Perturbation, Optimization & Statistics workshop

Probabilistic inference by randomly perturbing max-solvers

Tamir Hazan
University of Haifa

Complex structures dominate machine learning applications:

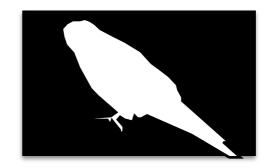
- Complex structures dominate machine learning applications:
 - Computer vision



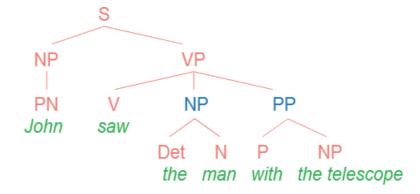


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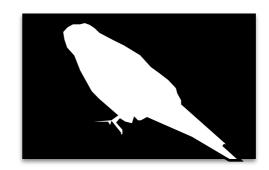


- Natural language processing

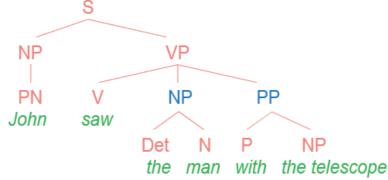


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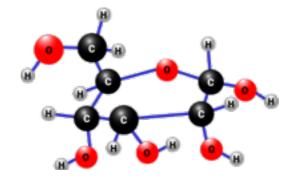




- Natural language processing



Computational biology



- and more..

Outline

- Random perturbation why and how?
 - Sampling likely structures as fast as finding the most likely one.

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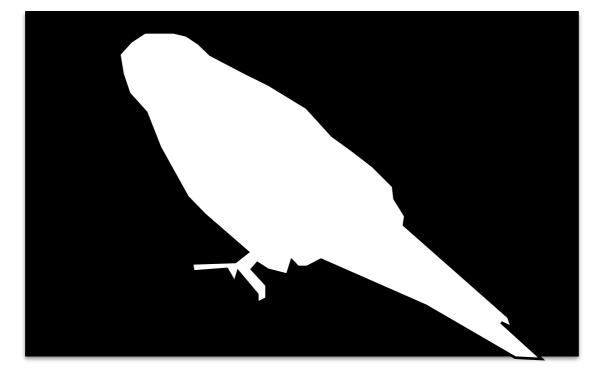
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- Application: interactive annotation.
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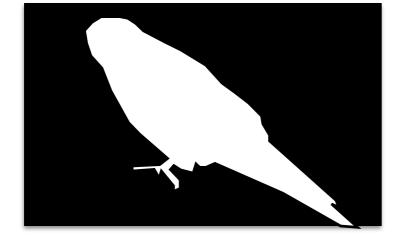
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high score



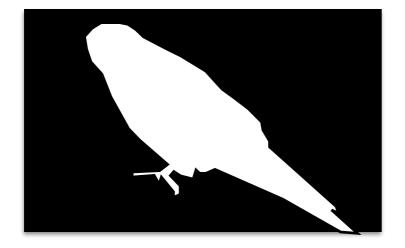
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$$\theta(y_1, ..., y_n) = \sum_{i \in V} \theta_i(y_i) + \sum_{i,j \in E} \theta_{i,j}(y_i, y_j)$$

high score



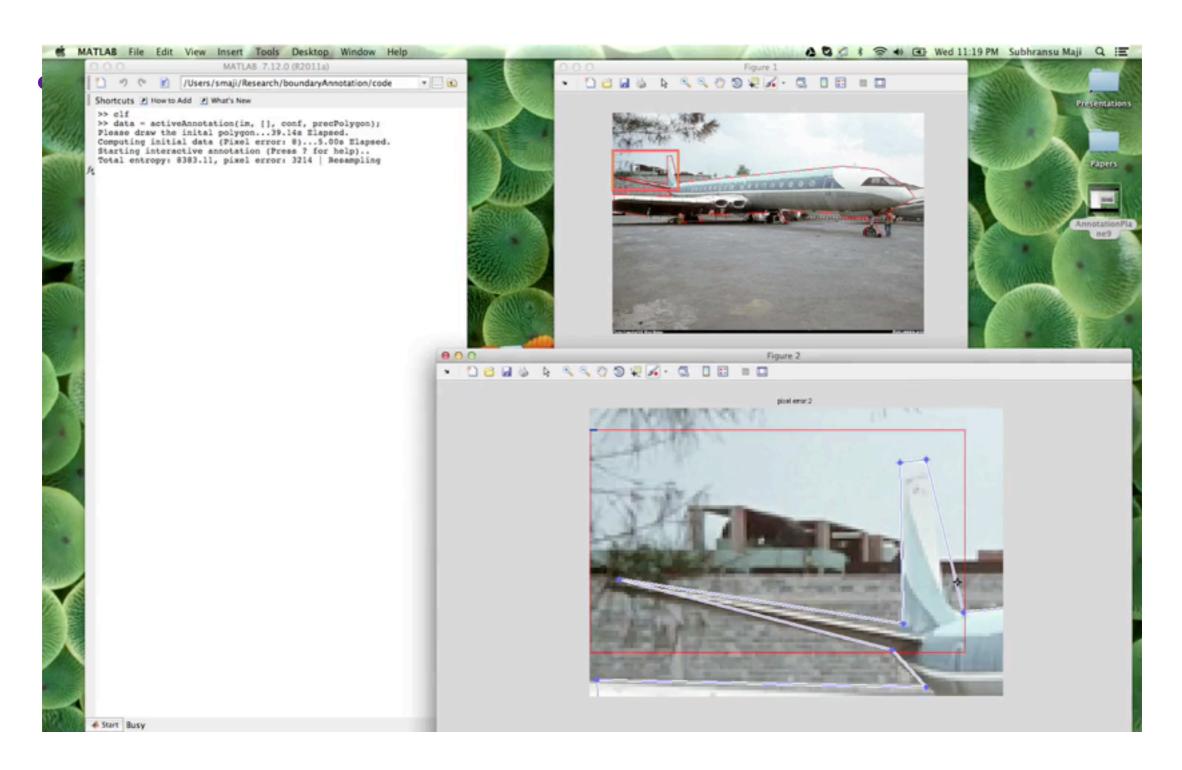
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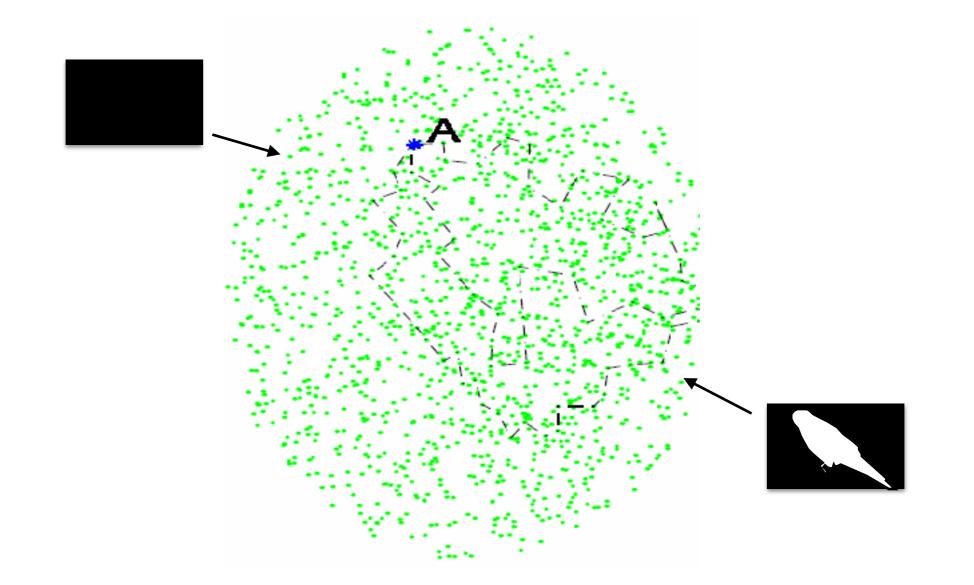
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 For machine learning we need to efficiently infer from distributions over complex structures.



$$p(y_1, ..., y_n) = \frac{1}{Z} \exp\left(\sum_{i} \theta_i(y_i) + \sum_{i,j} \theta_{i,j}(y_i, y_j)\right)$$

MCMC samplers:



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- MCMC samplers:
 - Gibbs sampling, Metropolis-Hastings, Swendsen-Wang
- Many efficient sampling algorithms for special cases:
 - Counting bi-partite matchings in planar graphs (Kasteleyn 61)
 - Ising models (Jerrum 93)
 - Approximating the permanent (Jerrum 04)
 - Many others...

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- Gibbs distribution has a significant impact on statistics and computer science
 - Efficient sampling in Ising models (Jerrum 93)
 - Attractive pairwise potentials

$$\theta_{i,j}(y_i, y_j) = \begin{cases} w_{i,j} & \text{if } y_i = y_j \\ -w_{i,j} & \text{otherwise} \end{cases}$$

$$w_{i,j} \ge 0$$

No data terms

$$\theta_i(y_i) = 0$$

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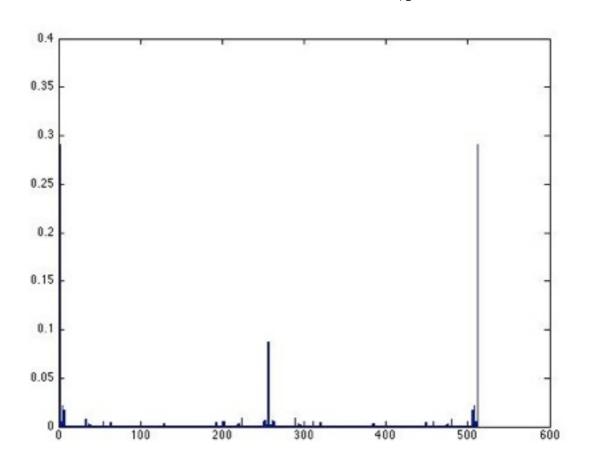
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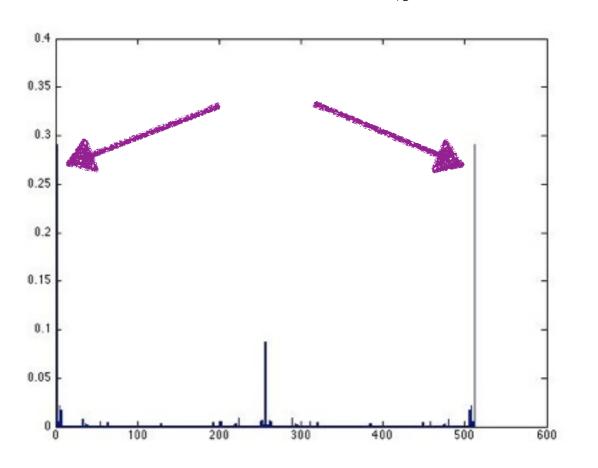
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 Nicely behaved distribution that is centered around the (I,...,I) or (0,...,0)

 Sampling from the Gibbs distribution is provably hard in Al applications (Goldberg 05, Jerrum 93)



• x_i RGB color of pixel i

$$\theta_i(y_i) = \log p(y_i|x_i)$$

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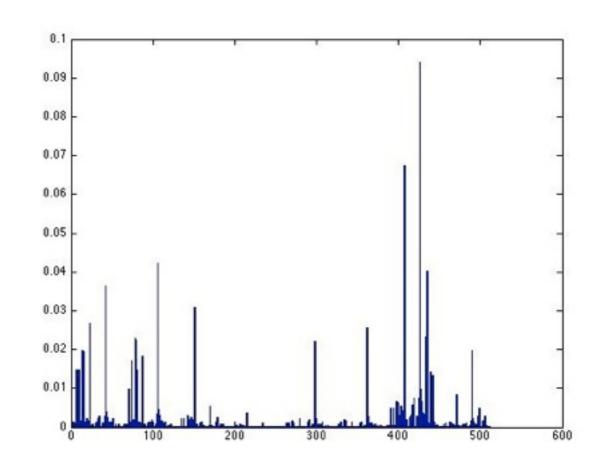
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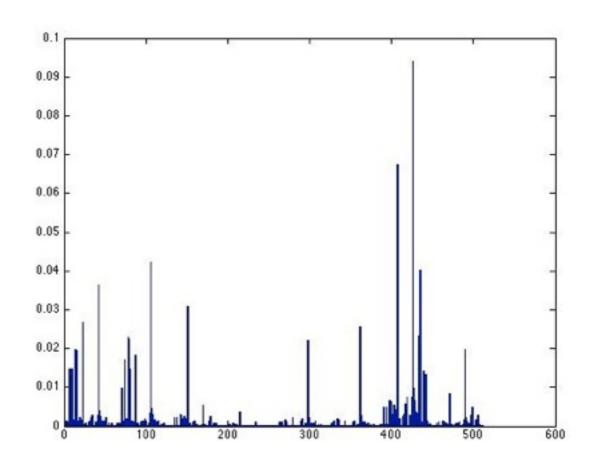
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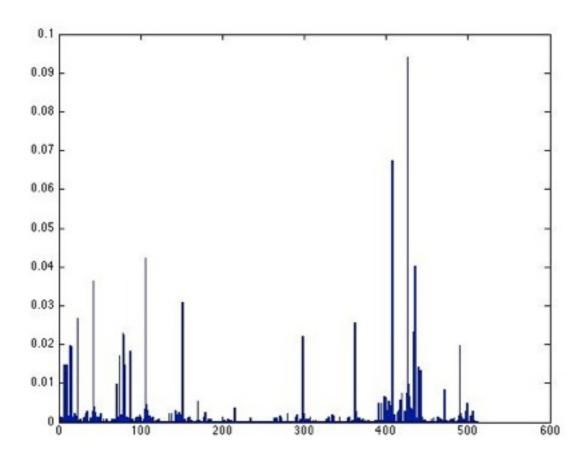
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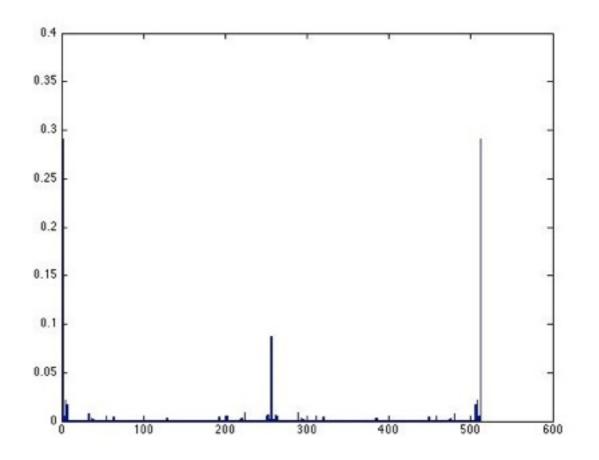
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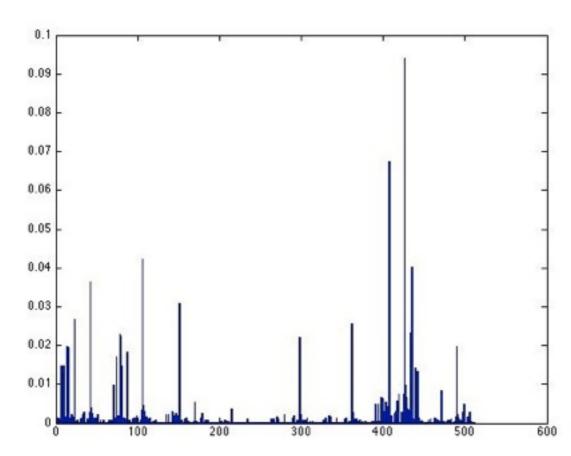
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 Recall: sampling from the Gibbs distribution is easy in Ising models (Jerrum 93)



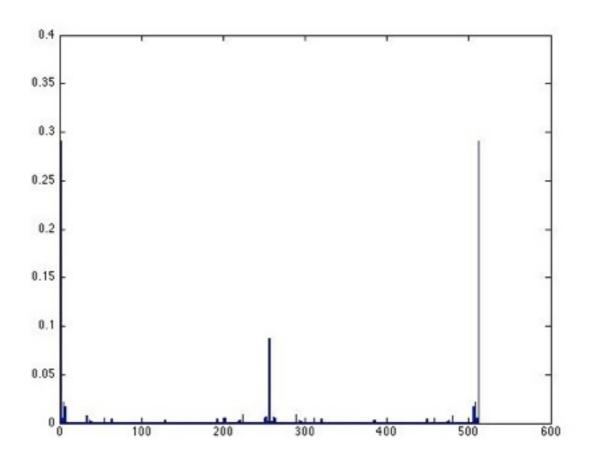
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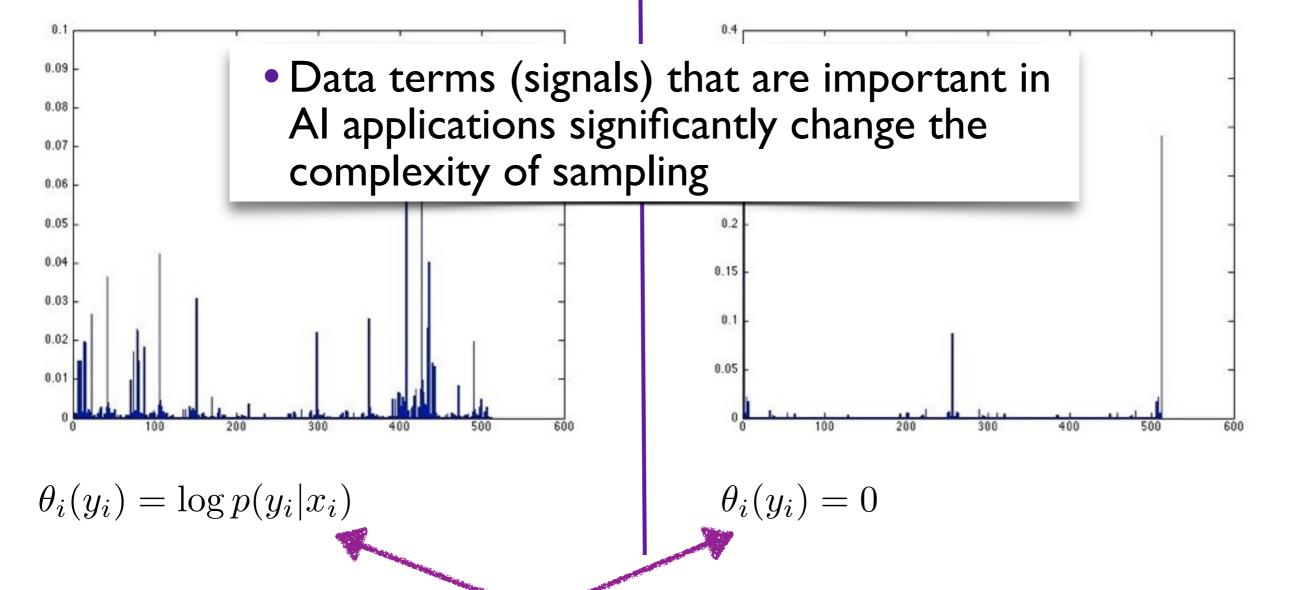
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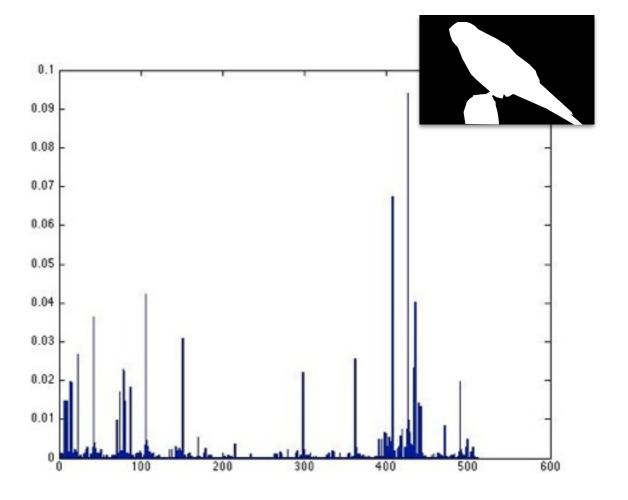
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 Instead of sampling, it may be significantly faster to find the most likely structure

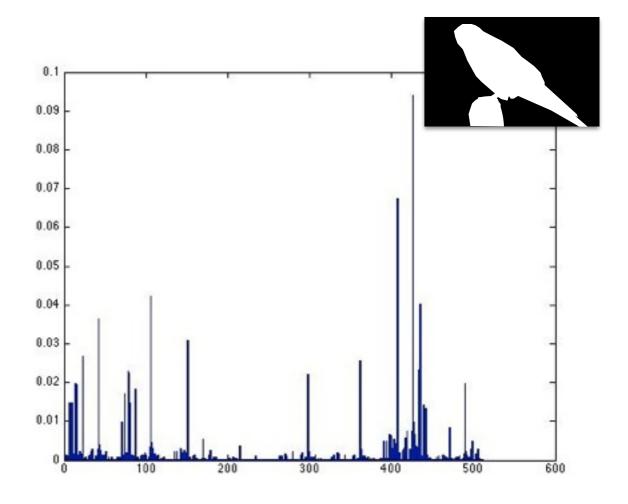
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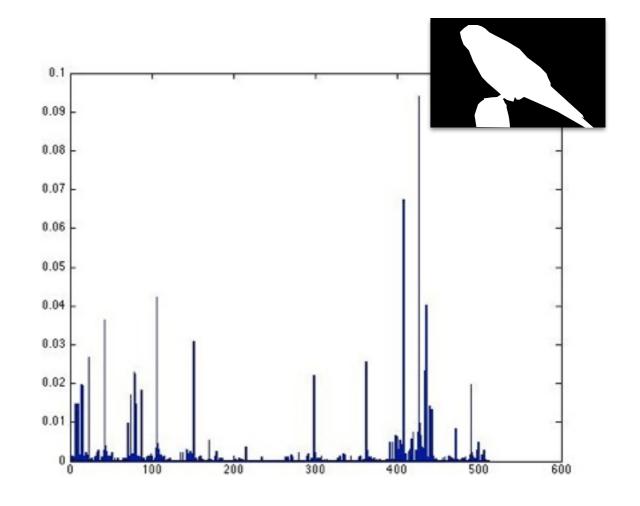
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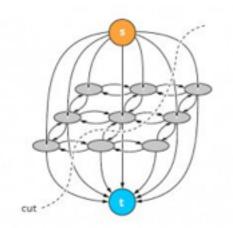


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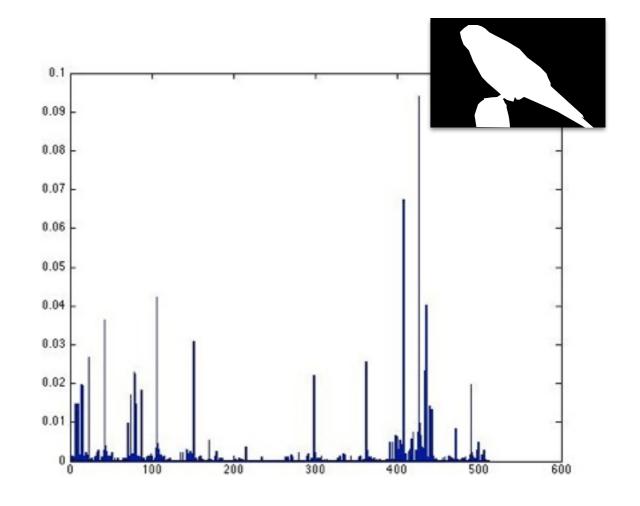
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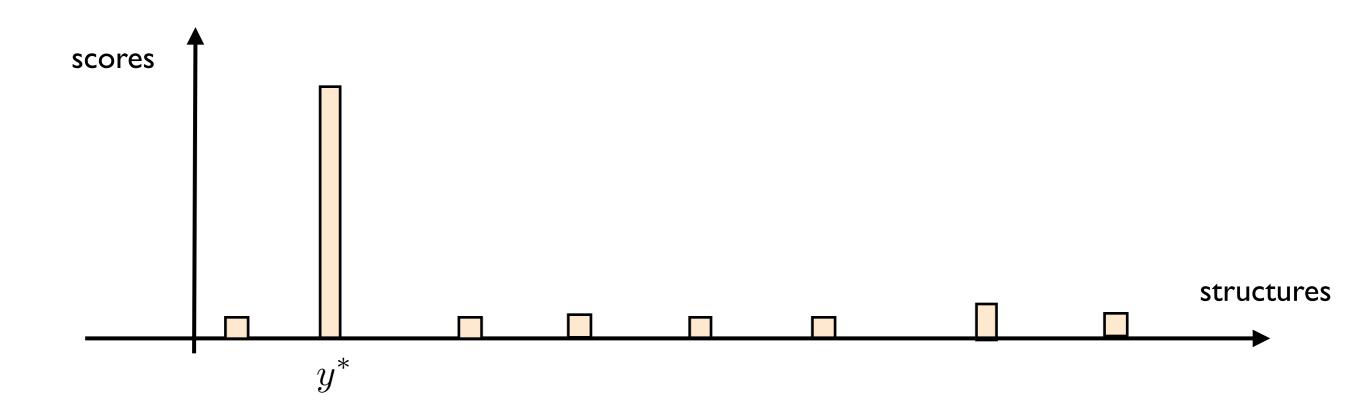


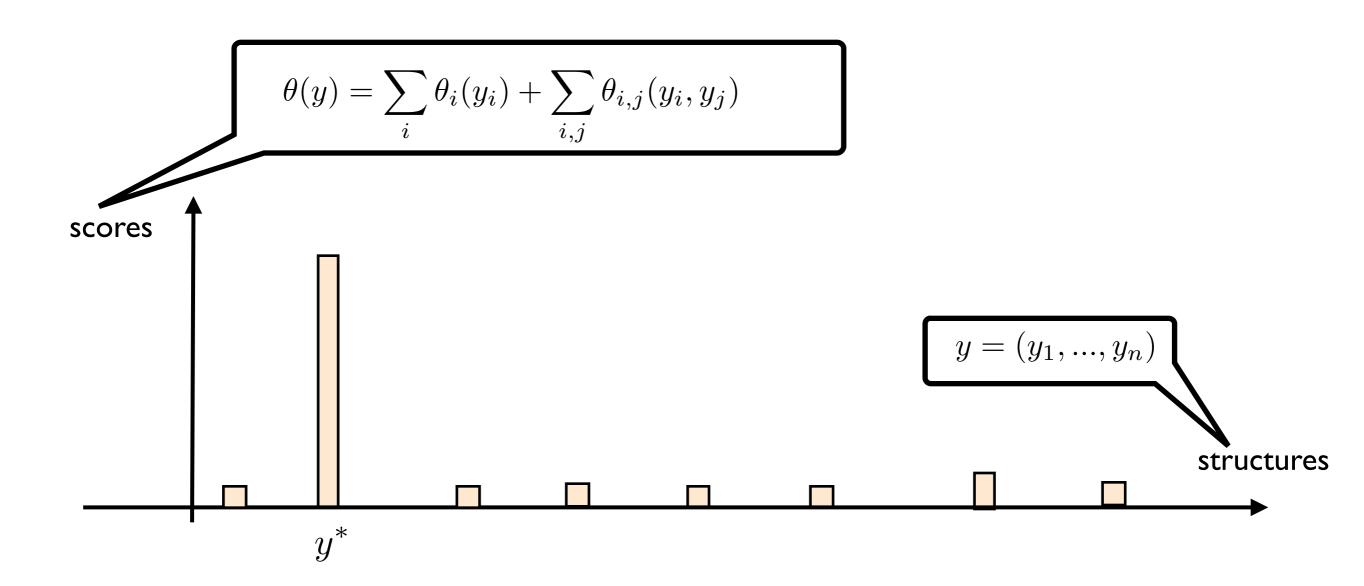
$$y^* = \arg \max_{y_1,...,y_n} \sum_{i} \theta_i(y_i) + \sum_{i,j} \theta_{i,j}(y_i, y_j)$$

- Maximum a-posterior (MAP) inference.
- Many efficient optimization algorithms for special cases:
 - Beliefs propagation: trees (Pearl 88), perfect graphs (Jebara 10),
 - Graph-cuts for image segmentation
 - branch and bound (Rother 09), branch and cut (Gurobi)
 - Linear programming relaxations (Schlesinger 76, Wainwright 05, Kolmogorov 06, Werner 07, Sontag 08, Hazan 10, Batra 10, Nowozin 10, Pletscher 12, Kappes 13, Savchynskyy 13, Tarlow 13, Kohli 13, Jancsary 13, Schwing 13)
 - CKY for parsing
 - Many others...

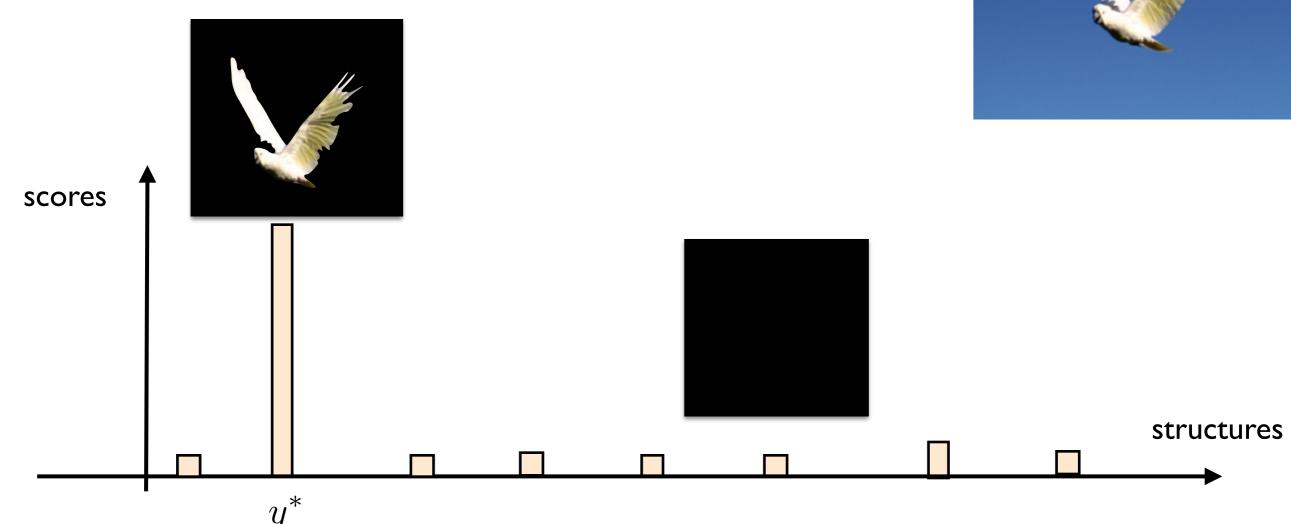
The challenge

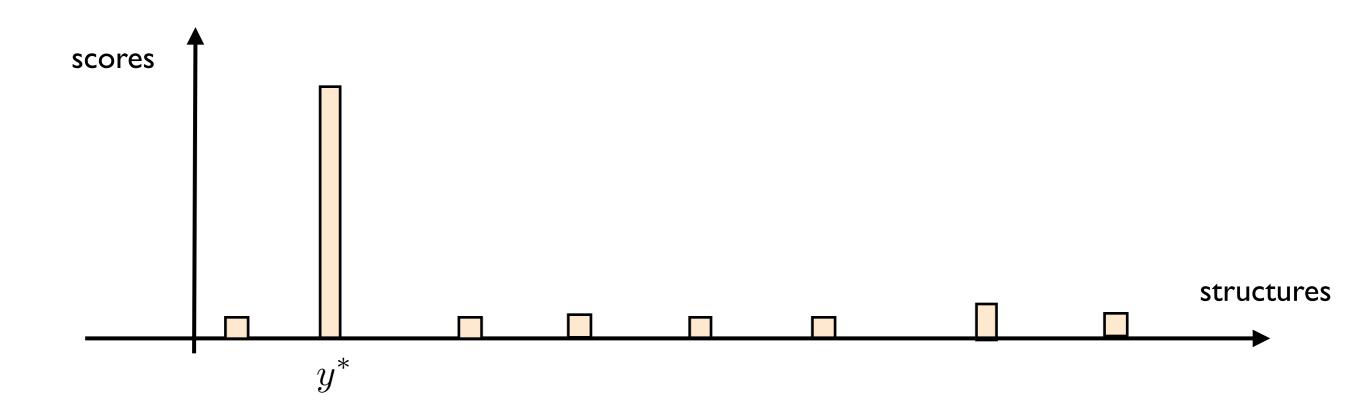
Sampling from the likely high dimensional structures (with millions of variables, e.g., image segmentation with 12 million pixels) as efficient as optimizing





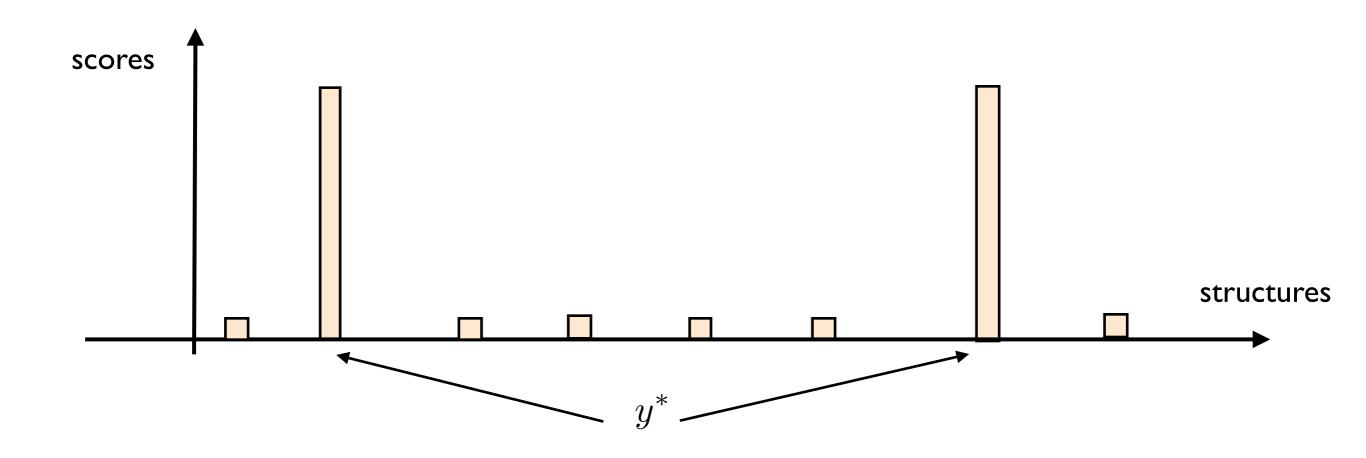




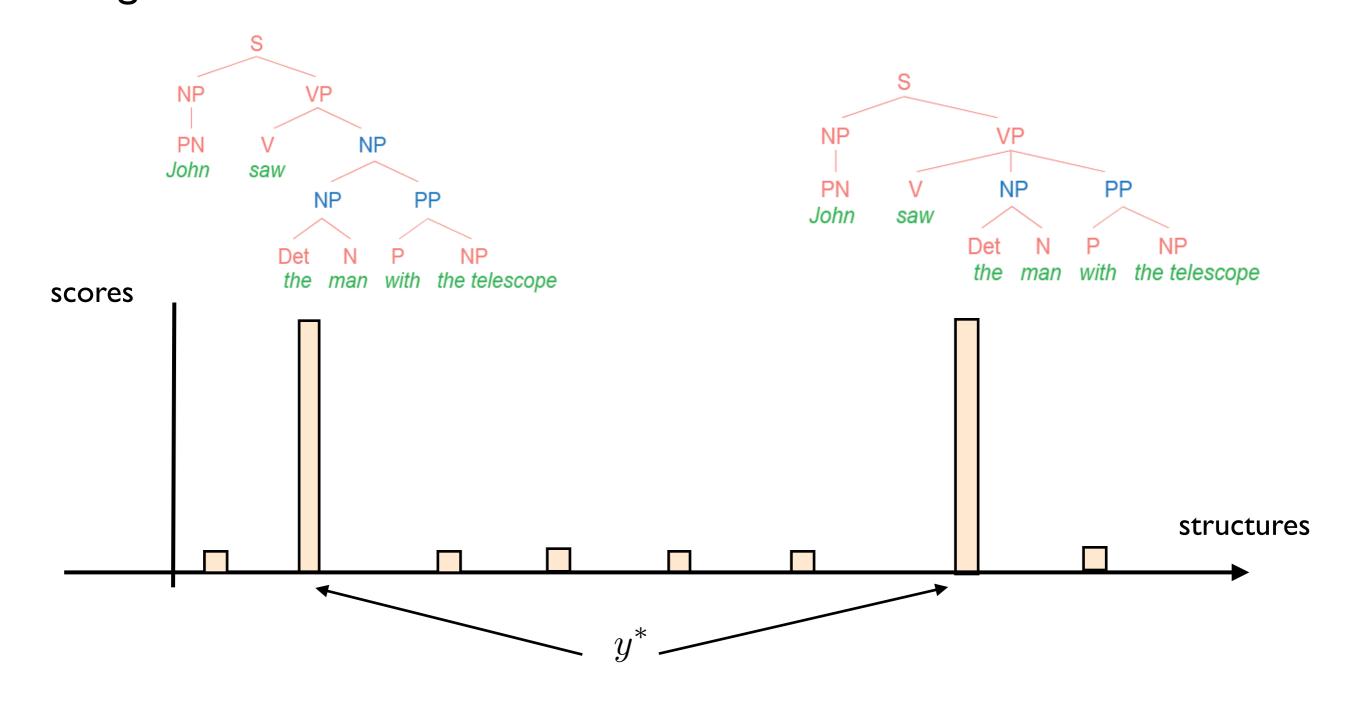


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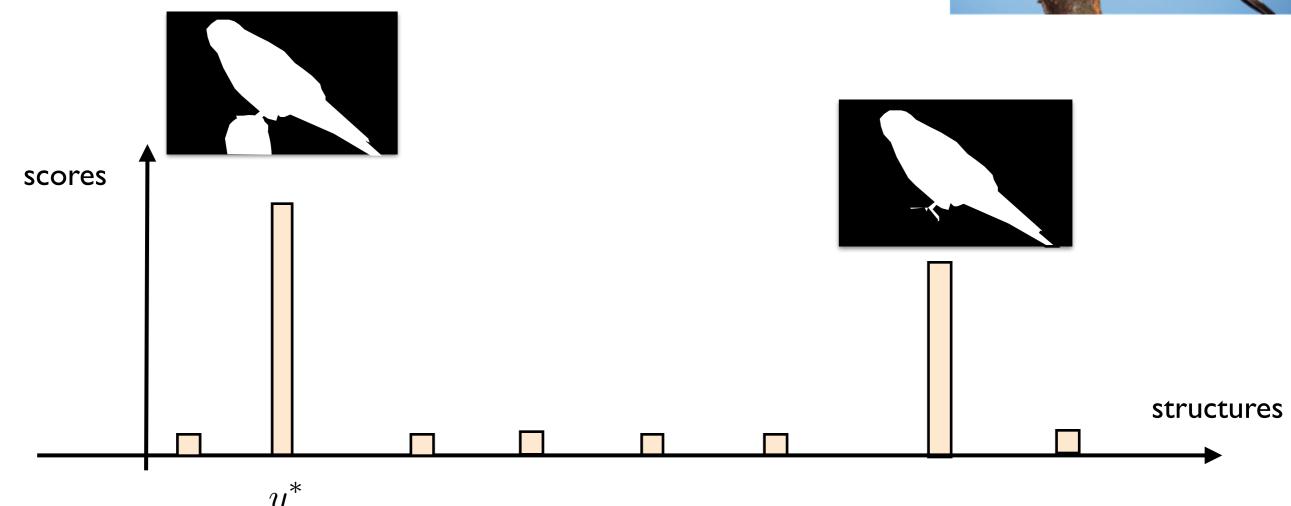


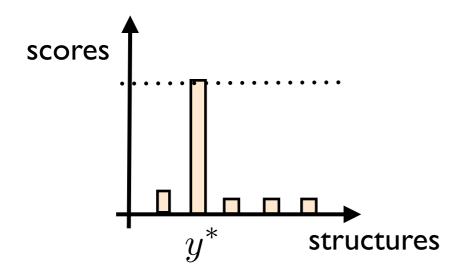
 The maximizing structure is not robust in case of computationally limited models



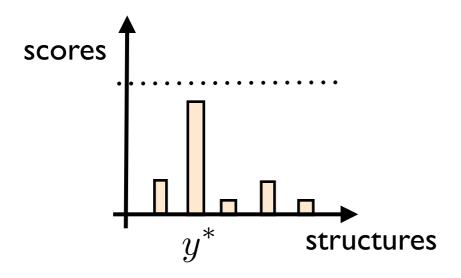
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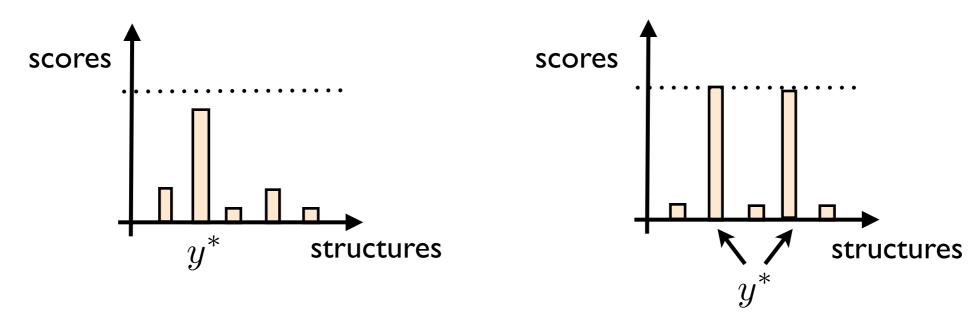




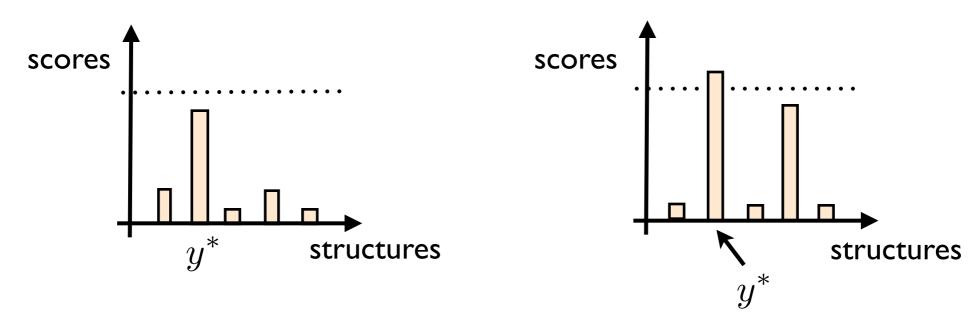
- Randomly perturbing the system reveals its complexity
 - little effect when the maximizing structure is "evident"



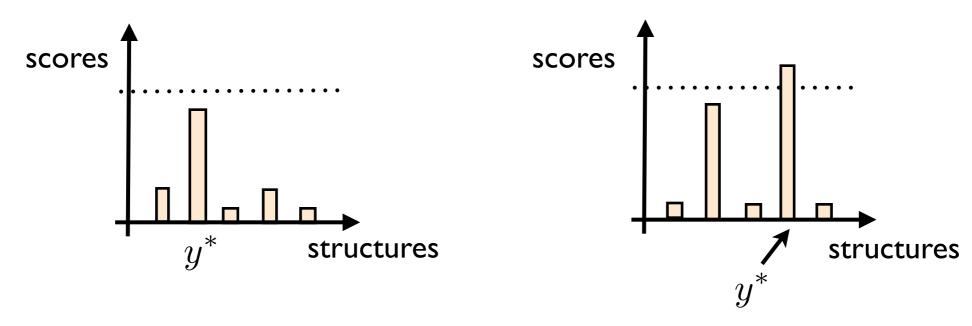
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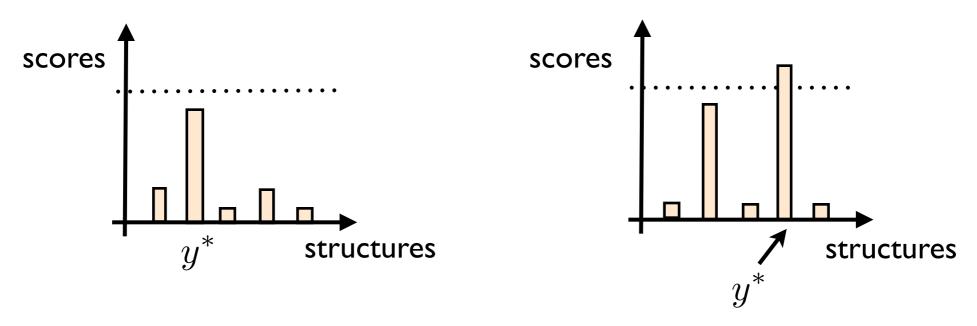
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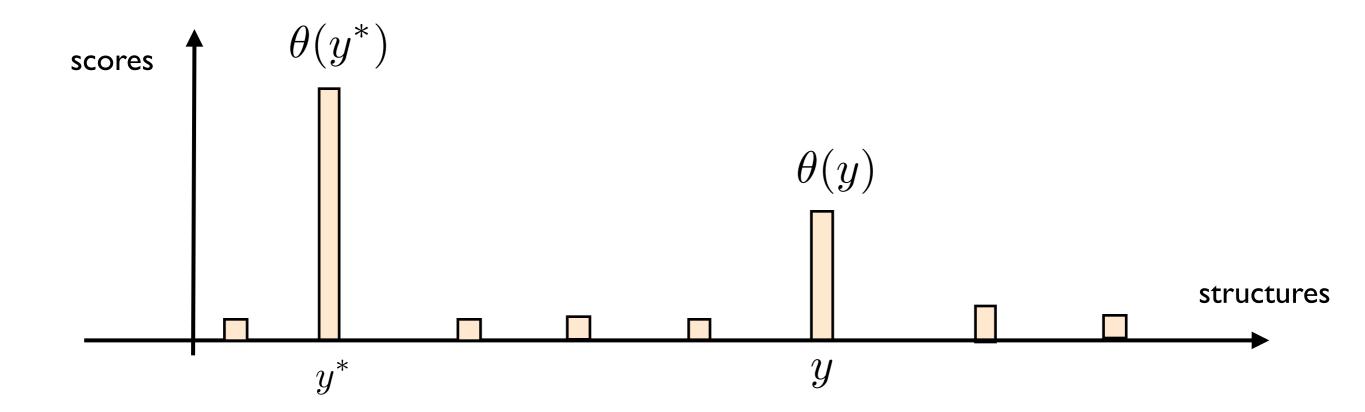
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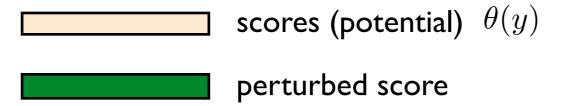
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- Related work:
 - McFadden 74 (Discrete choice theory)
 - Talagrand 94 (Canonical processes)

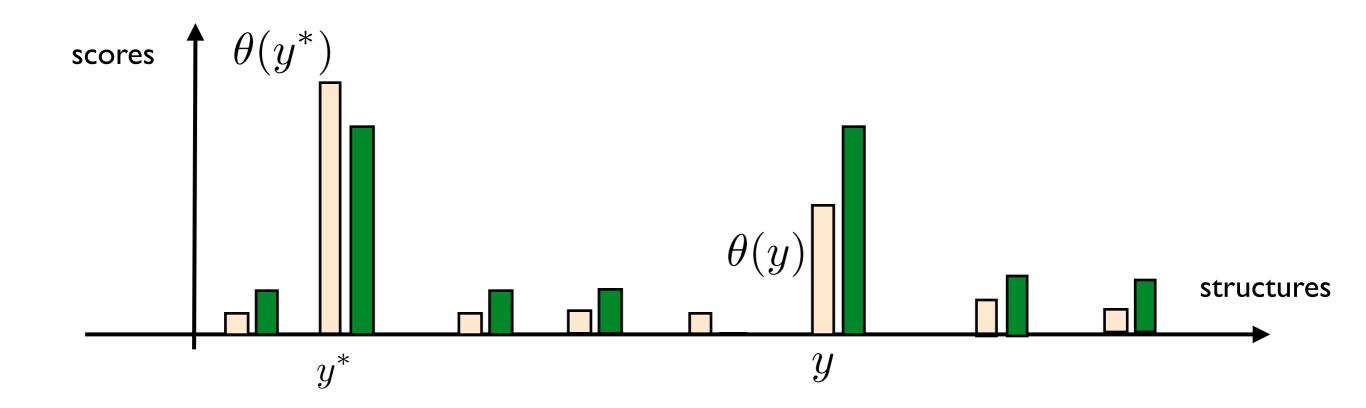
Notation:

scores (potential) $\theta(y)$

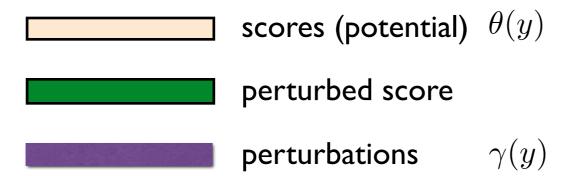


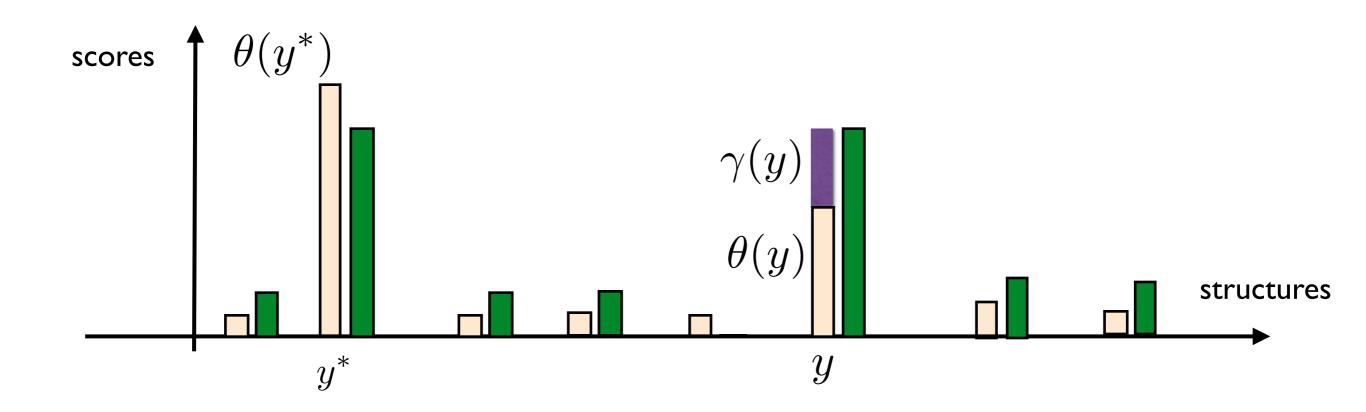
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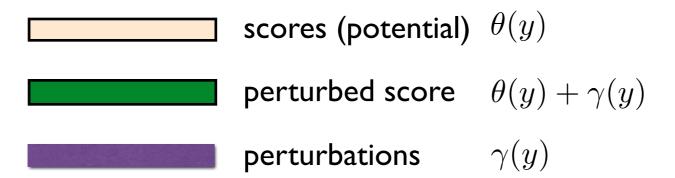


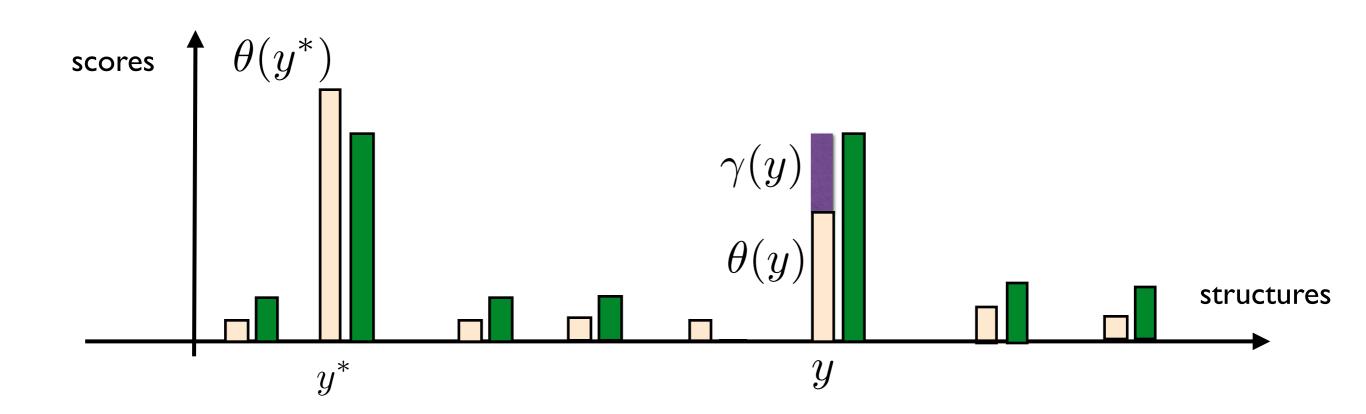
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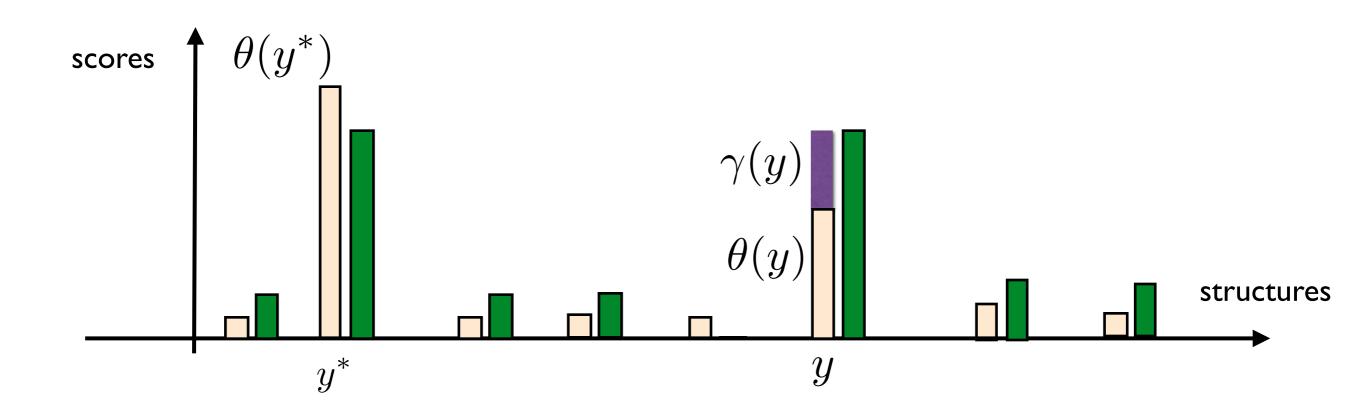


Notation:





- For every structure y, the perturbation value $\gamma(y)$ is a random variable (y is an index, traditional notation is γ_y).
- Perturb-max models: how stable is the maximal structure to random changes in the potential function.



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Theorem

Let $\gamma(y)$ be i.i.d. with Gumbel distribution with zero mean

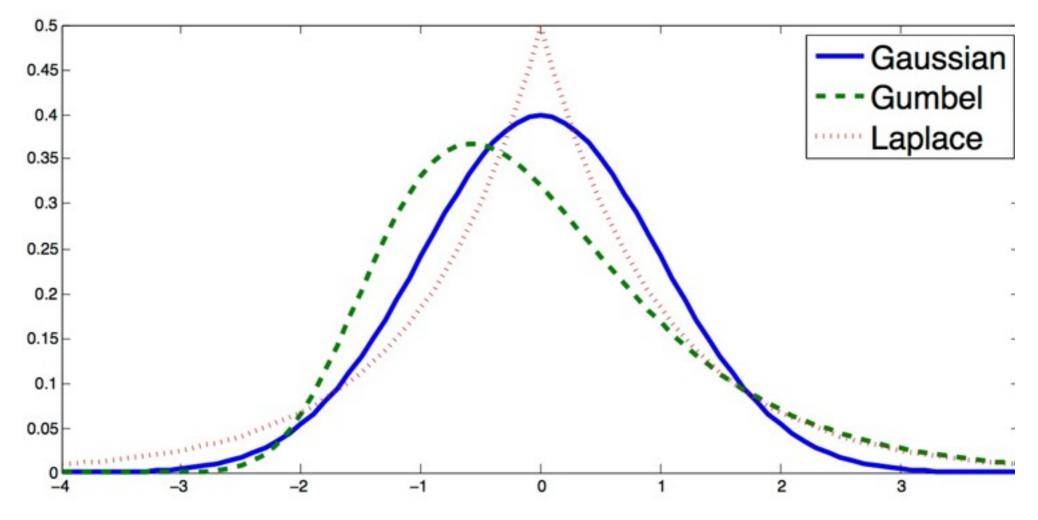
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$$f(t) = F'(t) = \exp(-t)F(t)$$





Theorem

Let $\gamma(y)$ be i.i.d. with Gumbel distribution with zero mean

$$F(t) \stackrel{def}{=} P[\gamma(y) \le t] = \exp(-\exp(-t))$$

then the perturb-max model is the Gibbs distribution

$$\frac{1}{Z} \exp(\theta(y)) = P_{\gamma \sim Gumbel}[y = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \gamma(\hat{y})\}]$$

- Why Gumbel distribution? $F(t) = \exp(-\exp(-t))$
- Since maximum of Gumbel variables is a Gumbel variable.

Let
$$\gamma(y)$$
 be i.i.d Gumbel ($P[\gamma(y) \leq t] = F(t)$). Then

$$\max_{y} \{\theta(y) + \gamma(y)\}$$

$$Z = \sum_{y} \exp(\theta(y))$$

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• Proof:
$$P_{\gamma}[\max_{y} \{\theta(y) + \gamma(y)\} \le t] = \prod_{y} F(t - \theta(y))$$

= $\exp(-\sum_{y} \exp(-(t - \theta(y)))) = F(t - \log Z)$



Max stability:

$$\log\left(\sum_{y} \exp(\theta(y))\right) = E_{\gamma \sim Gumbel} \left[\max_{y} \{\theta(y) + \gamma(y)\}\right]$$

Implications (taking gradients):

$$\frac{1}{Z} \exp(\theta(y)) = P_{\gamma \sim Gumbel}[y = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \gamma(\hat{y})\}]$$

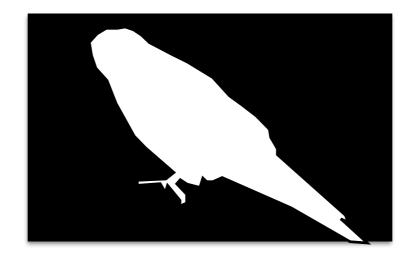
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 Representing the Gibbs distribution using perturb-max models may require exponential number of perturbations

$$P_{\gamma}[y = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \gamma(\hat{y})\}]$$

 Use low dimension perturbations [Papandreou & Yuille I I, Tarlow et. al I 2]

$$P_{\gamma}[y = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i)\}]$$

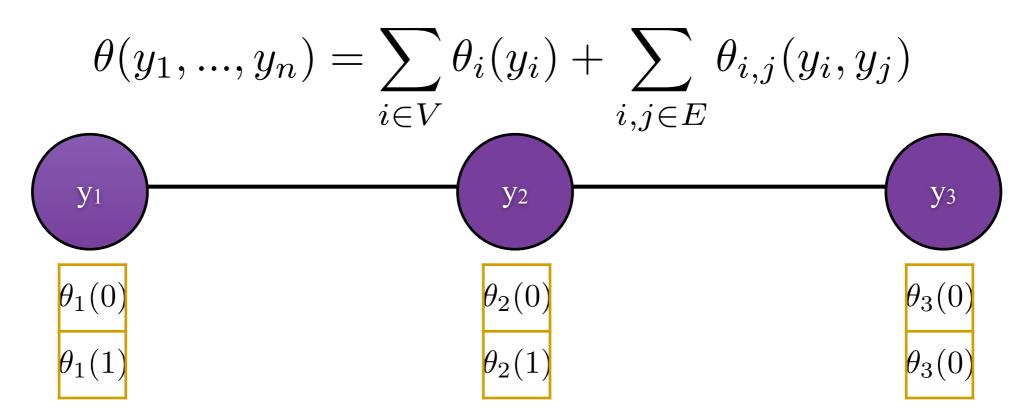
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$$\theta_{1,2}(0, 0)\theta_{1,2}(0, 1)$$

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$$\theta_{2,3}(0, 0)\theta_{2,3}(0, 1)$$

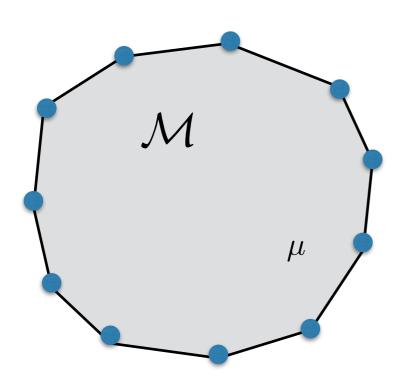
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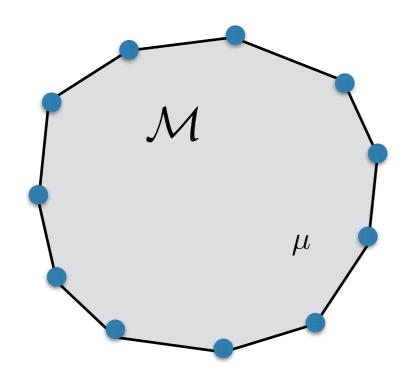
$$\theta_{3}(0)$$

$$\theta_{3}(0)$$

$$\theta(y_1, ..., y_n) = \sum_{i \in V} \theta_i(y_i) + \sum_{i,j \in E} \theta_{i,j}(y_i, y_j)$$

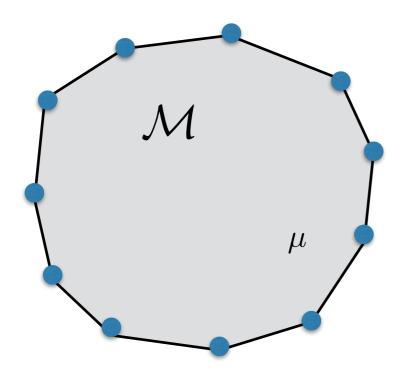


$$\theta(y_1, ..., y_n) = \sum_{i \in V} \theta_i(y_i) + \sum_{i,j \in E} \theta_{i,j}(y_i, y_j)$$



$$\mu = \begin{pmatrix} \mu_1(0), \mu_1(1), \mu_2(0), \mu_2(1), \mu_3(0), \mu_3(1), \\ \mu_{1,2}(0,0), \mu_{1,2}(0,1), \mu_{1,2}(1,0), \mu_{1,2}(1,1), \\ \mu_{2,3}(0,0), \mu_{2,3}(0,1), \mu_{2,3}(1,0), \mu_{2,3}(1,1) \end{pmatrix}$$

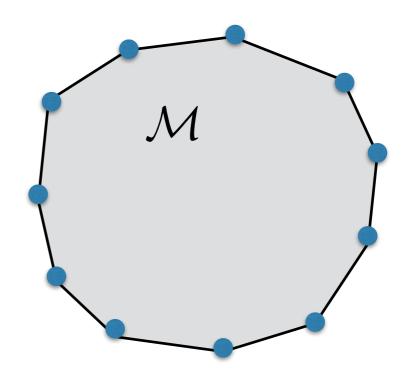
$$\theta(y_1,...,y_n) = \sum_{i \in V} \theta_i(y_i) + \sum_{i,j \in E} \theta_{i,j}(y_i,y_j)$$



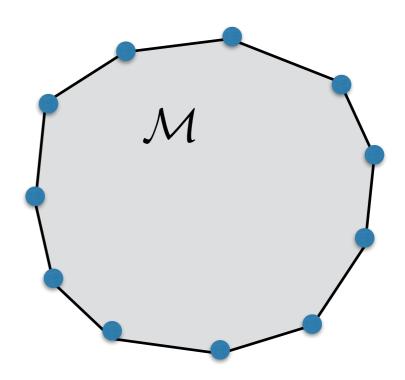
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$$\exists p(y_1, y_2, y_3) \text{ s.t. } \mu_1(y_1) = \sum_{y_2, y_3} p(y_1, y_2, y_3), \dots$$

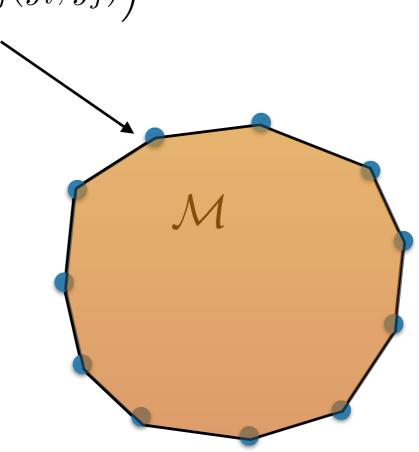
$$\mu_{1,2}(y_1, y_2) = \sum_{y_3} p(y_1, y_2, y_3), \dots$$



$$p(y) \propto \exp\left(\sum_{i} \theta_{i}(y_{i}) + \sum_{i,j} \theta_{i,j}(y_{i}, y_{j})\right)$$



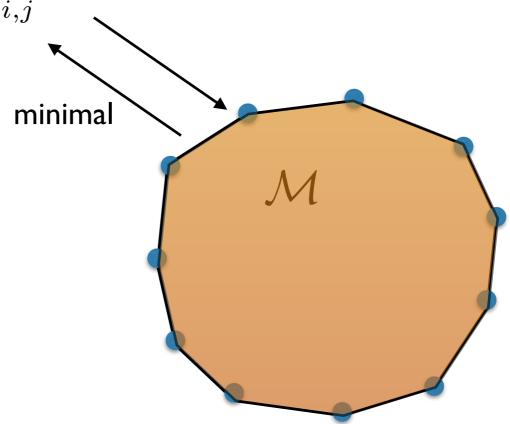
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 minimal

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[Wainwright & Jordan 08]



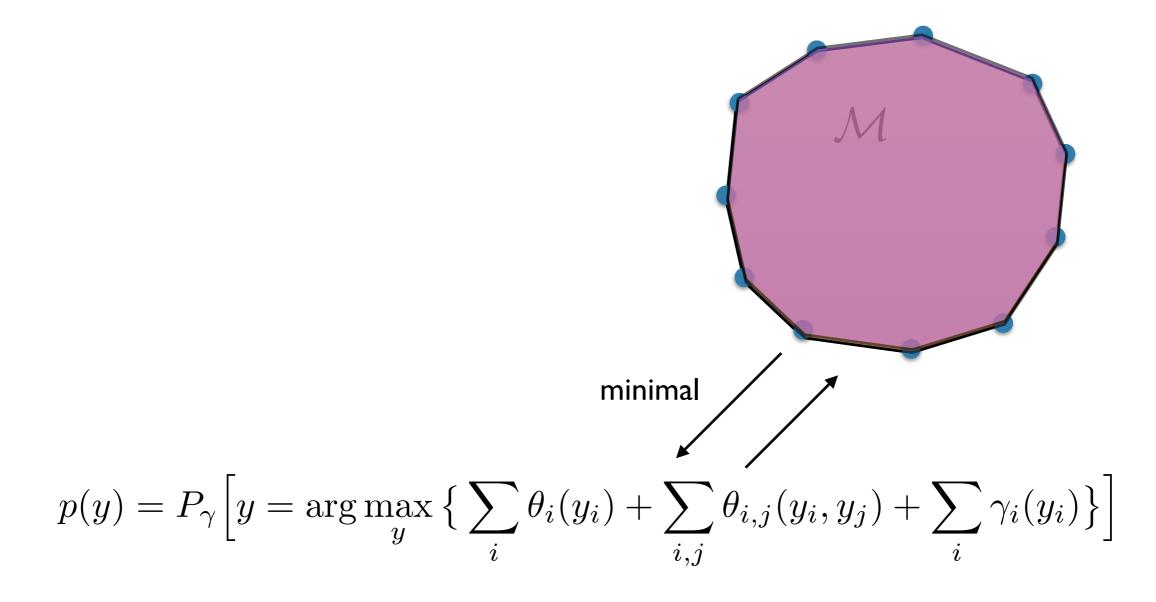
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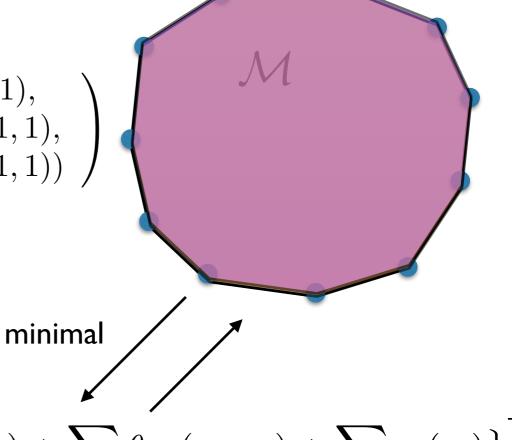
$$p(y) = P_{\gamma} \left[y = \arg \max_{y} \left\{ \sum_{i} \theta_{i}(y_{i}) + \sum_{i,j} \theta_{i,j}(y_{i}, y_{j}) + \sum_{i} \gamma_{i}(y_{i}) \right\} \right]$$

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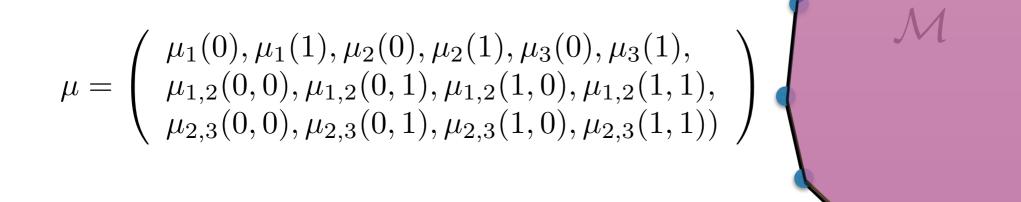
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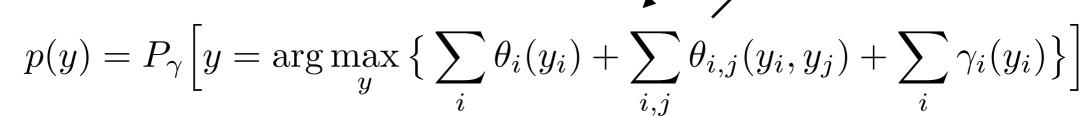


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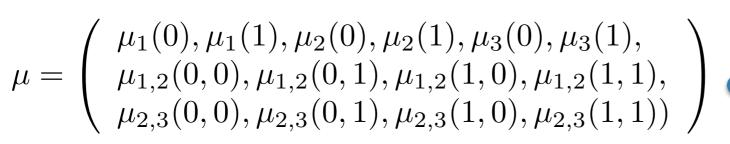


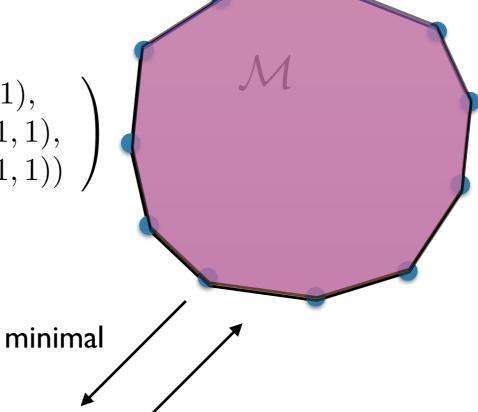


minimal

Proof idea:

$$\mu_i(y_i) = \frac{\partial E_{\gamma} \left[\max_y \left\{ \sum_i \theta_i(y_i) + \sum_{i,j} \theta_{i,j}(y_i, y_j) + \sum_i \gamma_i(y_i) \right\} \right]}{\partial \theta_i(y_i)}$$





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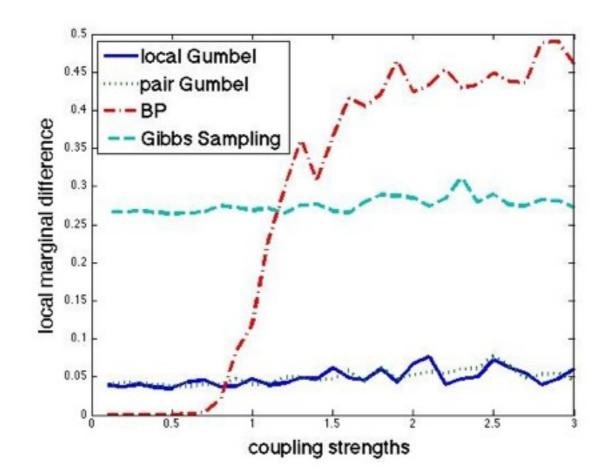
$$\mu_{i,j}(y_i, y_j) = \frac{\partial E_{\gamma} \left[\max_{y} \left\{ \sum_{i} \theta_i(y_i) + \sum_{i,j} \theta_{i,j}(y_i, y_j) + \sum_{i} \gamma_i(y_i) \right\} \right]}{\partial \theta_{i,j}(y_i, y_j)}$$

Outline

- Random perturbation why and how?
 - Sampling likely structures as fast as finding the most likely one.
- Connections and Alternatives to Gibbs distribution:
 - the marginal polytope
 - non-MCMC sampling for Gibbs with perturb-max
- Application: interactive annotation.
 - New entropy bounds for perturb-max models.

Non-MCMC sampling

- Perturb-max sample from tree-shaped Gibbs distribution [Gane, H, Jaakkola 14].
- Perturb-max + rejections sample from the Gibbs distribution on general graphs [H, Maji, Jaakkola 13].
- In practice, perturb-max marginals approximate the Gibbs marginals for general graphs [Papandreou & Yuille 11].



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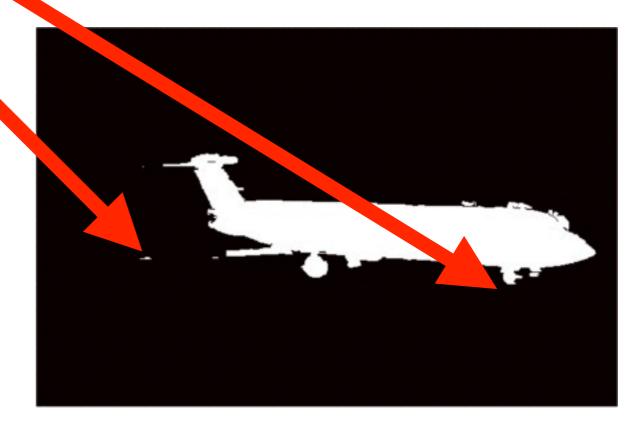


• Image annotation is a time consuming (and tedious) task. Can computers do it for us?

• Why not to use the most likely annotation instead?

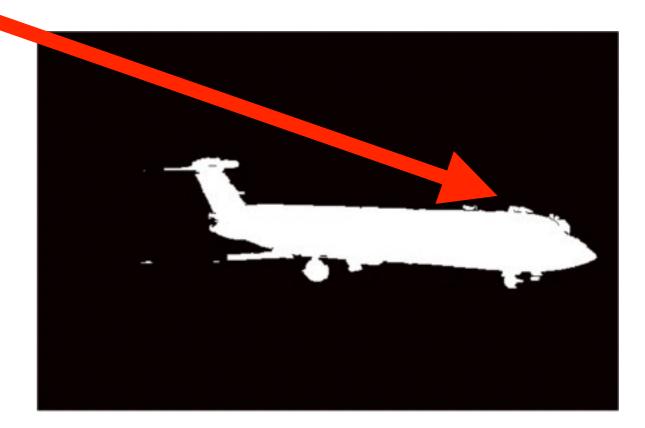
- Why not to use the most likely annotation instead?
- Most likely annotation is inaccurate around
 - "thin" areas





- Why not to use the most likely annotation instead?
- Most likely annotation is inaccurate around
 - "thin" areas
 - clutter





Interactive image annotation

• Perturb-max models show the boundary of decision.

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Interactive image annotation

Perturb-max models show the boundary of decision.



• Interactive annotation directs the human annotator to areas of uncertainty - significantly reduces annotation time [Maji, H., Jaakkola 14].

• Entropy
$$H(p_{ heta}) = -\sum_{y} p_{ heta}(y) \log p_{ heta}(y)$$

- Entropy = uncertainty
 - It is a nonnegative function over probability distributions.
 - It attains its maximal value for the uniform distribution.
 - It attains its minimal value for the zero-one distribution.
- Computing the entropy requires summing over exponential many configurations $y = (y_1, ..., y_n)$
- Can we bound it with perturb-max approach?

Perturb-max models

$$p_{\theta}(y) \stackrel{def}{=} P_{\gamma}[y = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i)\}]$$

Entropy

$$H(p_{\theta}) = -\sum_{y} p_{\theta}(y) \log p_{\theta}(y)$$

• Entropy bound $H(p_{\theta}) \leq E_{\gamma} \Big[\sum \gamma_i(y_i^*)\Big]$

$$y^* = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i)\}$$

$$U(p_{\theta}) = E_{\gamma} \left[\sum_{i=1}^{n} \gamma_i(y_i^*) \right] \qquad y^* = \arg\max_{\hat{y}} \{ \theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i) \}$$

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• $U(p_{\theta})$ is an uncertainty measure

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 - U(zero-one distribution) = 0

$$\theta(\hat{y}) = 0, \quad \forall y \neq \hat{y} \quad \theta(y) = -\infty$$

$$E[\gamma_i(\hat{y}_i)] = 0$$

$$U(p_{\theta}) = E_{\gamma} \left[\sum_{i=1}^{n} \gamma_i(y_i^*) \right] \qquad y^* = \arg\max_{\hat{y}} \{ \theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i) \}$$

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$$\theta(y) \equiv 0$$

higher $\theta(y)$ favor lower $\gamma(y)$ at the expanse of higher $\gamma(\hat{y})$

- How does it compare to standard entropy bounds?
- Perturb-max entropy bound:

$$H(p_{\theta}) \le E\left[\sum_{i} \gamma_{i}(y_{i}^{*})\right] = \sum_{i} E\left[\gamma_{i}(y_{i}^{*})\right]$$

Standard entropy independence bound:

$$H(p_{\theta}) \leq \sum_{i} H(p_{\theta}(y_i))$$

$$p_{\theta}(y_i) = P_{\gamma}[y_i = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^{n} \gamma_i(\hat{y}_i)\}]$$

 Perturb-max entropy bound requires less samples since sampled average tail decreases exponentially.

Perturb-max entropy bounds

Spin glass, 5x5 grid

$$\sum_{i} \theta_{i}(y_{i}) + \sum_{i,j} \theta_{i,j}(y_{i}, y_{j})$$

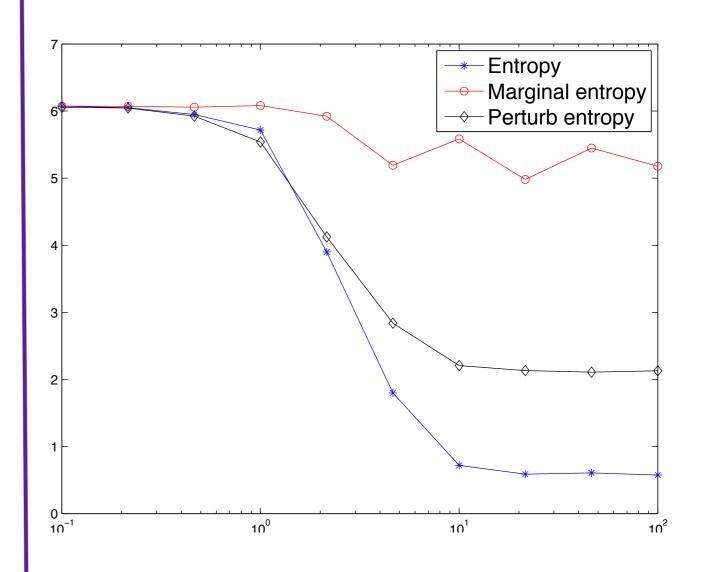
$$y_{i} \in \{-1, 1\}$$

$$\theta_{i}(y_{i}) = w_{i}y_{i}$$

$$w_{i} \sim N(0, 1)$$

$$\theta_{i,j}(y_{i}, y_{j}) = w_{i,j}y_{i}y_{j}$$

• attractive $w_{i,j} \geq 0$. Graph-cuts.



• Theorem:
$$H(p_{\theta}) \leq E\Big[\sum_i \gamma_i(y_i^*)\Big]$$

$$y^* = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^n \gamma_i(\hat{y}_i)\}$$

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$$\log Z(\hat{\theta}) \le E_{\gamma} \left[\max_{y} \left\{ \hat{\theta}(y) + \sum_{i} \gamma_{i}(y_{i}) \right\} \right]$$

The flashback slide



Max stability:

$$\log\left(\sum_{y} \exp(\theta(y))\right) = E_{\gamma \sim Gumbel} \left[\max_{y} \{\theta(y) + \gamma(y)\}\right]$$

• Theorem:
$$H(p_{\theta}) \leq E\Big[\sum_i \gamma_i(y_i^*)\Big]$$

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$$\begin{split} H(p) &= \min_{\hat{\theta}} \Big\{ \log Z(\hat{\theta}) - \sum_{y} \hat{\theta}(y) p(y) \Big\} \\ &\log Z(\hat{\theta}) \leq E_{\gamma} \Big[\max_{y} \big\{ \hat{\theta}(y) + \sum_{i} \gamma_{i}(y_{i}) \big\} \Big] \\ \frac{p_{\theta}}{H(p)} &\leq \min_{\hat{\theta}} \Big\{ E_{\gamma} \Big[\max_{y} \big\{ \hat{\theta}(y) + \sum_{i} \gamma_{i}(y_{i}) \big\} \Big] - \sum_{y} \hat{\theta}(y) p(y) \Big\} \end{split}$$

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oof idea: conjugate duality
$$H(p) = \min_{\hat{\theta}} \Big\{ \log Z(\hat{\theta}) - \sum_{y} \hat{\theta}(y) p(y) \Big\}$$

$$\log Z(\hat{\theta}) \leq E_{\gamma} \Big[\max_{y} \big\{ \hat{\theta}(y) + \sum_{i} \gamma_{i}(y_{i}) \big\} \Big]$$

$$p_{\theta} \qquad \qquad \hat{\theta}^{*} = \theta \qquad \qquad \hat{\theta}^{*} = \theta \qquad p_{\theta}$$

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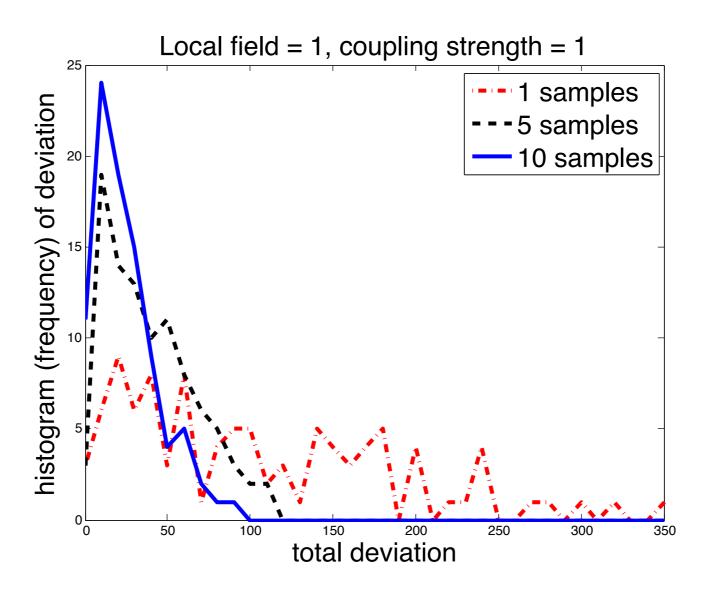
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The upper bounds hold in expectation.

$$H(p_{\theta}) \leq E\left[\sum_{i} \gamma_{i}(y_{i}^{*})\right]$$

$$y^{*} = \arg\max_{\hat{y}} \{\theta(\hat{y}) + \sum_{i=1}^{n} \gamma_{i}(\hat{y}_{i})\}$$

 The distance between the sampled average and the true expectation decays exponentially



$$P[\text{avg of M samples} \le \text{expectation} + r] \le \exp\left(-\frac{M}{20}\min(r, \frac{r^2}{n})\right)$$

[Orabona, H., Sarwate, Jaakkola 14], [Nguyen 14]

• Why is it hard to get exponential decay?

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$$P\left[\sum_{i} \gamma_{i}(y_{i}^{*}) > r\right] \leq \frac{E\left[\exp\left(\sum_{i} \gamma_{i}(y_{i}^{*})\right)\right]}{\exp(r)}$$

Why it is hard to get exponential decay?

moment generating function

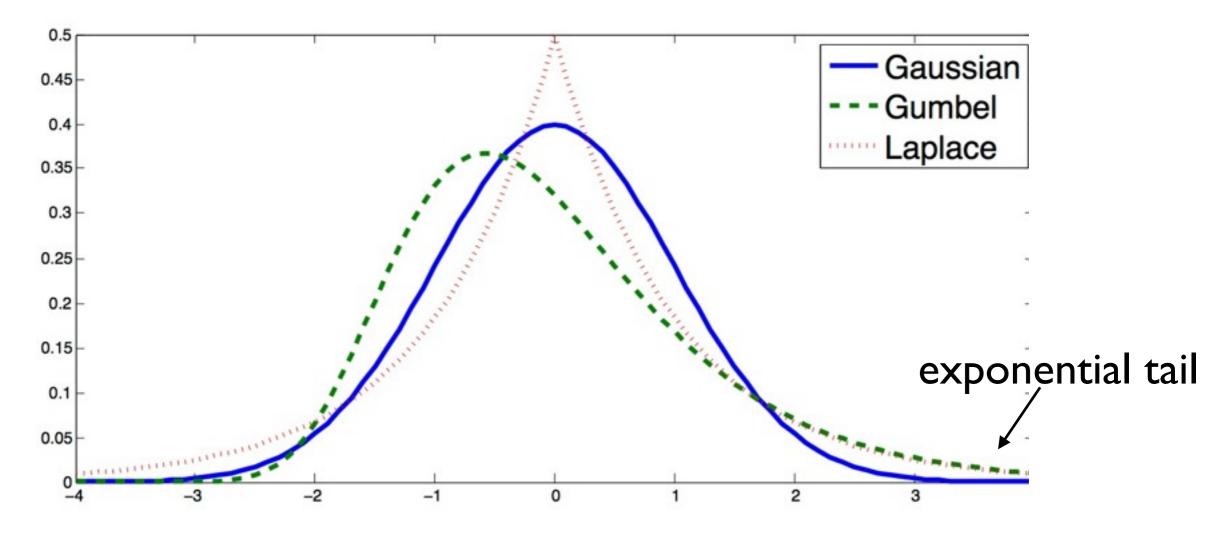
$$P\left[\sum_{i} \gamma_{i}(y_{i}^{*}) > r\right] \leq \frac{E\left[\exp\left(\sum_{i} \gamma_{i}(y_{i}^{*})\right)\right]}{\exp(r)}$$

- Measure concentration requires to bound the moment generating function
 - Hoeffding concentration requires bounded perturbations.
 - McDiarmid concentration requires bounded differences.
 - Our perturbations are unbounded with exponential tail.

The exponential tail of Gumbel distribution

$$E\left[\exp\left(\sum_{i} \gamma_{i}(y_{i}^{*})\right)\right] = \int q(\gamma_{i}(y_{i})) \to \exp\left(-\gamma_{i}(y_{i})\right)$$

$$i \to \exp\left(\sum_{i} \gamma_{i}(y_{i}^{*})\right)$$



• A function concentrates around its expectation if it does not change too much.

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 - Use tensorization to deal with one dimension at a time

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$$Var\left[\sum_{i} \gamma_{i}(y_{i}^{*})\right] = \sum_{j,y_{j}} Var_{\gamma_{j},y_{j}} \left[\sum_{i} \gamma_{i}(y_{i}^{*})\right]$$

- A function concentrates around its expectation if it does not change too much.
 - Use tensorization to deal with one dimension at a time

$$Var\left[\sum_{i} \gamma_{i}(y_{i}^{*})\right] = \sum_{j,y_{j}} Var_{\gamma_{j},y_{j}} \left[\sum_{i} \gamma_{i}(y_{i}^{*})\right]$$

- Bound any dimension's variance with its perturb-max probability (a Poincare inequality)

$$Var_{j,y_j}[\cdot] \le P_{\gamma_j(y_j)}[y_j = \arg\max_{y} \{\theta(y) + \sum_{i} \gamma_i(y_i)\}]$$

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- Connections and Alternatives to Gibbs distribution:
 - the marginal polytope
 - non-MCMC sampling for Gibbs distributions with perturb-max
- Application: interactive annotation.
 - New entropy bounds for perturb-max models.

Open problems

- Perturb-max models:
 - How do perturb-max models generalize Follow the Perturbed Leader [Manfred Warmuth, Jacob Abernethy]
 - Adversarial learning objective [lan Goodfellow]
 - Perturb-max models stabilize the prediction. Do they connect computational and statistical stability [Yury Makarychev]?
 - Perturb-max models in continuous space [Maddison et. al 14]
 - When does fixing variables in the max-function amount to statistical conditioning?
 - When do perturb-max models preserve the most likely assignment?
 - How do the perturbations dimension affect the model properties?
 - How to encourage diverse sampling?

Thank you

