln^2 \Delta}$ vertices of A_j are assigned the same color as another vertex in A_j , then there are at least $\frac{|P_j|}{3 \ln^2 \Delta}$ pairs of vertices $a, b \in A_j$, where $c(a) = c(b)$. The probability that there are at least $\frac{|P_j|}{3 \ln^2 \Delta}$ such pairs of vertices is at most:

$$\left(\# \text{ ways to choose } \frac{|P_j|}{3 \ln^2 \Delta} \text{ pairs from at most } \frac{|P_j|}{\ln \Delta} + \frac{|P_j|}{\ln^2 \Delta} \right)$$

$$\begin{aligned}
& \times \Pr(\text{every such pair } (a, b) \text{ has } c(a) = c(b)) \\
& < \left(|P_j| \left(\frac{1}{\ln \Delta} + \frac{1}{\ln^2 \Delta} \right) \right)^{\frac{2|P_j|}{3 \ln^2 \Delta}} \cdot \left(\frac{1}{L_i} \right)^{\frac{|P_j|}{3 \ln^2 \Delta}} \\
& \leq \left(\frac{2|P_j|^2 \left(\frac{1}{\ln \Delta} \right)^2}{L_i} \right)^{\frac{|P_j|}{3 \ln^2 \Delta}} \\
& \leq \left(\frac{8\sqrt{\Delta} \left(\frac{1}{\ln \Delta} \right)^2}{\frac{\sqrt{\Delta}}{\ln \Delta}} \right)^{\frac{|P_j|}{3 \ln^2 \Delta}} \quad [\text{by Fact 1}] \\
& \leq \left(\frac{1}{\ln^3 \Delta} \right)^{\frac{\Delta^{1/4}}{3 \ln^2 \Delta}} \leq \frac{1}{6} \Delta^{-\ln \Delta - 2} \text{ for } \Delta \text{ sufficiently large}
\end{aligned}$$

Therefore, if $\left| |A_j| - \frac{|P_j|}{\ln \Delta} \right| \leq \frac{|P_j|}{\ln^2 \Delta}$, then

$$\begin{aligned}
\Pr \left(|B_j| < |P_j| \left(\frac{1}{\ln \Delta} - \frac{2}{\ln^2 \Delta} \right) \right) & \leq \Pr \left(|B_j| < |A_j| - \frac{|P_j|}{\ln^2 \Delta} \right) \quad (3) \\
& < \frac{1}{6} \Delta^{-\ln \Delta - 2}
\end{aligned}$$

Our final set of random choices is the activation and assignment choices for all vertices outside of P_j . Assuming that the outcomes of the other trials were such that $|B_j| \geq |P_j| \left(\frac{1}{\ln \Delta} - \frac{2}{\ln^2 \Delta} \right)$, we prove that $|C_j|$ is large enough with high probability under this final set of choices. Note that since B_j contains only vertices which were assigned unique colors among P_j , only the colors assigned to vertices outside of P_j can cause a vertex of B_j to become uncolored. We start by computing the expected value of $|C_j|$. This computation is very similar to that in the proof of Lemma 5.1 and we can use the same argument as in Claim 5.1.1 to see that $E(|C_j|)$ is minimized over the rationals when all the t_i 's are equal. Any vertex of $B_j - C_j$ must be uncolored because of a neighbor outside of P_j . Hence, $E(C_j)$ is bounded as follows.

$$\begin{aligned}
E(|C_j|) &= \sum_{u \in B_j} \Pr(u \text{ keeps its color}) \\
&\geq \sum_{u \in B_j} \sum_{\gamma \in L_i(u)} \frac{1}{L_i} \prod_{w \in N_i(u, \gamma)} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right) \\
&= \sum_{u \in B_j} \sum_{\gamma \in L_i(u)} \frac{1}{L_i} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{t_i(u, \gamma)} \\
&\geq \sum_{u \in B_j} \sum_{\gamma \in L_i(u)} \frac{1}{L_i} \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{\frac{x_i}{L_i}} \\
&= |B_j| \left(1 - \frac{1}{\ln \Delta \cdot L_i} \right)^{\frac{x_i}{L_i}}
\end{aligned}$$

$$\begin{aligned}
&\geq |B_j| \left(1 - \frac{X_i}{\ln \Delta \cdot L_i^2} \right) \text{ for } \Delta \text{ sufficiently large} \\
&\geq |B_j| \left(1 - \frac{1}{\ln \Delta} \right) \text{ [by Fact 3]}
\end{aligned}$$

Instead of working with C_j directly, we show concentration of the difference $R = |B_j| - |C_j|$. Note that changing the outcome of any one trial on a vertex not in P_j can at most add or remove one vertex from C_j . Also, for any $s > 0$, we can certify $R \geq s$, by finding s vertices of C_j that are uncolored and looking at the neighbor that got assigned the same color. Recalling that $|P_j| \geq \Delta^{1/4}$, and using the fact that $E(R) \leq |P_j|$, we can apply Talagrand's Inequality with $c = r = 1$ and $t = \frac{|P_j|}{\ln^2 \Delta}$ to obtain:

$$\begin{aligned}
\Pr \left(|R - E(R)| > \frac{|P_j|}{\ln^2 \Delta} \right) &< 4e^{-\frac{|P_j|^2 / \ln^4 \Delta}{32(|P_j| + |P_j|)}} = 4e^{-\frac{|P_j|}{64 \ln^4 \Delta}} \\
&\leq 4e^{-\frac{\Delta^{1/4}}{64 \ln^4 \Delta}} < \frac{1}{6} \Delta^{-\ln \Delta - 2}
\end{aligned}$$

Now, using the linearity of expectation and our bound on $E(|C_j|)$, we get that if $|B_j| \geq |P_j| \left(\frac{1}{\ln \Delta} - \frac{2}{\ln^2 \Delta} \right)$, then

$$\begin{aligned}
&\Pr \left(K_j < |P_j| \cdot \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) \\
&\leq \Pr \left(|C_j| < |P_j| \left(\frac{1}{\ln \Delta} - \frac{2}{\ln^2 \Delta} \right) \left(1 - \frac{1}{\ln \Delta} \right) - \frac{|P_j|}{\ln^2 \Delta} \right) \\
&\leq \Pr \left(|C_j| < |B_j| \left(1 - \frac{1}{\ln \Delta} \right) - \frac{|P_j|}{\ln^2 \Delta} \right) < \frac{1}{6} \Delta^{-\ln \Delta - 2}
\end{aligned} \tag{4}$$

Since $K_j \geq |C_j|$, equations (2), (3) and (4) yield:

$$\begin{aligned}
\Pr \left(K_j < |P_j| \cdot \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) &< \frac{1}{6} \Delta^{-\ln \Delta - 2} + \frac{1}{6} \Delta^{-\ln \Delta - 2} + \frac{1}{6} \Delta^{-\ln \Delta - 2} \\
&= \frac{1}{2} \Delta^{-\ln \Delta - 2}
\end{aligned}$$

Note that $K = \sum_{j=1}^k K_j$, $t_i(v, c) = \sum_{j=1}^k |P_j|$, and $k < \Delta$ to get the following:

$$\Pr \left(K < t_i(v, c) \cdot \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) < \Delta \cdot \frac{1}{2} \Delta^{-\ln \Delta - 2} = \frac{1}{2\Delta} \Delta^{-\ln \Delta}$$

Since $t_{i+1}(v, c) \leq t_i(v, c) - K$, this completes our proof. \square

We are now ready to show that x_i decreases quickly enough. We can use Lemmas 5.3 and 5.4 to prove Lemma 4.4.

Proof of Lemma 4.4. As there are at most $\sqrt{(1+\alpha)\Delta}$ colors in L'_{i+1} , Lemma 5.4 implies that with probability at least

$$1 - \sqrt{(1+\alpha)\Delta} \cdot \frac{1}{2\Delta} \Delta^{-\ln \Delta} \geq 1 - \frac{1}{2} \Delta^{-\ln \Delta}$$

for every $c \in L'_{i+1}$ we have:

$$t_{i+1}(v, c) \leq \left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) t_i(v, c) \quad (5)$$

By Lemma 5.3, with probability at least $1 - \frac{1}{2} \Delta^{-\ln \Delta}$, we have:

$$y'_{i+1}(v) \leq \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^2 \Delta}\right) X_i \quad (6)$$

From the proof of Lemma 4.3, we know that the following holds with probability at least $\frac{1}{4} \Delta^{-\ln \Delta}$:

$$l_{i+1}(v) \geq E(l_{i+1}(v)) - \frac{L_i}{\ln^2 \Delta} \quad (7)$$

Let x_i^* be the following new parameter.

$$x_i^*(v) = \sum_{c \in L_i^*(v)} t_i(v, c)$$

This is the same as x_i , except the value is measured before the truncation step.

If (5), (6) and (7) hold, then:

$$\begin{aligned} x_{i+1}^*(v) &= \sum_{c \in L_{i+1}^{*'}(v)} t_{i+1}(v, c) + \sum_{c \in L_{i+1}^*(v) \setminus L_{i+1}^{*'}(v)} t_{i+1}(v, c) \\ &\leq \left(\sum_{c \in L_{i+1}^{*'}(v)} t_{i+1}(v, c) \right) + \sqrt{(1+\alpha)\Delta} \cdot \Delta^{2/5} \text{ [by definition of } L_i^{*'}(v)] \\ &\leq \left[\sum_{c \in L_{i+1}^{*'}(v)} \left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) t_i(v, c) \right] + \frac{X_i}{\ln^2 \Delta} \text{ [by (5) and Fact 2]} \\ &\leq \left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) \left(\frac{E(l_{i+1}(v))}{L_i} + \frac{1}{\ln^2 \Delta}\right) X_i + \frac{X_i}{\ln^2 \Delta} \text{ [by (6)]} \\ &\leq \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) \left(\frac{E(l_{i+1}(v))}{L_i}\right) + \frac{2}{\ln^2 \Delta} \right] X_i \\ &\leq \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) \left(\frac{l_{i+1}(v) + \frac{L_i}{\ln^2 \Delta}}{L_i}\right) + \frac{2}{\ln^2 \Delta} \right] X_i \text{ [by (7)]} \\ &\leq \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta}\right) \left(\frac{l_{i+1}(v)}{L_i}\right) + \frac{3}{\ln^2 \Delta} \right] X_i \end{aligned}$$

To bound x_{i+1} , we recall that in Step 5, we are removing the c 's with the largest t_i 's. We will keep a proportion of $\frac{L_{i+1}}{l_{i+1}(v)}$ of the colors in v 's list, so, by our choice of which c 's we remove, we get the following:

$$\begin{aligned}
x_{i+1}(v) &\leq \frac{L_{i+1}}{l_{i+1}(v)} \cdot x_{i+1}^*(v) \\
&\leq \frac{L_{i+1}}{l_{i+1}(v)} \cdot \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) \left(\frac{l_{i+1}(v)}{L_i} \right) + \frac{3}{\ln^2 \Delta} \right] X_i \\
&\leq \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) \left(\frac{L_{i+1}}{L_i} \right) + \frac{3}{\ln^2 \Delta} \right] X_i \\
&\leq \left[\left(1 - \frac{1}{(1 + \frac{\alpha}{4}) \ln \Delta} \right) \left(1 - \frac{1}{(1 + \frac{3\alpha}{4}) \ln \Delta} \right) + \frac{3}{\ln^2 \Delta} \right] X_i \\
&\leq \left(1 - \frac{2}{(1 + \frac{\alpha}{2}) \ln \Delta} \right) X_i \quad \text{for } \Delta \text{ sufficiently large} \\
&\quad \text{and } \alpha \text{ sufficiently small} \\
&\leq \left(1 - \frac{1}{(1 + \frac{\alpha}{2}) \ln \Delta} \right)^2 X_i = X_{i+1}
\end{aligned}$$

This holds as long as (5), (6) and (7) hold, which occurs with probability at least $\frac{1}{4}\Delta^{-\ln \Delta} + \frac{1}{2}\Delta^{-\ln \Delta} + \frac{1}{4}\Delta^{-\ln \Delta} = \Delta^{-\ln \Delta}$. \square

6 Concluding Remarks

While our bound is asymptotically optimal, there are still ways in which it could be improved. As mentioned above, there are examples where $\chi_a(G) = \sqrt{\Delta} + 1$ (e.g. K_n^2 for certain n). In response to Reed's paper [21], Bohman and Holzman [3] constructed a graph G with lists $L(v)$ for every vertex v such that: (i) for each vertex v and color $c \in L(v)$, the number of neighbors u of v with $c \in L(u)$ is at most d , (ii) for each vertex v , $|L(v)| = d + 1$, and (iii) G is uncolorable from these lists. If we convert this into an example for adaptable coloring, as described in Section 1, we get an uncolorable graph G with maximum degree $\Delta = d(d + 1)$ and lists of size $d + 1$; thus $\text{ch}_a(G) \geq \sqrt{\Delta} + 2$.

Question 1. Is there a constant k such that $\text{ch}_a(G) \leq \sqrt{\Delta} + k$ for every graph G ?

Determining whether χ_a grows with χ may be the most important question in this area. We mentioned above that χ_a has a lower bound of order $\chi/\sqrt{n \ln \chi}$ [12]. We would like to know whether there is a bound that is only in terms of χ .

Question 2. Is there a function f tending to infinity such that $\chi_a(G) \geq f(\chi(G))$ for every graph G ?

Conceivably, f could be as high as $\Theta(\sqrt{\chi})$. The complete graph has $\chi_a = \Theta(\sqrt{\chi})$ [6], but no example has been found with a smaller adaptable chromatic number in terms of χ .

Question 3. Are there any graphs G with $\chi_a(G) = o(\sqrt{\chi(G)})$?

Intuitively, χ_a is typically much less than χ . There are constructions, however, that demonstrate that χ_a can be equal to χ [14, 15]. These constructions rely on blowing up graphs in a way that maintains χ while increasing χ_a ; they have at least χ^2 vertices but contain the χ -clique. Thus these graphs are far from critical. It would be interesting to see how close χ_a can get to χ for critical graphs. It is easily seen that a critical graph cannot have $\chi_a = \chi$ (for example this follows from Proposition 2.1 in [15]), but nothing more is known.

Question 4. Are there any critical graphs G with $\chi_a(G) = \chi(G) - 1$?

The largest we can get χ_a to be in terms of χ for critical graphs may in fact be much worse than this.

Question 5. Are there any critical graphs G with $\chi_a(G) = \omega(\sqrt{\chi(G)})$?

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