

Linear Subspace Models

Goal: Explore linear models of data.

Motivation: A central question in vision concerns how we represent a collection of data vectors, such as images of an object under a wide range of viewing conditions (lighting and viewpoints).

- We consider the construction of low-dimensional bases for an ensemble of training data using principal components analysis (PCA).
- We introduce PCA, its derivation, its properties, and some of its uses.
- We briefly critique its suitability for object detection.

Readings: Sections 22.1–22.3 of the Forsyth and Ponce.

Matlab Tutorials: colourTutorial.m, trainEigenEyes.m and detectEigenEyes.m

Representing Images of Human Eyes

Question: Suppose we have a dataset of scaled, aligned images of human eyes. How can we find an efficient representation of them?



Left Eyes

Right Eyes

Generative Model: For example, suppose we can approximate each image in the data set with a parameterized model of the form

$$I(\vec{x}) \approx g(\vec{x}, \vec{a}),$$

where \vec{a} is a (low-dimensional) vector of coefficients.

Possible uses:

- reduce the dimension of the data set (compression)
- generate novel instances (density estimation)
- (possibly) detection/recognition

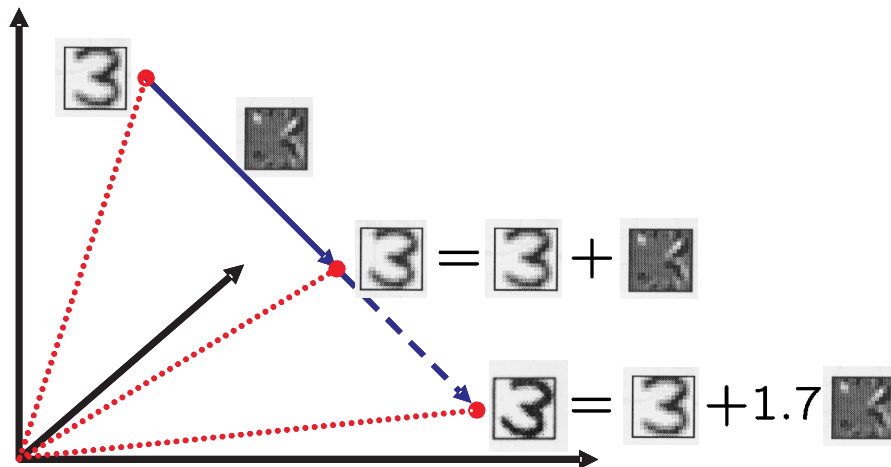
Subspace Appearance Models

Idea: Images are not random, especially those of an object, or similar objects, under different viewing conditions.



Rather, than storing every image, we might try to represent the images more effectively, e.g., in a lower dimensional *subspace*.

For example, let's represent each $N \times N$ image as a point in an N^2 -dimensional vector space (e.g., ordering the pixels lexicographically to form the vectors).



(red points denote images, blue vectors denote image differences)

How do we find a low-dimensional basis to accurately model (approximate) each image of the training ensemble (as a linear combination of basis images)?

Linear Subspace Models

We seek a linear basis with which each image in the ensemble is approximated as a linear combination of basis images $b_k(\vec{\mathbf{x}})$:

$$I(\vec{\mathbf{x}}) \approx m(\vec{\mathbf{x}}) + \sum_{k=1}^K a_k b_k(\vec{\mathbf{x}}), \quad (1)$$

where $m(\vec{\mathbf{x}})$ is the mean of the image ensemble. The *subspace coefficients* $\vec{\mathbf{a}} = (a_1, \dots, a_K)$ comprise the representation.

With some abuse of notation, assuming basis images $b_k(\vec{\mathbf{x}})$ with N^2 pixels, let's define

- $\vec{\mathbf{b}}_k$ – an $N^2 \times 1$ vector with pixels arranged in lexicographic order
- \mathbf{B} – a matrix with columns $\vec{\mathbf{b}}_k$, *i.e.*, $\mathbf{B} = [\vec{\mathbf{b}}_1, \dots, \vec{\mathbf{b}}_K] \in \mathcal{R}^{N^2 \times K}$

With this notation we can rewrite Eq. (1) in matrix algebra as

$$\vec{\mathbf{I}} \approx \vec{\mathbf{m}} + \mathbf{B} \vec{\mathbf{a}}. \quad (2)$$

In what follows, we assume that the mean of the ensemble is $\vec{\mathbf{0}}$.

(Otherwise, if the ensemble we have is not mean zero, we can estimate the mean and subtract it from each image.)

Choosing The Basis

Orthogonality: Let's assume orthonormal basis functions,

$$\| \vec{\mathbf{b}}_k \|_2 = 1 \quad , \quad \vec{\mathbf{b}}_j^T \vec{\mathbf{b}}_k = \delta_{jk} .$$

Subspace Coefficients: It follows from the linear model in Eq. (2) and the orthogonality of the basis functions that

$$\vec{\mathbf{b}}_k^T \vec{\mathbf{I}} \approx \vec{\mathbf{b}}_k^T \mathbf{B} \vec{\mathbf{a}} = \vec{\mathbf{b}}_k^T [\vec{\mathbf{b}}_1, \dots, \vec{\mathbf{b}}_K] \vec{\mathbf{a}} = a_k$$

This selection of coefficients, $\vec{\mathbf{a}} = \mathbf{B}^T \vec{\mathbf{I}}$, minimizes the sum of squared errors (or sum of squared pixel differences, SSD):

$$\min_{\vec{\mathbf{a}} \in \mathcal{R}^K} \| \vec{\mathbf{I}} - \mathbf{B} \vec{\mathbf{a}} \|_2^2$$

Basis Images: In order to select the basis functions $\{\vec{\mathbf{b}}_k\}_{k=1}^K$, suppose we have a training set of images

$$\{ \vec{\mathbf{I}}_l \}_{l=1}^L \quad , \quad \text{with } L \gg K$$

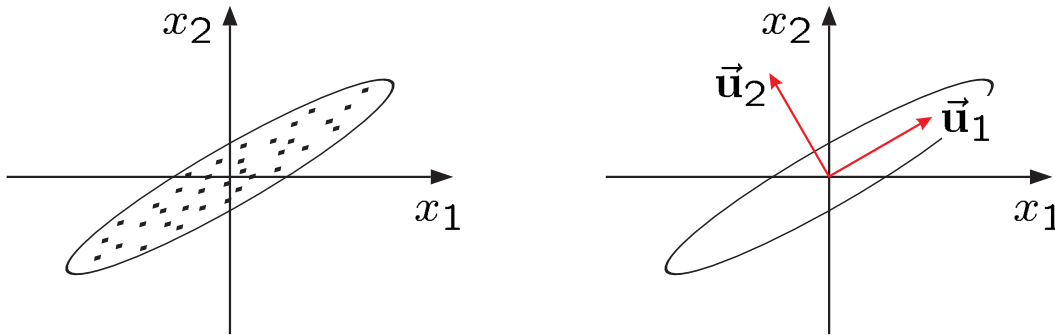
Recall we are assuming the images are mean zero: i.e., $\frac{1}{L} \sum_l \vec{\mathbf{I}}_l = 0$,

Let's select the basis, \mathbf{B} , to minimize squared reconstruction error:

$$\sum_{l=1}^L \min_{\vec{\mathbf{a}}_l} \| \vec{\mathbf{I}}_l - \mathbf{B} \vec{\mathbf{a}}_l \|_2^2$$

Intuitions

Example: let's consider a set of images $\{\vec{\mathbf{I}}_l\}_{l=1}^L$, each with only two pixels. So, each image can be viewed as a 2D point, $\vec{\mathbf{I}}_l \in \mathcal{R}^2$.



For a model with only one basis image, what should $\vec{\mathbf{b}}_1$ be?

Approach: Fit an ellipse to the distribution of the image data, and choose $\vec{\mathbf{b}}_1$ to be a unit vector in the direction of the major axis.

Define the ellipse as $\vec{\mathbf{x}}^T \mathbf{C}^{-1} \vec{\mathbf{x}} = 1$, where \mathbf{C} is the sample covariance matrix of the image data,

$$\mathbf{C} = \frac{1}{L} \sum_{l=1}^L \vec{\mathbf{I}}_l \vec{\mathbf{I}}_l^T$$

The eigenvectors of \mathbf{C} provide the major axis, i.e.,

$$\mathbf{C} \mathbf{U} = \mathbf{U} \mathbf{D}$$

for orthogonal matrix $\mathbf{U} = [\vec{\mathbf{u}}_1, \vec{\mathbf{u}}_2]$, and diagonal matrix \mathbf{D} with elements $d_1 \geq d_2 \geq 0$. The direction $\vec{\mathbf{u}}_1$ associated with the largest eigenvalue, d_1 , is the direction of the major axis, so let $\vec{\mathbf{b}}_1 = \vec{\mathbf{u}}_1$.

Principal Components Analysis

Theorem: (*Minimum reconstruction error*) The orthogonal basis \mathbf{B} , of rank $K < N^2$, that minimizes the squared reconstruction error over training data, $\{\vec{\mathbf{I}}_l\}_{l=1}^L$, i.e.,

$$\sum_{l=1}^L \min_{\vec{\mathbf{a}}_l} \|\vec{\mathbf{I}}_l - \mathbf{B} \vec{\mathbf{a}}_l\|_2^2$$

is given by the first K eigenvectors of the data covariance matrix

$$\mathbf{C} = \frac{1}{L} \sum_{l=1}^L \vec{\mathbf{I}}_l \vec{\mathbf{I}}_l^T \in \mathcal{R}^{N^2 \times N^2}, \text{ for which } \mathbf{C} \mathbf{U} = \mathbf{U} \mathbf{D}$$

where $\mathbf{U} = [\vec{\mathbf{u}}_1, \dots, \vec{\mathbf{u}}_{N^2}]$ is orthogonal, and $\mathbf{D} = \text{diag}(d_1, \dots, d_{N^2})$ with $d_1 \geq d_2 \geq \dots \geq d_{N^2}$.

That is, the optimal basis vectors are $\vec{\mathbf{b}}_k = \vec{\mathbf{u}}_k$, for $k = 1 \dots K$. The corresponding basis images $\{b_k(\vec{\mathbf{x}})\}_{k=1}^K$ are often called eigen-images.

Proof: see the derivation below.

Derivation of PCA

To begin, we want to find \mathbf{B} in order to minimize squared error in subspace approximations to the images of the training ensemble.

$$E = \sum_{l=1}^L \min_{\vec{\mathbf{a}}_l} \|\vec{\mathbf{I}}_l - \mathbf{B} \vec{\mathbf{a}}_l\|_2^2$$

Given the assumption that the columns of \mathbf{B} are orthonormal, the optimal coefficients are $\vec{\mathbf{a}}_l = \mathbf{B}^T \vec{\mathbf{I}}_l$, so

$$E = \sum_{l=1}^L \min_{\vec{\mathbf{a}}_l} \|\vec{\mathbf{I}}_l - \mathbf{B} \vec{\mathbf{a}}_l\|_2^2 = \|\vec{\mathbf{I}}_l - \mathbf{B} \mathbf{B}^T \vec{\mathbf{I}}_l\|_2^2 \quad (3)$$

Furthermore, writing the each training image as a column in a matrix $\mathbf{A} = [\vec{\mathbf{I}}_1, \dots, \vec{\mathbf{I}}_L]$, we have

$$E = \sum_{l=1}^L \|\vec{\mathbf{I}}_l - \mathbf{B} \mathbf{B}^T \vec{\mathbf{I}}_l\|_2^2 = \|\mathbf{A} - \mathbf{B} \mathbf{B}^T \mathbf{A}\|_F^2 = \text{trace}[\mathbf{A} \mathbf{A}^T] - \text{trace}[\mathbf{B}^T \mathbf{A} \mathbf{A}^T \mathbf{B}]$$

You get this last step by expanding the square and noting $\mathbf{B}^T \mathbf{B} = \mathbf{I}_K$, and using the properties of *trace*, e.g., $\text{trace}[\mathbf{A}] = \text{trace}[\mathbf{A}^T]$, and also $\text{trace}[\mathbf{B}^T \mathbf{A} \mathbf{A}^T \mathbf{B}] = \text{trace}[\mathbf{A}^T \mathbf{B} \mathbf{B}^T \mathbf{A}]$.

So to minimize the average squared error in the approximation we want to find \mathbf{B} to maximize

$$E' = \text{trace}[\mathbf{B}^T \mathbf{A} \mathbf{A}^T \mathbf{B}] \quad (4)$$

Now, let's use the fact that for the data covariance, \mathbf{C} we have $\mathbf{C} = \frac{1}{L} \mathbf{A} \mathbf{A}^T$. Moreover, as defined above the SVD of \mathbf{C} can be written as $\mathbf{C} = \mathbf{U} \mathbf{D} \mathbf{U}^T$. So, let's substitute the SVD into E' :

$$E' = \text{trace}[\mathbf{B}^T \mathbf{U} \mathbf{D} \mathbf{U}^T \mathbf{B}] \quad (5)$$

where of course \mathbf{U} is orthogonal, and \mathbf{D} is diagonal.

Now we just have to show that we want to choose \mathbf{B} such that the trace strips off the first K elements of \mathbf{D} to maximize E' . Intuitively, note that $\mathbf{B}^T \mathbf{U}$ must be rank K since \mathbf{B} is rank K . And note that the diagonal elements of \mathbf{D} are ordered. Also the trace is invariant under matrix rotation. So, the highest rank K trace we can hope to get is by choosing \mathbf{B} so that, when combined with \mathbf{U} we keep the first K columns of \mathbf{D} . That is, the columns of \mathbf{B} should be the first K orthonormal rows of \mathbf{U} . We need to make this a little more rigorous, but that's it for now...

Other Properties of PCA

Maximum Variance: PCA also gives the K -D subspace that captures the greatest fraction of the total variance in the training data.

- For $a_1 = \vec{\mathbf{b}}_1^T \vec{\mathbf{I}}$, the direction $\vec{\mathbf{b}}_1$ that maximizes the coef variance $E[a_1^2] = \vec{\mathbf{b}}_1^T \mathbf{C} \vec{\mathbf{b}}_1$, s.t. $\vec{\mathbf{b}}_1^T \vec{\mathbf{b}}_1 = 1$, is the first eigenvector of \mathbf{C} .
- The second maximizes $\vec{\mathbf{b}}_2^T \mathbf{C} \vec{\mathbf{b}}_2$ subject to $\vec{\mathbf{b}}_2^T \vec{\mathbf{b}}_2 = 1$ and $\vec{\mathbf{b}}_1^T \vec{\mathbf{b}}_2 = 0$.
- For $a_k = \vec{\mathbf{b}}_k^T \vec{\mathbf{I}}$, and $\vec{\mathbf{a}} = (a_1, \dots, a_K)$, the subspace coefficient covariance is $E[\vec{\mathbf{a}} \vec{\mathbf{a}}^T] = \text{diag}(d_1, \dots, d_K)$. That is, the diagonal entries of \mathbf{D} are marginal variances of the subspace coefficients:

$$\sigma_k^2 \equiv E[a_k^2] = d_k .$$

So the total variance *captured* in the subspace is sum of first K eigenvalues of \mathbf{C} .

- Total variance *lost* owing to the subspace projection (i.e., the out-of-subspace variance) is the sum of the last $N^2 - K$ eigenvalues:

$$\frac{1}{L} \sum_{l=1}^L \left[\min_{\vec{\mathbf{a}}_l} \|\vec{\mathbf{I}}_l - \mathbf{B} \vec{\mathbf{a}}_l\|_2^2 \right] = \sum_{k=K+1}^{N^2} \sigma_k^2$$

Decorrelated Coefficients: \mathbf{C} is diagonalized by its eigenvectors, so \mathbf{D} is diagonal, and the subspace coefficients are uncorrelated.

- Under a Gaussian model of the images (where the images are drawn from an N^2 -dimensional Gaussian pdf), this means that the coefficients are also statistically independent.

PCA and Singular Value Decomposition

The singular value decomposition (SVD) of the data matrix \mathbf{A} ,

$$\mathbf{A} = [\vec{\mathbf{I}}_1, \dots, \vec{\mathbf{I}}_L] , \quad \mathbf{A} \in \mathcal{R}^{N^2 \times L} , \quad \text{where usually } L \ll N^2 .$$

is given by

$$\mathbf{A} = \mathbf{U} \mathbf{S} \mathbf{V}^T$$

where $\mathbf{U} \in \mathcal{R}^{N^2 \times L}$, $\mathbf{S} \in \mathcal{R}^{L \times L}$, $\mathbf{V} \in \mathcal{R}^{L \times L}$. The columns of \mathbf{U} and \mathbf{V} are orthogonal, i.e., $\mathbf{U}^T \mathbf{U} = \mathbf{I}_{L \times L}$ and $\mathbf{V}^T \mathbf{V} = \mathbf{I}_{L \times L}$, and matrix \mathbf{S} is diagonal, $\mathbf{S} = \text{diag}(s_1, \dots, s_L)$ where $s_1 \geq s_2 \geq \dots \geq s_L \geq 0$.

Theorem: The best rank- K approximation to \mathbf{A} under the Frobenius norm, $\tilde{\mathbf{A}}$, is given by

$$\tilde{\mathbf{A}} = \sum_{k=1}^K s_k \vec{\mathbf{u}}_k \vec{\mathbf{v}}_k^T = \mathbf{B} \mathbf{B}^T \mathbf{A} , \quad \text{where } \min_{\text{rank}(\tilde{\mathbf{A}})=K} \|\mathbf{A} - \tilde{\mathbf{A}}\|_F^2 = \sum_{k=K+1}^{N^2} s_k^2 ,$$

and $\mathbf{B} = [\vec{\mathbf{u}}_1, \dots, \vec{\mathbf{u}}_K]$. $\tilde{\mathbf{A}}$ is also the best rank- K approximation under the L_2 matrix norm.

What's the relation to PCA and the covariance of the training images?

$$\mathbf{C} = \frac{1}{L} \sum_{l=1}^L \vec{\mathbf{I}}_l \vec{\mathbf{I}}_l^T = \frac{1}{L} \mathbf{A} \mathbf{A}^T = \frac{1}{L} \mathbf{U} \mathbf{S} \mathbf{V}^T \mathbf{V} \mathbf{S}^T \mathbf{U}^T = \frac{1}{L} \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

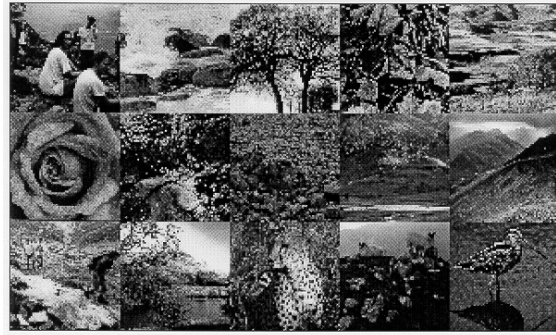
So the squared singular values of \mathbf{A} are proportional to the first L eigenvalues of \mathbf{C} :

$$d_k = \begin{cases} \frac{1}{L} s_k^2 & \text{for } k = 1, \dots, L \\ 0 & \text{for } k > L \end{cases}$$

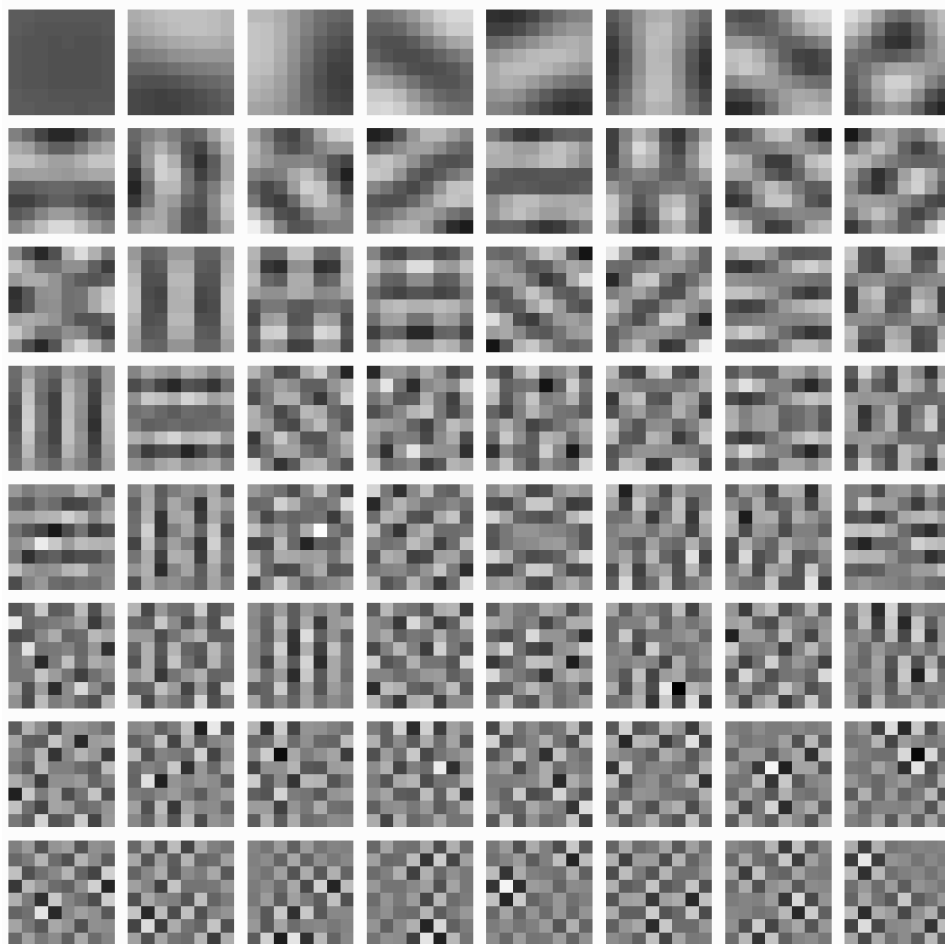
And the singular vectors of \mathbf{A} are just the first L eigenvectors of \mathbf{C} .

Eigen-Images for Generic Images?

Fourier components are eigenfunctions of generic image ensembles.



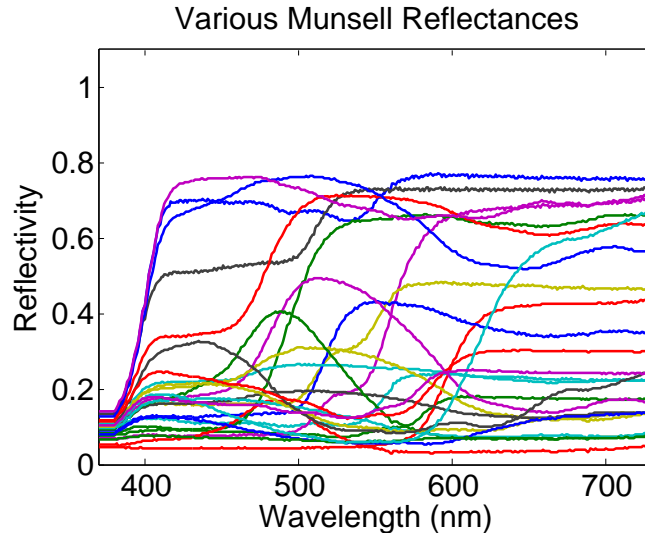
Why? Covariance matrices for stationary processes are Toeplitz.



PCA yields unique eigen-images up to rotations of invariant subspaces (e.g., Fourier components with the same marginal variance).

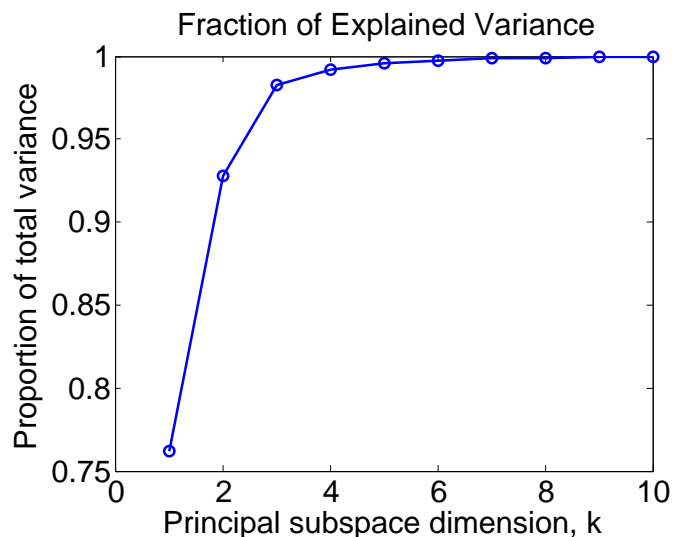
Eigen-Reflectances

Consider an ensemble of surface reflectance functions (BRDFs), $r(\lambda)$.



What is the effective dimension of these reflectances?

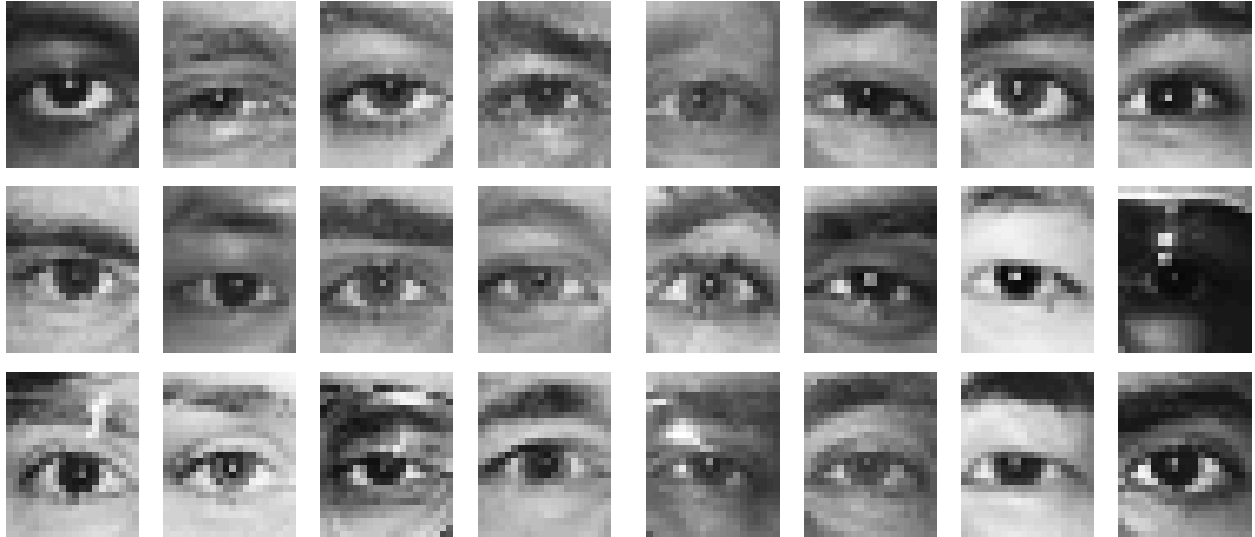
Let $V_k \equiv \sum_{j=1}^k \sigma_j^2$. Then the fraction of total variance explained by the first k PCA components is $Q_k \equiv V_k/V_L$.



Reflectances $r(\lambda)$, for wavelengths λ within the visible spectrum, are effectively 3 dimensional (see `colourTutorial.m`).

Eye Subspace Model

Subset of 1196 eye images (25×20):

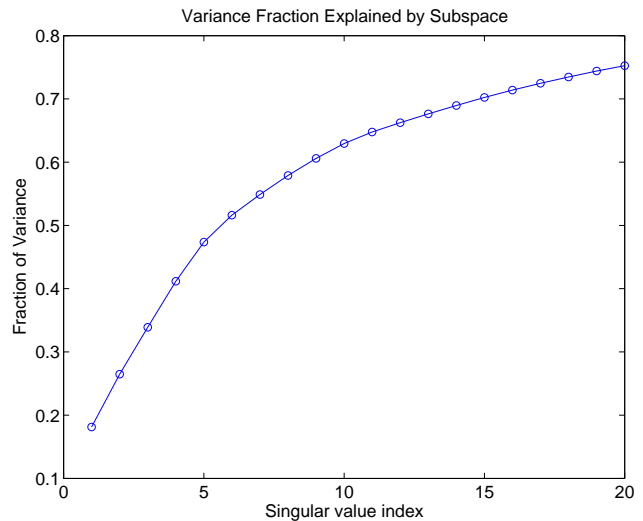
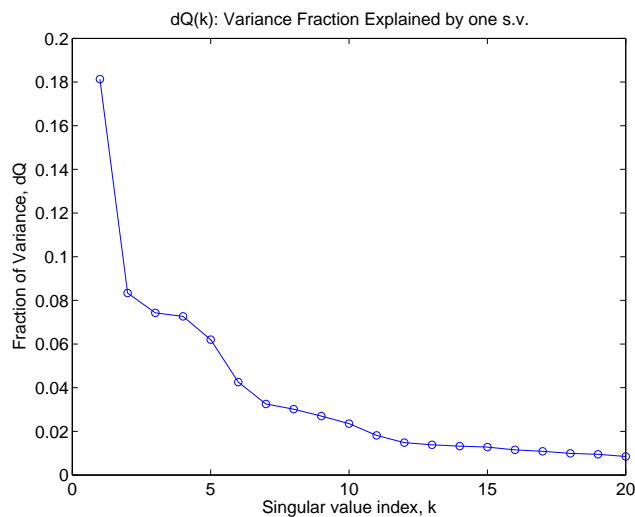


Left Eyes

Right Eyes

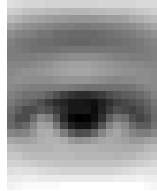
With $V_k \equiv \sum_{j=1}^k \sigma_j^2$,

- $dQ_k \equiv \sigma_k^2 / V_L$ is the fraction of total variance contributed by the k^{th} principal component (left)
- $Q_k \equiv V_k / V_L$ is the fraction of total variance captured by the subspace formed from the first k principal components (right)



Eye Subspace Model

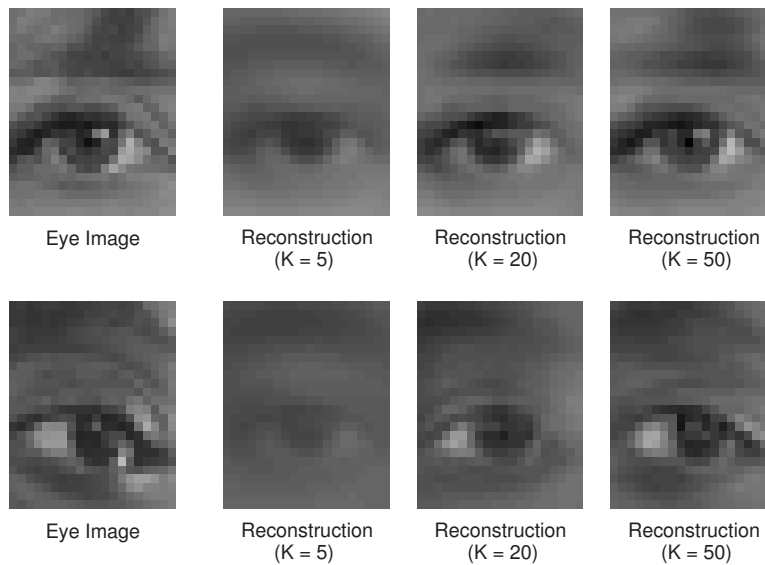
Mean Eye:



Basis Images (1–6, and 10:5:35):



Reconstructions (for $K = 5, 20, 50$):



Generative Eye Model

Generative model, \mathcal{M} , for random eye images:

$$\vec{\mathbf{I}} = \vec{\mathbf{m}} + \left(\sum_{k=1}^K a_k \vec{\mathbf{b}}_k \right) + \vec{\mathbf{e}}$$

where $\vec{\mathbf{m}}$ is the mean eye image, $a_k \sim \mathcal{N}(0, \sigma_k^2)$, σ_k^2 is the sample variance associated with the k^{th} principal direction in the training data, and $\vec{\mathbf{e}} \sim \mathcal{N}(0, \sigma_e^2 \mathbf{I}_{N^2})$ where $\sigma_e^2 = \frac{1}{N^2} \sum_{k=K+1}^{N^2} \sigma_k^2$ is the per pixel out-of-subspace variance.

Random Eye Images:



Random draws from generative model (with $K = 5, 10, 20, 50, 100, 200$)

So the likelihood of an image of an eye given this model \mathcal{M} is

$$p(\vec{\mathbf{I}} | \mathcal{M}) = \left(\prod_{k=1}^K p(a_k | \mathcal{M}) \right) p(\vec{\mathbf{e}} | \mathcal{M})$$

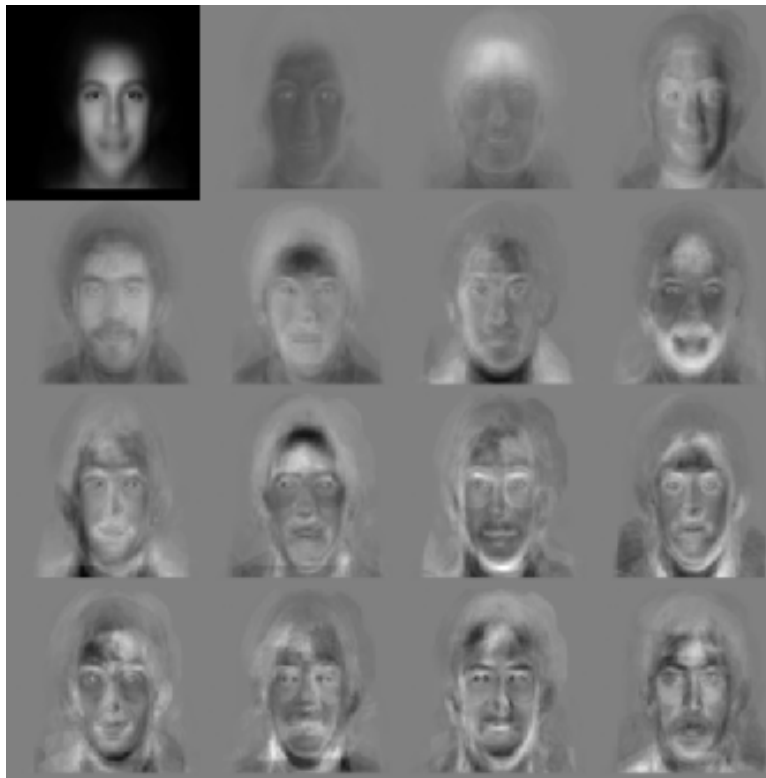
where

$$p(a_k | \mathcal{M}) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{a_k^2}{2\sigma_k^2}}, \quad p(\vec{\mathbf{e}} | \mathcal{M}) = \prod_{j=1}^{N^2} \frac{1}{\sqrt{2\pi}\sigma_e} e^{-\frac{e_j^2}{2\sigma_e^2}}.$$

Face Detection

The wide-spread use of PCA for object recognition began with the work Turk and Pentland (1991) for face detection and recognition.

Shown below is the model learned from a collection of frontal faces, normalized for contrast, scale, and orientation, with the backgrounds removed prior to PCA.

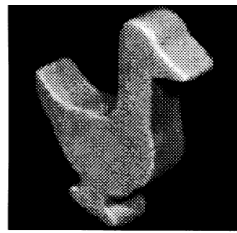


Here are the mean image (upper-left) and the first 15 eigen-images. The first three show strong variations caused by illumination. The next few appear to correspond to the occurrence of certain features (hair, hairline, beard, clothing, etc).

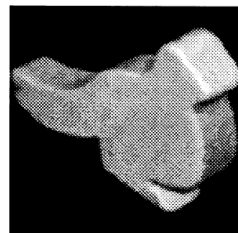
Object Recognition

Murase and Nayar (1995)

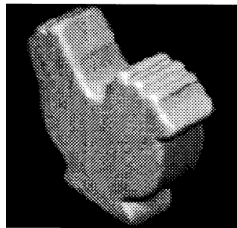
- images of multiple objects, taken from different positions on the viewsphere and with different lighting directions
- each object occupies a manifold in the subspace (as a function of position on the viewsphere and lighting direction)
- recognition: nearest neighbour assuming dense sampling of object pose variations in the training set.



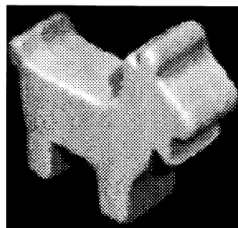
A



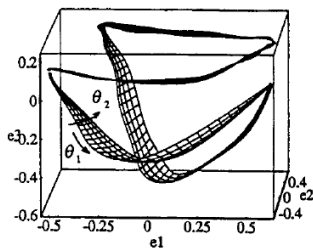
B



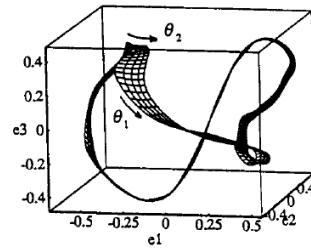
C



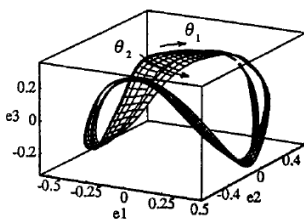
D



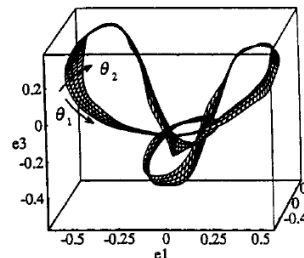
A



B



C



D

Summary

The generative model:

- PCA finds the subspace (of a specified dimension) that maximizes (projected) signal variance.
- A single Gaussian model is naturally associated with a PCA representation. The principal axes are the principal directions of the Gaussian's covariance.
- This can be a simple way to find low-dimensional representations that are effective as prior models and for optimization.

Issues:

- The single Gaussian model is often rather crude. PCA coefficients often exhibit significantly more structure (cf. Murase & Nayar).
- As a result of this unmodelled structure, detectors based on single Gaussian models are often poor. (see the Matlab tutorial `detectEigenEyes.m`)

We'll discuss alternative strategies for detection and recognition later in the course.

Further Readings

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- Chennubhotla, C. and Jepson, A.D. (2001) Sparse PCA (S-PCA): Extracting multi-scale structure from data. *Proc. IEEE ICCV*, Vancouver, pp. 641-647.
- T. Cootes, G. Edwards, and C.J. Taylor, Active Appearance Models, *Proc. IEEE ECCV*, 1998.
- Golub, G. and van Loan, C. (1984) *Matrix Computations*. Johns Hopkins Press, Baltimore.
- Murase, H. and Nayar, S. (1995) Visual learning and recognition of 3D objects from appearance. *Int. J. Computer Vision* 14:5–24.
- M. Black and A. Jepson (1998) EigenTracking: Robust matching and tracking of articulated objects using a view-based representation. *Int. J. Computer Vision* 26(1):63-84.
- Moghaddam, B., Jebara, T. and Pentland, A. (2000) Bayesian face recognition. *Pattern Recognition*, 33(11):1771-1782
- Turk, M. and Pentland, A. (1991) Face recognition using eigenfaces, *J. Cognitive Neuroscience*, 3(1):71–86.