## Linear Subspace Models

Goal: Explore linear models of a data set.

## Motivation:

A central question in vision concerns how we represent a collection of data vectors. The data vectors may be rasterized images, for example.

- We consider the construction of low-dimensional bases for an ensemble of training data using principal components analysis (PCA).
- We introduce PCA, its derivation, its properties, and some of its uses.
- We very briefly critique its suitability for object detection.

Readings: Sections 22.1-22.3 of the Forsyth and Ponce. Matlab Tutorials: colourTutorial.m, trainEigenEyes.m and detectEigenEyes.m

## Representing Images of Human Eyes

Question: Suppose we have a dataset of scaled, centered images of human eyes. How can we find an efficient representation of such a data set?


Left Eyes


Right Eyes

Generative Model. Suppose we can approximate each image in the data set with a parameterized model of the form,

$$
I(\vec{x}) \approx g(\vec{x} ; \vec{a}) .
$$

Here $\vec{a}$ is a vector of coefficients.

## Possible uses:

- compress or reduce the dimension of the data set,
- generate novel instances,
- (possibly) recognition.


## Subspace Appearance Models

Idea: Images are not random, especially those of an object, or similar objects, under different viewing conditions.

$$
3 \quad 3 \quad 3 \quad 3
$$

Rather, than storing every image, we might try to represent the images more effectively, e.g., in a lower dimensional subspace.

For example, let's represent each $N \times N$ image as a point in an $N^{2}-$ dim vector space (e.g., ordering the pixels lexicographically to form the vectors).

(red points denote images, blue vectors denote image differences)
How do we find a low-dimensional basis to accurately model (approximate) each image of the training ensemble (as a linear combination of basis images)?

## Linear Subspace Models

We seek a linear basis with which each image in the ensemble is approximated as a linear combination of basis images $b_{k}(\overrightarrow{\mathbf{x}})$

$$
\begin{equation*}
I(\overrightarrow{\mathbf{x}}) \approx m(\vec{x})+\sum_{k=1}^{K} a_{k} b_{k}(\overrightarrow{\mathbf{x}}) \tag{1}
\end{equation*}
$$

here $m(\vec{x})$ is the mean of the image ensemble. The subspace coefficients $\overrightarrow{\mathbf{a}}=\left(a_{1}, \ldots, a_{K}\right)$ comprise the representaion.

With some abuse of notation, assuming basis images $b_{k}(\overrightarrow{\mathbf{x}})$ with $N^{2}$ pixels, let's define
$\overrightarrow{\mathbf{b}}_{k}-$ an $N^{2} \times 1$ vector with pixels arranged in lexicographic order $\mathbf{B}-$ a matrix with columns $\overrightarrow{\mathbf{b}}_{k}$, i.e., $\mathbf{B}=\left[\overrightarrow{\mathbf{b}}_{1}, \ldots, \overrightarrow{\mathbf{b}}_{K}\right] \in \mathcal{R}^{N^{2} \times K}$

With this notation we can rewrite Eq. (1) in matrix algebra as

$$
\begin{equation*}
\overrightarrow{\mathbf{I}} \approx \vec{m}+\mathbf{B} \overrightarrow{\mathrm{a}} \tag{2}
\end{equation*}
$$

In what follows, we assume that the mean of the ensemble is $\overrightarrow{0}$. (Otherwise, if the ensemble we have is not mean zero, we can estimate the mean and subtract it from each image.)

## Choosing The Basis

Orthogonality: Let's assume orthogonal basis functions,

$$
\left\|\overrightarrow{\mathbf{b}}_{k}\right\|_{2}=1, \quad \overrightarrow{\mathbf{b}}_{j}^{T} \overrightarrow{\mathbf{b}}_{k}=\delta_{j k}
$$

Subspace Coefficients: It follows from the linear model in Eq. (2) and the orthogonality of the basis functions that

$$
\overrightarrow{\mathbf{b}}_{k}^{T} \overrightarrow{\mathbf{I}} \approx \overrightarrow{\mathbf{b}}_{k}^{T} \mathbf{B} \overrightarrow{\mathbf{a}}=\overrightarrow{\mathbf{b}}_{k}^{T}\left[\overrightarrow{\mathbf{b}}_{1}, \ldots, \overrightarrow{\mathbf{b}}_{K}\right] \overrightarrow{\mathbf{a}}=a_{k}
$$

This selection of coefficients, $\overrightarrow{\mathbf{a}}=\mathbf{B}^{T} \overrightarrow{\mathbf{I}}$, minimizes the sum of squared errors (or sum of squared pixel differences, SSD ):

$$
\min _{\overrightarrow{\mathbf{a}} \in \mathcal{R}^{K}}\|\overrightarrow{\mathbf{I}}-\mathbf{B} \overrightarrow{\mathbf{a}}\|_{2}^{2}
$$

Basis Images: In order to select the basis functions $\left\{\overrightarrow{\mathbf{b}}_{k}\right\}_{k=1}^{K}$, suppose we have a training set of images

$$
\left\{\overrightarrow{\mathbf{I}}_{l}\right\}_{l=1}^{L}, \quad \text { with } L \gg K
$$

Recall we are assuming the images are mean zero.
Finally, let's select the basis, $\left\{\overrightarrow{\mathbf{b}}_{k}\right\}_{k=1}^{K}$, to minimize squared reconstruction error:

$$
\sum_{l=1}^{L} \min _{\overrightarrow{\mathbf{a}}_{l}}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \overrightarrow{\mathbf{a}}_{l}\right\|_{2}^{2}
$$

## Intuitions

Example: Let's consider a set of images $\left\{\overrightarrow{\mathbf{I}}_{l}\right\}_{l=1}^{L}$, each with only two pixels. So, each image can be viewed as a 2 D point, $\overrightarrow{\mathbf{I}}_{l} \in \mathcal{R}^{2}$.



For a model with only one basis image, what should $\overrightarrow{\mathbf{b}}_{1}$ be?
Approach: Fit an ellipse to the distribution of the image data, and choose $\overrightarrow{\mathrm{b}}_{1}$ to be a unit vector in the direction of the major axis.

Define the ellipse as $\overrightarrow{\mathbf{x}}^{T} C^{-1} \overrightarrow{\mathbf{x}}=1$, where $C$ is the sample covariance matrix of the image data,

$$
\mathbf{C}=\frac{1}{L} \sum_{l=1}^{L} \overrightarrow{\mathbf{I}}_{l} \overrightarrow{\mathbf{I}}_{l}^{T}
$$

The eigenvectors of $\mathbf{C}$ provide the major axis, i.e.,

$$
\mathbf{C U}=\mathbf{U D}
$$

for orthogonal matrix $\mathbf{U}=\left[\overrightarrow{\mathbf{u}}_{1}, \overrightarrow{\mathbf{u}}_{2}\right]$, and diagonal matrix $\mathbf{D}$ with elements $d_{1} \geq d_{2} \geq 0$. The direction $\overrightarrow{\mathbf{u}}_{1}$ associated with the largest eigenvalue is the direction of the major axis, so let $\overrightarrow{\mathbf{b}}_{1}=\overrightarrow{\mathbf{u}}_{1}$.

## Principal Components Analysis

Theorem: (Minimum reconstruction error) The orthogonal basis B, of rank $K<N^{2}$, that minimizes the squared reconstruction error over training data, $\left\{\overrightarrow{\mathbf{I}}_{l}\right\}_{l=1}^{L}$, i.e.,

$$
\sum_{l=1}^{L} \min _{\overrightarrow{\mathbf{a}}_{l}}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \overrightarrow{\mathbf{a}}_{l}\right\|_{2}^{2}
$$

is given by the first $K$ eigenvectors of the data covariance matrix

$$
\mathbf{C}=\frac{1}{L} \sum_{l=1}^{L} \overrightarrow{\mathbf{I}}_{l} \overrightarrow{\mathbf{I}}_{l}^{T} \in \mathcal{R}^{N^{2} \times N^{2}}, \text { for which } \quad \mathbf{C} \mathbf{U}=\mathbf{U} \mathbf{D}
$$

where $\mathbf{U}=\left[\overrightarrow{\mathbf{u}}_{1}, \ldots, \overrightarrow{\mathbf{u}}_{N^{2}}\right]$ is orthogonal, and $\mathbf{D}=\operatorname{diag}\left(d_{1}, \ldots, d_{N^{2}}\right)$ with $d_{1} \geq d_{2} \geq \ldots \geq d_{N^{2}}$.

That is, the optimal basis vectors are $\overrightarrow{\mathbf{b}}_{k}=\overrightarrow{\mathbf{u}}_{k}$, for $k=1 \ldots K$. The corresponding basis images $\left\{b_{k}(\overrightarrow{\mathbf{x}})\right\}_{k=1}^{K}$ are often called eigen-images.

Proof: see the derivation below.

## Derivation of PCA

To begin, we want to find $\mathbf{B}$ in order to minimize squared error in subspace approximations to the images of the training ensemble.

$$
E=\sum_{l=1}^{L} \min _{\overrightarrow{\mathbf{a}}_{l}}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \overrightarrow{\mathbf{a}}_{l}\right\|_{2}^{2}
$$

Given the assumption that the columns of $\mathbf{B}$ are orthonormal, the optimal coefficients are $\overrightarrow{\mathbf{a}}_{l}=\mathbf{B}^{T} \overrightarrow{\mathbf{I}}_{l}$, so

$$
\begin{equation*}
E=\sum_{l=1}^{L} \min _{\overrightarrow{\mathbf{a}}_{l}}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \overrightarrow{\mathbf{a}}_{l}\right\|_{2}^{2}=\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \mathbf{B}^{T} \overrightarrow{\mathbf{I}}_{l}\right\|_{2}^{2} \tag{3}
\end{equation*}
$$

Furthermore, writing the each training image as a column in a matrix $\mathbf{A}=\left[\overrightarrow{\mathbf{I}}_{1}, \ldots, \overrightarrow{\mathbf{I}}_{L}\right]$, we have
$E=\sum_{l=1}^{L}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \mathbf{B}^{T} \overrightarrow{\mathbf{I}}_{l}\right\|_{2}^{2}=\left\|\mathbf{A}-\mathbf{B} \mathbf{B}^{T} \mathbf{A}\right\|_{F}^{2}=\operatorname{trace}\left[\mathbf{A} \mathbf{A}^{T}\right]-\operatorname{trace}\left[\mathbf{B}^{T} \mathbf{A} \mathbf{A}^{T} \mathbf{B}\right]$
You get this last step by expanding the square and noting $\mathbf{B}^{T} \mathbf{B}=\mathbf{I}_{K}$, and using the properties of trace, e.g., trace $[\mathbf{A}]=\operatorname{trace}\left[\mathbf{A}^{T}\right]$, and also trace $\left[\mathbf{B}^{T} \mathbf{A} \mathbf{A}^{T} \mathbf{B}\right]=\operatorname{trace}\left[\mathbf{A}^{T} \mathbf{B B}^{T} \mathbf{A}\right]$.
So to minmize the average squared error in the approximation we want to find $\mathbf{B}$ to maximize

$$
\begin{equation*}
E^{\prime}=\operatorname{trace}\left[\mathbf{B}^{T} \mathbf{A} \mathbf{A}^{T} \mathbf{B}\right] \tag{4}
\end{equation*}
$$

Now, let's use the fact that for the data covariance, $\mathbf{C}$ we have $\mathbf{C}=\frac{1}{L} \mathbf{A} \mathbf{A}^{T}$. Moreover, as defined above the SVD of $\mathbf{C}$ can be written as $\mathbf{C}=\mathbf{U} \mathbf{D ~ U}^{T}$. So, let's substitute the SVD into $E^{\prime}$ :

$$
\begin{equation*}
E^{\prime}=\operatorname{trace}\left[\mathbf{B}^{T} \mathbf{U} \mathbf{D} \mathbf{U}^{T} \mathbf{B}\right] \tag{5}
\end{equation*}
$$

where of course $\mathbf{U}$ is orthogonal, and $\mathbf{D}$ is diagonal.

Now we just have to show that we want to choose $\mathbf{B}$ such that the trace strips off the first $K$ elements of $\mathbf{D}$ to maximize $E^{\prime}$. Intuitively, note that $\mathbf{B}^{T} \mathbf{U}$ must be rank $K$ since $\mathbf{B}$ is rank $K$. And note that the diagonal elements of $\mathbf{D}$ are ordered. Also the trace is invariant under matrix rotation. So, the highest rank $K$ trace we can hope to get is by choosing $\mathbf{B}$ so that, when combined with $\mathbf{U}$ we keep the first $K$ columns of $\mathbf{D}$. That is, the columns of $\mathbf{B}$ should be the first $K$ orthonormal rows of $\mathbf{U}$. We need to make this a little more rigorous, but that's it for now...

## Other Properties of PCA

Maximum Variance: The $K$-D subspace approximation captures the greatest possible variance in the training data.

- For $a_{1}=\overrightarrow{\mathbf{b}}_{1}^{T} \overrightarrow{\mathbf{I}}$, the direction $\overrightarrow{\mathbf{b}}_{1}$ that maximizes the variance $\mathrm{E}\left[a_{1}^{2}\right]=$ $\overrightarrow{\mathbf{b}}_{1}^{T} \mathbf{C} \overrightarrow{\mathbf{b}}_{1}$, subject to $\overrightarrow{\mathbf{b}}_{1}^{T} \overrightarrow{\mathbf{b}}_{1}=1$, is the first eigenvector of $\mathbf{C}$.
- The second maximizes $\overrightarrow{\mathbf{b}}_{2}^{T} \mathbf{C} \overrightarrow{\mathbf{b}}_{2}$ subject to $\overrightarrow{\mathbf{b}}_{2}^{T} \overrightarrow{\mathbf{b}}_{2}=1$ and $\overrightarrow{\mathbf{b}}_{1}^{T} \overrightarrow{\mathbf{b}}_{2}=0$.
- For $a_{k}=\overrightarrow{\mathbf{b}}_{k}^{T} \overrightarrow{\mathbf{I}}$, and $\overrightarrow{\mathbf{a}}=\left(a_{1}, \ldots, a_{K}\right)$, the subspace coefficient covariance is $\mathrm{E}\left[\overrightarrow{\mathbf{a}} \overrightarrow{\mathbf{a}}^{T}\right]=\operatorname{diag}\left(d_{1}, \ldots, d_{K}\right)$. That is, the diagonal entries of $\mathbf{D}$ are marginal variances of the subspace coefficients:

$$
\sigma_{k}^{2} \equiv \mathrm{E}\left[a_{k}^{2}\right]=d_{k} .
$$

So the total variance captured in the subspace is sum of first $K$ eigenvalues of $\mathbf{C}$.

- Total variance lost owing to the subspace projection (i.e., the out-of-subspace variance) is the sum of the last $N^{2}-K$ eigenvalues:

$$
\frac{1}{L} \sum_{l=1}^{L}\left[\min _{\overrightarrow{\mathbf{a}}_{l}}\left\|\overrightarrow{\mathbf{I}}_{l}-\mathbf{B} \overrightarrow{\mathbf{a}}_{l}\right\|_{2}^{2}\right]=\sum_{k=K+1}^{N^{2}} \sigma_{k}^{2}
$$

Decorrelated Coefficients: C is diagonalized by its eigenvectors, so $\mathbf{D}$ is diagonal, and the subspace coefficients are uncorrelated.

- Under a Gaussian model of the images (where the images are drawn from an $N^{2}$-dimensional Gaussian pdf), this means that the coefficients are also statistically independent.


## PCA and Singular Value Decomposition

The singular value decomposition of the data matrix $\mathbf{A}$,

$$
\mathbf{A}=\left[\overrightarrow{\mathbf{I}}_{1}, \ldots, \overrightarrow{\mathbf{I}}_{L}\right], \quad \mathbf{A} \in \mathcal{R}^{N^{2} \times L}, \quad \text { where usually } L \ll N^{2} .
$$

is given by

$$
\mathbf{A}=\mathbf{U} \mathbf{S} \mathbf{V}^{T}
$$

where $\mathbf{U} \in \mathcal{R}^{N^{2} \times L}, \mathbf{S} \in \mathcal{R}^{L \times L}, \mathbf{V} \in \mathcal{R}^{L \times L}$. The columns of $\mathbf{U}$ and $\mathbf{V}$ are orthogonal, i.e., $\mathbf{U}^{T} \mathbf{U}=\mathbf{I}_{L \times L}$ and $\mathbf{V}^{T} \mathbf{V}=\mathbf{I}_{L \times L}$, and matrix $\mathbf{S}$ is diagonal, $\mathbf{S}=\operatorname{diag}\left(s_{1}, \ldots, s_{L}\right)$ where $s_{1} \geq s_{2} \geq$ $\ldots \geq s_{L} \geq 0$.

Theorem: The best rank- $K$ approximation to $\mathbf{A}$ under the Frobenius norm, $\tilde{\mathbf{A}}$, is given by

$$
\tilde{\mathbf{A}}=\sum_{k=1}^{K} s_{k} \overrightarrow{\mathbf{u}}_{k} \overrightarrow{\mathbf{v}}_{k}^{T}=\mathbf{B} \mathbf{B}^{T} \mathbf{A}, \text { where } \min _{\operatorname{rank}(\tilde{\mathbf{A}})=K}\|\mathbf{A}-\tilde{\mathbf{A}}\|_{F}^{2}=\sum_{k=K+1}^{N^{2}} s_{k}^{2},
$$

and $\mathbf{B}=\left[\overrightarrow{\mathbf{u}}_{1}, \ldots, \overrightarrow{\mathbf{u}}_{K}\right] . \tilde{\mathbf{A}}$ is also the best rank- $K$ approximation under the $L_{2}$ matrix norm.

What's the relation to PCA and the covariance of the training images?

$$
\mathbf{C}=\frac{1}{L} \sum_{l=1}^{L} \overrightarrow{\mathbf{I}}_{l} \overrightarrow{\mathbf{I}}_{l}^{T}=\frac{1}{L} \mathbf{A} \mathbf{A}^{T}=\frac{1}{L} \mathbf{U} \mathbf{S} \mathbf{V}^{T} \mathbf{V} \mathbf{S}^{T} \mathbf{U}^{T}=\frac{1}{L} \mathbf{U} \mathbf{S}^{2} \mathbf{U}^{T}
$$

So the squared singular values of $\mathbf{A}$ are proportional to the first $L$ eigenvalues of $\mathbf{C}$ :

$$
d_{k}=\left\{\begin{array}{cl}
\frac{1}{L} s_{k}^{2} & \text { for } k=1, \ldots, L \\
0 & \text { for } k>L
\end{array}\right.
$$

And the singular vectors of $\mathbf{A}$ are just the first $L$ eigenvectors of $\mathbf{C}$.

## Eigen-Images for Generic Images?

Fourier components are eigenfunctions of generic image ensembles.


Why? Covariance matrices for stationary processes are Toeplitz.


PCA yields unique eigen-images up to rotations of invariant subspaces (e.g., Fourier components with the same marginal variance).

## Eigen-Reflectances

Consider an ensemble of surface reflectances $r(\lambda)$.


What is the effective dimension of these reflectances? Define $V_{k} \equiv \sum_{j=1}^{k} \sigma_{j}^{2}$. Then the fraction of total variance explained by the first $k$ PCA components is $Q_{k} \equiv V_{k} / V_{L}$.

Fraction of Explained Variance


Reflectances $r(\lambda)$, for wavelengths $\lambda$ within the visible spectrum, are effectively 3 dimensional (see colourTutorial.m).

## Eye Subspace Model

Subset of 1196 eye images $(25 \times 20)$ :


Left Eyes


Right Eyes

$$
\text { Defn: Let } V_{k} \equiv \sum_{j=1}^{k} s_{j}^{2}, \quad d Q_{k} \equiv s_{k}^{2} / V_{L}, \quad \text { and } Q_{k} \equiv V_{k} / V_{L}:
$$




Left plot shows $d Q_{k}$, the fraction of the total variance contributed by the $k^{t h}$ principal component.
Right plot shows $Q_{k}$ the fraction of the total variance captured by the subspace formed from the first $k$ principal components.

## Eye Subspace Model

## Mean Eye:



Basis Images (1-6, and 10:5:35):


Reconstructions (for $K=5,20,50$ ):


Eye Image


Eye Image


Reconstruction ( $\mathrm{K}=5$ )


Reconstruction ( $\mathrm{K}=5$ )


Reconstruction ( $\mathrm{K}=20$ )


Reconstruction ( $\mathrm{K}=20$ )


Reconstruction ( $\mathrm{K}=50$ )


Reconstruction ( $\mathrm{K}=50$ )

## Generative Eye Model

Generative model, $\mathcal{M}$, for random eye images:

$$
\overrightarrow{\mathbf{I}}=\overrightarrow{\mathbf{m}}+\left(\sum_{k=1}^{K} a_{k} \overrightarrow{\mathbf{b}}_{k}\right)+\overrightarrow{\mathbf{e}}
$$

where $\overrightarrow{\mathbf{m}}$ is the mean eye image, $a_{k} \sim \mathcal{N}\left(0, \sigma_{k}^{2}\right), \sigma_{k}^{2}$ is the sample variance associated with the $k^{t h}$ principal direction in the training data, and $\overrightarrow{\mathbf{e}} \sim \mathcal{N}\left(0, \sigma_{e}^{2} \mathbf{I}_{N^{2}}\right)$ where $\sigma_{e}^{2}=\frac{1}{N^{2}} \sum_{k=K+1}^{N^{2}} \sigma_{k}^{2}$ is the per pixel out-of-subspace variance.

## Random Eye Images:



Random draws from generative model (with $\mathrm{K}=5,10,20,50,100,200$ )

So the probability of an image given this model $\mathcal{M}$ is

$$
p(\overrightarrow{\mathbf{I}} \mid \mathcal{M})=\left(\prod_{k=1}^{K} p\left(a_{k} \mid \mathcal{M}\right)\right) p(\overrightarrow{\mathbf{e}} \mid \mathcal{M})
$$

where

$$
p\left(a_{k} \mid \mathcal{M}\right)=\frac{1}{\sqrt{2 \pi} \sigma_{k}} e^{-\frac{a_{k}^{2}}{2 \sigma_{k}^{2}}}, \quad p(\overrightarrow{\mathbf{e}} \mid \mathcal{M})=\prod_{j=1}^{N^{2}} \frac{1}{\sqrt{2 \pi} \sigma_{e}} e^{-\frac{e_{j}^{2}}{2 \sigma_{e}^{2}}} .
$$

## Face Detection

The wide-spread use of PCA for object recognition began with the work Turk and Pentland (1991) for face detection and recognition.

Shown below is the model learned from a collection of frontal faces, normalized for contrast, scale, and orientation, with the backgrounds removed prior to PCA.


Here are the mean image (upper-left) and the first 15 eigen-images. The first three show strong variations caused by illumination. The next few appear to correspond to the occurrence of certain features (hair, hairline, beard, clothing, etc).

## Object Recognition

Murase and Nayar (1995)

- images of multiple objects, taken from different positions on the viewsphere
- each object occupies a manifold in the subspace (as a function of position on the viewsphere)
- recognition: nearest neighbour assuming dense sampling of object pose variations in the training set.



## Summary

## The generative model:

- PCA finds the subspace (of a specified dimension) that maximizes projected signal variance.
- A single Gaussian model is naturally associated with a PCA representation. The principal axes are the principal directions of the Gaussian's covariance.


## Issues:

- The single Gaussian model is often rather crude. PCA coeff's can exhibit significantly more structure (cf. Murase \& Nayar).
- As a result of this unmodelled structure, detectors based on single Gaussian models are often poor. See the Matlab tutorial detectEigenEyes.m.
- We discuss alternative detection strategies later in this course.


## Further Readings

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