

# CSC 412 (Lecture 4): Undirected Graphical Models

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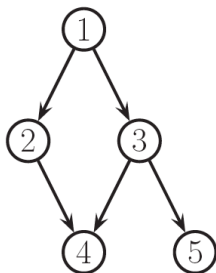
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## Undirected Graphical Models:

- Semantics of the graph: conditional independence
- Parameterization
  - Clique
  - Potentials
  - Gibbs Distribution
  - Partition function
  - Hammersley-Clifford Theorem
- Factor Graphs
- Learning

# Directed Graphical Models

- Represent large joint distribution using "local" relationships specified by the graph
- Each random variable is a node
- The edges specify the statistical dependencies
- We have seen directed acyclic graphs



# Directed Acyclic Graphs

- Represent distribution of the form

$$p(y_1, \dots, y_N) = \prod_i p(y_i | y_{\pi_i})$$

with  $\pi_i$  the parents of the node  $i$

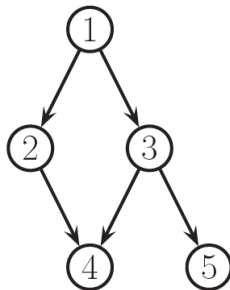
- Factorizes in terms of **local conditional probabilities**
- Each node has to maintain  $p(y_i | y_{\pi_i})$
- Each variable is CI of its non-descendants given its parents

$$\{y_i \perp y_{\tilde{\pi}_i} | y_{\pi_i}\} \quad \forall i$$

with  $y_{\tilde{\pi}_i}$  the nodes before  $y_i$  that are not its parents

- Such an ordering is a "**topological**" ordering (i.e., parents have lower numbers than their children)
- Missing edges imply conditional independence

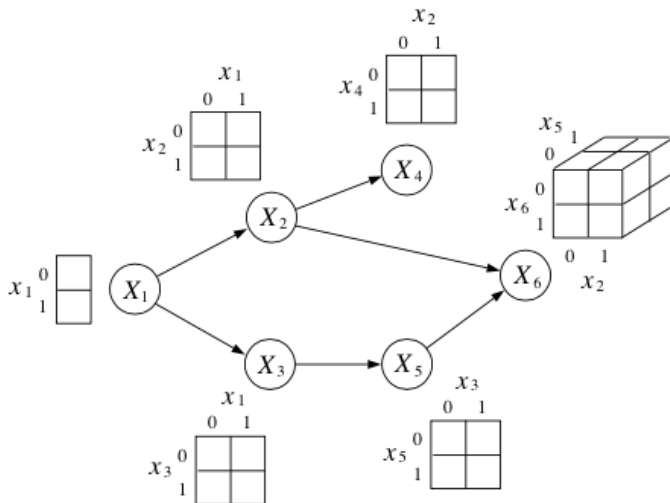
# Example



What's the joint probability distribution?

# Internal Representation

- For discrete variables, each node stores a **conditional probability table** (CPT)



# Are DGM Always Useful?

- Not always clear how to choose the direction for the edges
- Example: Modeling dependencies in an image

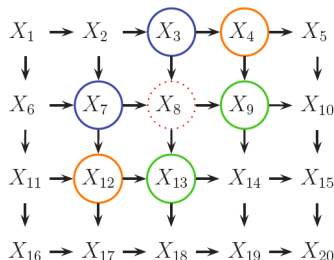
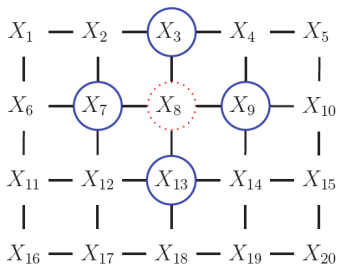


Figure : Causal MRF or a Markov mesh

- Unnatural conditional independence, e.g., see [Markov Blanket](#)  
 $mb(8) = \{3, 7\} \cup \{9, 13\} \cup \{12, 4\}$ , [parents](#), [children](#) and [co-parents](#)
- Alternative: Undirected Graphical models (UGMs)

# Undirected Graphical Models

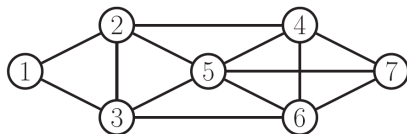
- Also called **Markov random field (MRF)** or **Markov network**
- As in DGM, the **nodes** in the graph represent the variables
- **Edges** represent probabilistic interaction between neighboring variables
- How to parametrize the graph?
  - In DGM we used CPD (conditional probabilities) to represent distribution of a node given others
  - For undirected graphs, we use a more **symmetric** parameterization that captures the affinities between related variables.





# Semantics of the Graph: Conditional Independence

- **Global Markov Property:**  $x_A \perp x_B | x_C$  iff C separates A from B (no path in the graph), e.g.,  $\{1, 2\} \perp \{6, 7\} | \{3, 4, 5\}$



- **Markov Blanket** (local property) is the set of nodes that renders a node  $t$  conditionally independent of all the other nodes in the graph

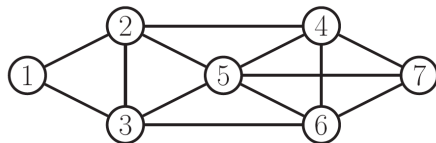
$$t \perp \mathcal{V} \setminus cl(t) | mb(t)$$

where  $cl(t) = mb(t) \cup t$  is the **closure** of node  $t$ . It is the set of neighbors, e.g.,  $mb(5) = \{2, 3, 4, 6, 7\}$ .

- **Pairwise Markov Property**

$$s \perp t | \mathcal{V} \setminus \{s, t\} \iff G_{st} = 0$$

# Dependencies and Examples

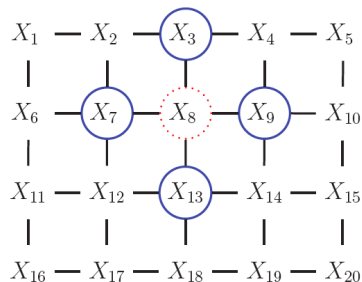


- Pairwise:  $1 \perp 7 | \text{rest}$
- Local:  $1 \perp \text{rest} | 2, 3$
- Global:  $1, 2 \perp 6, 7 | 3, 4, 5$

$$\begin{array}{c} G \rightleftarrows L \rightleftarrows P \\ \leftarrow \text{curved arrow} \rightarrow \\ p(x) > 0 \end{array}$$

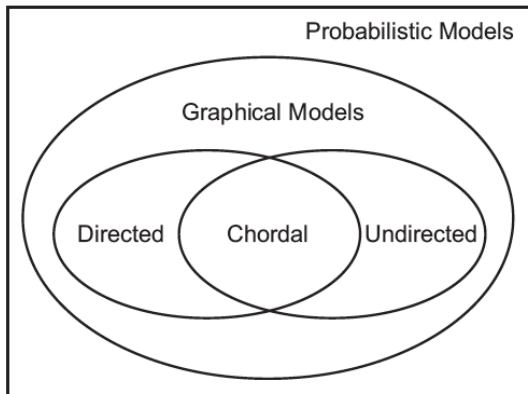
→ See page 119 of Koller and Friedman for a proof

# Image Example



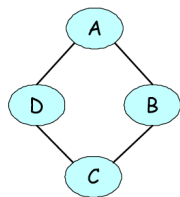
Complete the following statements:

- Pairwise:  $1 \perp 7 | \text{rest}?$ ,  $1 \perp 20 | \text{rest}?$ ,  $1 \perp 2 | \text{rest}?$
- Local:  $1 \perp \text{rest} | ?$ ,  $8 \perp \text{rest} | ?$
- Global:  $1, 2 \perp 15, 20 | ?$

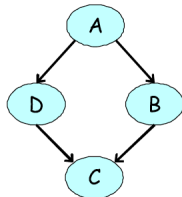


- From Directed to Undirected via [moralization](#)
- From Undirected to Directed via [triangulation](#)
- See (Kohler and Friedman) book if interested

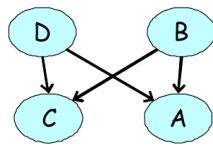
# Not all UGM can be represented as DGM



(a)



(b)

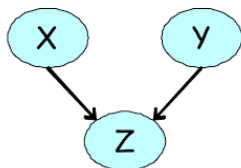


(c)

- Fig. (a) Two independencies:  $(A \perp C | D, B)$  and  $(B \perp D | A, C)$
- Can we encode this with a DGM?
- Fig. (b) First attempt: encodes  $(A \perp C | D, B)$  but it also implies that  $(B \perp D | A)$  but dependent given both  $A, C$
- Fig. (c) Second attempt: encodes  $(A \perp C | D, B)$ , but also implies that  $B$  and  $D$  are marginally independent.

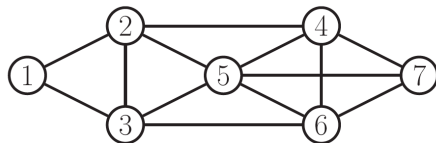
# Not all DGM can be represented as UGM

- Example is the V-structure



- Undirected model fails to capture the marginal independence ( $X \perp Y$ ) that holds in the directed model at the same time as  $\neg(X \perp Y|Z)$

# Cliques



- A **clique** in an undirected graph is a subset of its vertices such that every two vertices in the subset are connected by an edge  
→ i.e., the subgraph induced by the clique is complete
- The **maximal clique** is a clique that cannot be extended by including one more adjacent vertex
- The **maximum clique** is a clique of the largest possible size in a given graph
- What are the maximal cliques? And the maximum clique in the figure?

# Parameterization of an UGM

- $\mathbf{y} = (y_1, \dots, y_m)$  the set of all random variables
- Unlike DGM, since there is no topological ordering associated with an undirected graph, we can't use the chain rule to represent  $p(\mathbf{y})$
- Instead of associating conditional probabilities to each node, we associate **potential functions** or **factors** with each maximal clique in the graph
- For a clique  $c$ , we define the potential function or factor

$$\psi_c(\mathbf{y}_c | \theta_c)$$

to be any non-negative function, with  $\mathbf{y}_c$  the restriction to a subset of variables in  $\mathbf{y}$

- The joint distribution is then proportional to the product of clique potentials
- Any positive distribution whose CI are represented with an UGM can be represented this way (let's see this more formally)



# Factor Parameterization

## Theorem (Hammersley-Clifford)

A positive distribution  $p(\mathbf{y}) > 0$  satisfies the CI properties of an undirected graph  $G$  iff  $p$  can be represented as a product of factors, one per maximal clique, i.e.,

$$p(\mathbf{y}|\theta) = \frac{1}{Z(\theta)} \prod_{c \in \mathcal{C}} \psi_c(\mathbf{y}_c|\theta_c)$$

with  $\mathcal{C}$  the set of all (maximal) cliques of  $G$ , and  $Z(\theta)$  the **partition function** defined as

$$Z(\theta) = \sum_{\mathbf{y}} \prod_{c \in \mathcal{C}} \psi_c(\mathbf{y}_c|\theta_c)$$

## Proof.

Can be found in (Koller and Friedman book) □

We need the **partition function** as the potentials are not conditional distributions. In DGMs we don't need it

# The partition function

The joint distribution is

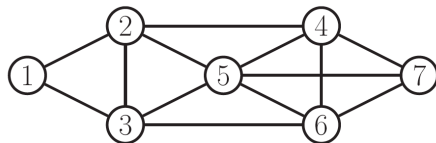
$$p(\mathbf{y}|\theta) = \frac{1}{Z(\theta)} \prod_{c \in \mathcal{C}} \psi_c(\mathbf{y}_c|\theta_c)$$

with the partition function

$$Z(\theta) = \sum_{\mathbf{y}} \prod_{c \in \mathcal{C}} \psi_c(\mathbf{y}_c|\theta_c)$$

- This is the hardest part of learning and inference. Why?
- **Factored structure** of the distribution makes it possible to more efficiently do the sums/integrals needed to compute it.

# Example



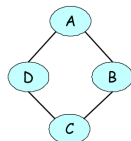
$$p(\mathbf{y}) \propto \psi_{1,2,3}(y_1, y_2, y_3) \psi_{2,3,5}(y_2, y_3, y_5) \psi_{2,4,5}(y_2, y_4, y_5) \\ \psi_{3,5,6}(y_3, y_5, y_6) \psi_{4,5,6,7}(y_4, y_5, y_6, y_7)$$

- Is this representation unique?
- What if I want a pairwise MRF?

# Representing Potentials

- If the variables are discrete, we can represent the potential or energy functions as tables of (non-negative) numbers

$$p(A, B, C, D) = \frac{1}{Z} \psi_{a,b}(A, B) \psi_{b,c}(B, C) \psi_{c,d}(C, D) \psi_{a,d}(A, D)$$



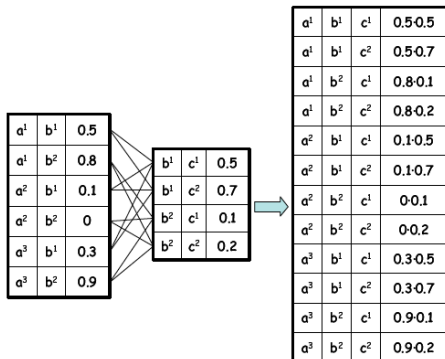
$\phi_1[A, B]$			$\phi_2[B, C]$			$\phi_3[C, D]$			$\phi_4[D, A]$		
$a^0$	$b^0$	30	$b^0$	$c^0$	100	$c^0$	$d^0$	1	$d^0$	$a^0$	100
$a^0$	$b^1$	5	$b^0$	$c^1$	1	$c^0$	$d^1$	100	$d^0$	$a^1$	1
$a^1$	$b^0$	1	$b^1$	$c^0$	1	$c^1$	$d^0$	100	$d^1$	$a^0$	1
$a^1$	$b^1$	10	$b^1$	$c^1$	100	$c^1$	$d^1$	1	$d^1$	$a^1$	100

- The potentials are NOT probabilities
- They represent compatibility between the different assignments

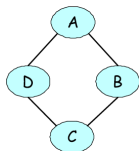
# Factor product

- Given 3 disjoint set of variables  $\mathbf{X}, \mathbf{Y}, \mathbf{Z}$ , and factors  $\psi_1(\mathbf{X}, \mathbf{Y}), \psi_2(\mathbf{Y}, \mathbf{Z})$ , the **factor product** is defined as

$$\psi_{x,y,z}(\mathbf{X}, \mathbf{Y}, \mathbf{Z}) = \psi_{x,y}(\mathbf{X}, \mathbf{Y})\phi_{y,z}(\mathbf{Y}, \mathbf{Z})$$



# Query about probabilities: marginalization

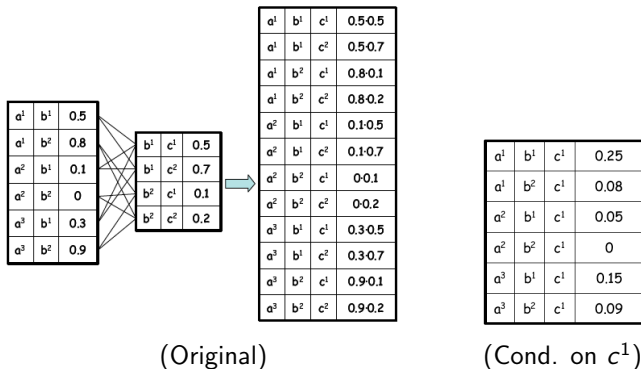


$\phi_1[A, B]$			$\phi_2[B, C]$			$\phi_3[C, D]$			$\phi_4[D, A]$		
$a^0$	$b^0$	30	$b^0$	$c^0$	100	$c^0$	$d^0$	1	$d^0$	$a^0$	100
$a^0$	$b^1$	5	$b^0$	$c^1$	1	$c^0$	$d^1$	100	$d^0$	$a^1$	1
$a^1$	$b^0$	1	$b^1$	$c^0$	1	$c^1$	$d^0$	100	$d^1$	$a^0$	1
$a^1$	$b^1$	10	$b^1$	$c^1$	100	$c^1$	$d^1$	1	$d^1$	$a^1$	100

Assignment				Unnormalized	Normalized
$a^0$	$b^0$	$c^0$	$d^0$	300000	0.04
$a^0$	$b^0$	$c^0$	$d^1$	300000	0.04
$a^0$	$b^0$	$c^1$	$d^0$	300000	0.04
$a^0$	$b^0$	$c^1$	$d^1$	30	$4.1 \cdot 10^{-6}$
$a^0$	$b^1$	$c^0$	$d^0$	500	$6.9 \cdot 10^{-5}$
$a^0$	$b^1$	$c^0$	$d^1$	500	$6.9 \cdot 10^{-5}$
$a^0$	$b^1$	$c^1$	$d^0$	5000000	0.69
$a^0$	$b^1$	$c^1$	$d^1$	500	$6.9 \cdot 10^{-5}$
$a^1$	$b^0$	$c^0$	$d^0$	100	$1.4 \cdot 10^{-5}$
$a^1$	$b^0$	$c^0$	$d^1$	1000000	0.14
$a^1$	$b^0$	$c^1$	$d^0$	100	$1.4 \cdot 10^{-5}$
$a^1$	$b^0$	$c^1$	$d^1$	100	$1.4 \cdot 10^{-5}$
$a^1$	$b^1$	$c^0$	$d^0$	10	$1.4 \cdot 10^{-6}$
$a^1$	$b^1$	$c^0$	$d^1$	100000	0.014
$a^1$	$b^1$	$c^1$	$d^0$	100000	0.014
$a^1$	$b^1$	$c^1$	$d^1$	100000	0.014

- What's the  $p(b^0)$ ? Marginalize the other variables!

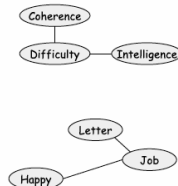
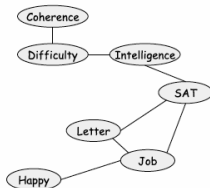
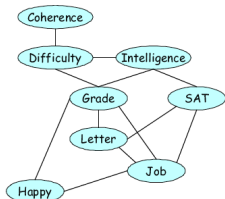
# Query about probabilities: conditioning



- **Conditioning** on an assignment  $\mathbf{u}$  to a subset of variables  $\mathbf{U}$  can be done by
  - 1 **Eliminating** all entries that are inconsistent with the assignment
  - 2 **Re-normalizing** the remaining entries so that they sum to 1

# Reduced Markov Networks

- Let  $\mathcal{H}$  be a Markov network over  $\mathbf{X}$  and let  $\mathbf{U} = u$  be the context. The reduced network  $\mathcal{H}[u]$  is a Markov network over the nodes  $\mathbf{W} = \mathbf{X} - \mathbf{U}$  where we have an edge between  $X$  and  $Y$  if there is an edge between them in  $\mathcal{H}$



- If  $\mathbf{U} = \text{Grade}$ ?
- If  $\mathbf{U} = \{\text{Grade}, \text{SAT}\}$ ?



- The **Gibbs Distribution** is defined as

$$p(\mathbf{y}|\theta) = \frac{1}{Z(\theta)} \exp\left(-\sum_c E(\mathbf{y}_c|\theta_c)\right)$$

where  $E(\mathbf{y}_c) > 0$  is the energy associated with the variables in clique  $c$

- We can convert this distribution to a UGM by

$$\psi(\mathbf{y}_c|\theta_c) = \exp(-E(\mathbf{y}_c|\theta_c))$$

- High probability states correspond to low energy configurations.
- These models are named **energy based models**

# Log Linear Models

- Represent the log potentials as a linear function of the parameters

$$\log \psi_c(\mathbf{y}_c) = \phi_c(\mathbf{y}_c)^T \theta_c$$

- The log probability is then

$$\log p(\mathbf{y}|\theta) = \sum_c \phi_c(\mathbf{y}_c)^T \theta_c - \log Z(\theta)$$

- This is called **log linear model**
- Example: we can represent tabular potentials

$$\psi(y_s = j, y_t = k) = \exp([\theta_{st}^T \phi_{st}]_{jk}) = \exp(\theta_{st}(j, k))$$

with  $\phi_{st}(y_s, y_t) = [\cdots, I(y_s = j, y_t = k), \cdots]$  and  $I$  the indicator function

## Example: Ising model

- Captures the energy of a set of interacting atoms.
- $y_i \in \{-1, +1\}$  represents direction of the atom spin.
- The graph is a **2D or 3D lattice**, and the energy of the edges is symmetric

$$\psi_{st}(y_s, y_t) = \begin{pmatrix} e^{w_{st}} & e^{-w_{st}} \\ e^{-w_{st}} & e^{w_{st}} \end{pmatrix}$$

with  $w_{st}$  the coupling strength between two nodes. If not connected  $w_{st} = 0$

- Often we assume all edges have the same strength, i.e.,  $w_{st} = J \neq 0$
- If all weights positive, then neighboring spins likely same spin (ferromagnets, associative Markov network)
- If weights are very strong, then two models, all +1 and all -1
- If weights negative, then anti-ferromagnets. Not all the constraints can be satisfied, and the prob. distribution has multiple modes
- Also individual node potentials that encode the bias of the individual atoms (i.e., external field)

## More on Ising Models

- Captures the energy of a set of interacting atoms.
- $y_i \in \{-1, +1\}$  represents direction of the atom spin.
- The **energy** associated is

$$P(\mathbf{y}) = \frac{1}{Z} \exp \left( \sum_{i,j} \frac{1}{2} w_{i,j} y_i y_j + \sum_i b_i y_i \right) = \frac{1}{Z} \exp \left( \frac{1}{2} \mathbf{y}^T \mathbf{W} \mathbf{y} + \mathbf{b}^T \mathbf{y} \right)$$

- The energy can be written as

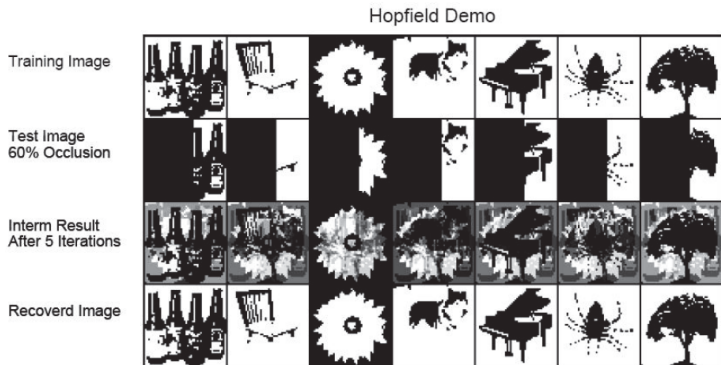
$$E(\mathbf{y}) = -\frac{1}{2} (\mathbf{y} - \boldsymbol{\mu})^T \mathbf{W} (\mathbf{y} - \boldsymbol{\mu}) + c$$

with  $\boldsymbol{\mu} = -\mathbf{W}^{-1} \mathbf{u}$ ,  $c = \frac{1}{2} \boldsymbol{\mu}^T \mathbf{W} \boldsymbol{\mu}$

- Looks like a Gaussian... but is it?
- Often modulated by a temperature  $p(\mathbf{y}) = \frac{1}{Z} \exp(-E(\mathbf{y})/T)$
- $T$  small makes distribution picky

# Example: Hopfield networks

- A Hopfield network is a fully connected Ising model with a symmetric weight matrix  $\mathbf{W} = \mathbf{W}^T$
- The main application of Hopfield networks is as an **associative memory**

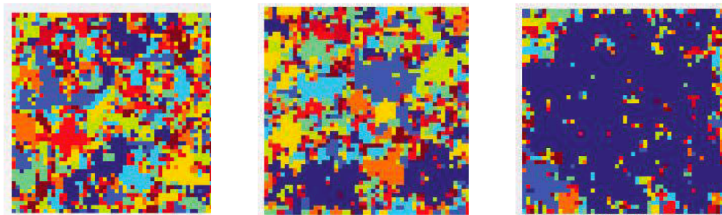


# Example: Potts Model

- Multiple discrete states  $y_i \in \{1, 2, \dots, K\}$
- Common to use

$$\psi_{st}(y_s, y_t) = \begin{pmatrix} e^J & 0 & 0 \\ 0 & e^J & 0 \\ 0 & 0 & e^J \end{pmatrix}$$

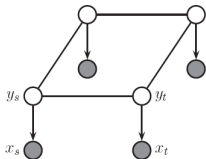
- If  $J > 0$  neighbors encourage to have the same label
- Phase transition: change of behavior,  $J = 1.44$  in example



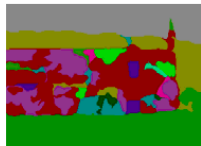
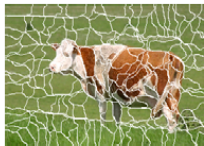
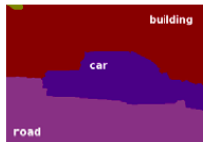
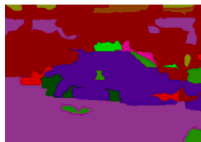
**Figure :** Sample from a 10-state Potts model of size  $128 \times 128$  for (a)  $J = 1.42$ , (b)  $J = 1.44$ , (c)  $J = 1.46$

# More on Potts

- Used in image segmentation: neighboring pixels are likely to have the same class label



$$p(\mathbf{y}|\mathbf{x}, \theta) = \frac{1}{Z} \prod_i \psi_i(y_i|\mathbf{x}) \prod_{i,j} \psi_{i,j}(y_i, y_j)$$



# Example: Gaussian MRF

- Is a pairwise MRF

$$p(\mathbf{y}|\theta) \propto \prod_{s \sim t} \psi_{st}(y_s, y_t) \prod_t \psi_t(y_t)$$

$$\psi_{st}(y_s, y_t) = \exp\left(-\frac{1}{2}y_s \Lambda_{st} y_t\right)$$

$$\psi_t(y_t) = \exp\left(-\frac{1}{2}\Lambda_{tt}y_t^2 + \eta_t y_t\right)$$

- The joint distribution is then

$$p(\mathbf{y}|\theta) \propto \exp\left[\eta^T \mathbf{y} - \frac{1}{2}\mathbf{y}^T \Lambda \mathbf{y}\right]$$

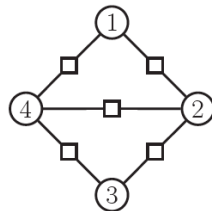
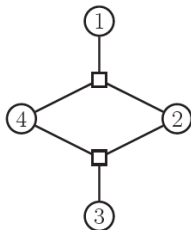
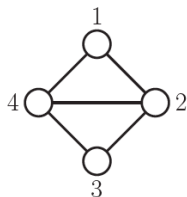
- This is a multivariate Gaussian with  $\Lambda = \Sigma^{-1}$  and  $\eta = \Lambda \mu$
- If  $\Lambda_{st} = 0$  (structural zero), then no pairwise connection and by factorization theorem

$$y_s \perp y_t | \mathbf{y}_{-(st)} \iff \Lambda_{st} = 0$$

- UGM are sparse precision matrices. Used for [structured learning](#)



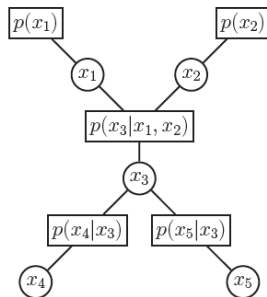
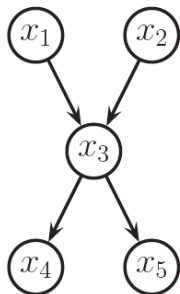
# Factor Graphs



- A **factor graph** is a graphical model representation that unifies directed and undirected models
- It is an undirected bipartite graph with two kinds of nodes.
  - Round nodes represent **variables**,
  - Square nodes represent **factors**and there is an **edge** from each variable to every factor that mentions it.
- Represents the distribution more uniquely than a graphical model

# Factor Graphs for Directed Models

- One factor per CPD (conditional distribution) and connect the factor to all the variables that use the CPD



# Learning using Gradient methods

- MRF in log-linear form

$$p(\mathbf{y}|\theta) = \frac{1}{Z(\theta)} \exp \left( \sum_c \theta_c^T \phi_c(\mathbf{y}_c) \right)$$

- Given training examples  $\mathbf{y}^{(i)}$ , the scaled log likelihood is

$$\ell(\theta) = -\frac{1}{N} \sum_i \log p(\mathbf{y}^{(i)}|\theta) = \frac{1}{N} \sum_i \left[ -\sum_c \theta_c^T \phi_c(\mathbf{y}_c^{(i)}) + \log Z^{(i)}(\theta) \right]$$

- Since MRFs are in the **exponential family**, this function is convex in  $\theta$
- We can find the global maximum, e.g., via gradient descent

$$\frac{\partial \ell}{\partial \theta_c} = \frac{1}{N} \sum_i \left[ -\phi_c(\mathbf{y}_c^{(i)}) + \frac{\partial}{\partial \theta_c} \log Z^{(i)}(\theta) \right]$$

- The first term is constant for each iteration of gradient descent, it is called the **empirical means**

# Moment Matching

$$\frac{\partial \ell}{\partial \theta_c} = \frac{1}{N} \sum_i \left[ -\phi_c(\mathbf{y}_c^{(i)}) + \frac{\partial}{\partial \theta_c} \log Z^{(i)}(\theta) \right]$$

- The derivative of the log partition function w.r.t.  $\theta_c$  is the expectation of the  $c$ 'th feature under the model

$$\frac{\partial \log Z(\theta)}{\partial \theta_c} = \sum_{\mathbf{y}} \phi_c(\mathbf{y}) p(\mathbf{y}|\theta) = E[\phi_c(\mathbf{y})]$$

Thus the gradient of the log likelihood is

$$\frac{\partial \ell}{\partial \theta_c} = \left[ -\frac{1}{N} \sum_i \phi_c(\mathbf{y}_c^{(i)}) \right] + E[\phi_c(\mathbf{y})]$$

- The second term is the **contrastive term** or **unclamped term** and requires **inference** in the model (it has to be done for each step in gradient descent)
- Dif. of the empirical distrib. and model's expectation of the feature vector

$$\frac{\partial \ell}{\partial \theta_c} = -E_{p_{emp}}[\phi_c(\mathbf{y})] + E_{p(\cdot|\theta)}[\phi_c(\mathbf{y})]$$

- At the optimum the moments are matched (i.e., **moment matching**)

# Approximated Methods

- In UGM, no closed form solution to the ML estimate of the parameters, need to do **gradient-based optimization**
- Computing each gradient step **requires inference** → very expensive (NP-hard in general)
- Many approximations exist: stochastic approaches, pseudo likelihood, etc