## Perturbation, Optimization \& Statistics workshop

# Probabilistic inference by randomly perturbing max-solvers 

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## Inference in machine learning

- Complex structures dominate machine learning applications:


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- Computer vision



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- Computational biology

- and more..


## Outline

- Random perturbation - why and how?
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y \in\{0,1\}^{n}
$$



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- For machine learning we need to efficiently infer from distributions over complex structures.


## Inference in machine learning



## Gibbs distribution

$$
p\left(y_{1}, \ldots, y_{n}\right)=\frac{1}{Z} \exp \left(\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)\right)
$$

- MCMC samplers:


## Gibbs distribution

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- MCMC samplers:
- Gibbs sampling, Metropolis-Hastings, Swendsen-Wang
- Many efficient sampling algorithms for special cases:
- Counting bi-partite matchings in planar graphs (Kasteleyn 61)
- Ising models (Jerrum 93)
- Approximating the permanent (Jerrum 04)
- Many others...


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- Efficient sampling in Ising models (Jerrum 93)
- Attractive pairwise potentials

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\theta_{i, j}\left(y_{i}, y_{j}\right)= \begin{cases}w_{i, j} & \text { if } y_{i}=y_{j} \\ -w_{i, j} & \text { otherwise }\end{cases}
$$

$$
w_{i, j} \geq 0
$$

- No data terms

$$
\theta_{i}\left(y_{i}\right)=0
$$

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$$



- Nicely behaved distribution that is centered around the $(I, \ldots, I)$ or $(0, \ldots, 0)$


## Sampling likely structures

- Sampling from the Gibbs distribution is provably hard in Al applications (Goldberg 05, Jerrum 93)

- $x_{i}$ RGB color of pixel i

$$
\theta_{i}\left(y_{i}\right)=\log p\left(y_{i} \mid x_{i}\right)
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- Recall: sampling from the Gibbs distribution is easy in Ising models (Jerrum 93)

$\theta_{i}\left(y_{i}\right)=0$


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## Sampling likely structures

- Sampling from the Gibbs distribution is provably hard in Al applications (Goldberg 05, lerrum 93)
- Data terms (signals) that are important in Al applications significantly change the complexity of sampling


$$
\theta_{i}\left(y_{i}\right)=0
$$

- Recall: sampling from the Gibbs distribution is easy in Ising models (Jerrum 93)


## Most likely structure

- Instead of sampling, it may be significantly faster to find the most likely structure


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- The most likely structure



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- Instead of sampling, it may be significantly faster to find the most likely structure
- Graph-cuts
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- The most likely structure



## Most likely structure

$$
y^{*}=\arg \max _{y_{1}, \ldots, y_{n}} \sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)
$$

- Maximum a-posterior (MAP) inference.
- Many efficient optimization algorithms for special cases:
- Beliefs propagation: trees (Pearl 88), perfect graphs (Jebara 10),
- Graph-cuts for image segmentation
- branch and bound (Rother 09), branch and cut (Gurobi)
- Linear programming relaxations (Schlesinger 76,Wainwright 05, Kolmogorov 06, Werner 07, Sontag 08, Hazan I0, Batra IO, Nowozin IO, Pletscher 12, Kappes 13, SavchynskyyI3, Tarlow 13, Kohli I3, Jancsary I3, Schwing I3)
- CKY for parsing
- Many others...


## The challenge

Sampling from the likely high dimensional structures (with millions of variables, e.g., image segmentation with 12 million pixels) as efficient as optimizing

## Most likely structure

- Selecting the maximizing structure is appropriate when one structure (e.g., segmentation / parse) dominates others



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## Most likely structure

- The maximizing structure is not robust in case of multiple high scoring alternatives


## Most likely structure

- The maximizing structure is not robust in case of ambiguities



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- The maximizing structure is not robust in case of computationally limited models



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- Randomly perturbing the system reveals its complexity
- little effect when the maximizing structure is "evident"


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- substantial effect when there are alternative high scoring structures
- Related work:
- McFadden 74 (Discrete choice theory)
- Talagrand 94 (Canonical processes)


## Random perturbations

- Notation:
$\square$ scores (potential) $\theta(y)$



## Random perturbations

- Notation:
scores (potential) $\theta(y)$
perturbed score



## Random perturbations

- Notation:

| $\square$ | scores (potential) $\theta(y)$ |
| :--- | :--- |
| $\square$ | perturbed score |
| $\square$ | perturbations |



## Random perturbations

- Notation:

| $\square$ scores (potential) | $\theta(y)$ |  |
| :--- | :--- | :--- |
| perturbed score | $\theta(y)+\gamma(y)$ |  |
|  | perturbations | $\gamma(y)$ |



## Random perturbations

- For every structure $\boldsymbol{y}$, the perturbation value $\gamma(y)$ is a random variable ( y is an index, traditional notation is $\gamma_{y}$ ).
- Perturb-max models: how stable is the maximal structure to random changes in the potential function.



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## Perturb-max models

- Theorem

Let $\gamma(y)$ be i.i.d. with Gumbel distribution with zero mean

$$
F(t) \stackrel{\text { def }}{=} P[\gamma(y) \leq t]=\exp (-\exp (-t))
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\begin{aligned}
& F(t) \stackrel{\text { def }}{=} P[\gamma(y) \leq t]=\exp (-\exp (-t)) \\
& f(t)=F^{\prime}(t)=\exp (-t) F(t)
\end{aligned}
$$



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then the perturb-max model is the Gibbs distribution

$$
\frac{1}{Z} \exp (\theta(y))=P_{\gamma \sim G u m b e l}\left[y=\arg \max _{\hat{y}}\{\theta(\hat{y})+\gamma(\hat{y})\}\right]
$$

## Perturb-max models

-Why Gumbel distribution? $F(t)=\exp (-\exp (-t))$

- Since maximum of Gumbel variables is a Gumbel variable.

Let $\gamma(y)$ be i.i.d Gumbel ( $P[\gamma(y) \leq t]=F(t))$. Then

$$
\max _{y}\{\theta(y)+\gamma(y)\}
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$y$

$$
Z=\sum_{y} \exp (\theta(y))
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has Gumbel distribution whose mean is $\log Z$

- Proof: $\quad P_{\gamma}\left[\max _{y}\{\theta(y)+\gamma(y)\} \leq t\right]=\prod_{y} F(t-\theta(y))$

$$
=\exp \left(-\sum_{y} \exp (-(t-\theta(y)))\right)=F(t-\log Z)
$$

## Perturb-max models

- Max stability:

$$
\log \left(\sum_{y} \exp (\theta(y))\right)=E_{\gamma \sim \text { Gumbel }}\left[\max _{y}\{\theta(y)+\gamma(y)\}\right]
$$

- Implications (taking gradients):

$$
\frac{1}{Z} \exp (\theta(y))=P_{\gamma \sim G u m b e l}\left[y=\arg \max _{\hat{y}}\{\theta(\hat{y})+\gamma(\hat{y})\}\right]
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$$

- Use low dimension perturbations [Papandreou \& Yuillel I, Tarlow et.all2]

$$
P_{\gamma}\left[y=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}\right]
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## The marginal polytope

$$
\theta\left(y_{1}, \ldots, y_{n}\right)=\sum_{i \in V} \theta_{i}\left(y_{i}\right)+\sum_{i, j \in E} \theta_{i, j}\left(y_{i}, y_{j}\right)
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$$
\mu=\left(\begin{array}{l}
\mu_{1}(0), \mu_{1}(1), \mu_{2}(0), \mu_{2}(1), \mu_{3}(0), \mu_{3}(1), \\
\mu_{1,2}(0,0), \mu_{1,2}(0,1), \mu_{1,2}(1,0), \mu_{1,2}(1,1), \\
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$$
\begin{gathered}
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\end{array}\right) \\
\exists p\left(y_{1}, y_{2}, y_{3}\right) \text { s.t. } \quad \mu_{1}\left(y_{1}\right)=\sum_{y_{2}, y_{3}} p\left(y_{1}, y_{2}, y_{3}\right), \ldots \\
\mu_{1,2}\left(y_{1}, y_{2}\right)=\sum_{y_{3}} p\left(y_{1}, y_{2}, y_{3}\right), \ldots
\end{gathered}
$$

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[Wainwright \& Jordan 08]


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$$
p(y)=P_{\gamma}\left[y=\arg \max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]
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\end{array}\right) \\
& p(y)=P_{\gamma}\left[y=\arg \max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]
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\end{array}\right)
$$



- Proof idea:

$$
\mu_{i}\left(y_{i}\right)=\frac{\partial E_{\gamma}\left[\max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]}{\partial \theta_{i}\left(y_{i}\right)}
$$

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$$



$$
p(y)=P_{\gamma}\left[y=\arg \max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]
$$

- Proof idea:

$$
\begin{aligned}
\mu_{i}\left(y_{i}\right) & =\frac{\partial E_{\gamma}\left[\max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]}{\partial \theta_{i}\left(y_{i}\right)} \\
\mu_{i, j}\left(y_{i}, y_{j}\right) & =\frac{\partial E_{\gamma}\left[\max _{y}\left\{\sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]}{\partial \theta_{i, j}\left(y_{i}, y_{j}\right)}
\end{aligned}
$$

## Outline

- Random perturbation - why and how?
- Sampling likely structures as fast as finding the most likely one.
- Connections and Alternatives to Gibbs distribution:
- the marginal polytope
- non-MCMC sampling for Gibbs with perturb-max
- Application: interactive annotation.
- New entropy bounds for perturb-max models.


## Non-MCMC sampling

- Perturb-max sample from tree-shaped Gibbs distribution [Gane, H, Jaakkola I4].
- Perturb-max + rejections sample from the Gibbs distribution on general graphs [H, Maji, Jaakkola I3].
- In practice, perturb-max marginals approximate the Gibbs marginals for general graphs [Papandreou \& Yuille II].



## Outline

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## Image annotation



- Image annotation is a time consuming (and tedious) task. Can computers do it for us?


## Image annotation

- Why not to use the most likely annotation instead?


## Image annotation

- Why not to use the most likely annotation instead?
- Most likely annotation is inaccurate around
- "thin" areas


## Image annotation

- Why not to use the most likely annotation instead?
- Most likely annotation is inaccurate around
- "thin" areas
- clutter



## Interactive image annotation

- Perturb-max models show the boundary of decision.


## Interactive image annotation

- Perturb-max models show the boundary of decision.



## Interactive image annotation

- Perturb-max models show the boundary of decision.

- Interactive annotation directs the human annotator to areas of uncertainty - significantly reduces annotation time [Maji, H., Jaakkola 14].


## Uncertainty

- Entropy

$$
H\left(p_{\theta}\right)=-\sum_{y} p_{\theta}(y) \log p_{\theta}(y)
$$

- Entropy = uncertainty
- It is a nonnegative function over probability distributions.
- It attains its maximal value for the uniform distribution.
- It attains its minimal value for the zero-one distribution.
- Computing the entropy requires summing over exponential many configurations $y=\left(y_{1}, \ldots, y_{n}\right)$
- Can we bound it with perturb-max approach?


## Uncertainty

- Perturb-max models

$$
p_{\theta}(y) \stackrel{\text { def }}{=} P_{\gamma}\left[y=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}\right]
$$

- Entropy

$$
H\left(p_{\theta}\right)=-\sum_{y} p_{\theta}(y) \log p_{\theta}(y)
$$

- Entropy bound $\quad H\left(p_{\theta}\right) \leq E_{\gamma}\left[\sum_{i=1}^{n} \gamma_{i}\left(y_{i}^{*}\right)\right]$

$$
y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
$$

## Uncertainty

$$
U\left(p_{\theta}\right)=E_{\gamma}\left[\sum_{i=1}^{n} \gamma_{i}\left(y_{i}^{*}\right)\right]
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- $U\left(p_{\theta}\right)$ is an uncertainty measure


## Uncertainty

$$
U\left(p_{\theta}\right)=E_{\gamma}\left[\sum_{i=1}^{n} \gamma_{i}\left(y_{i}^{*}\right)\right] \quad y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
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- $U\left(p_{\theta}\right)$ is an uncertainty measure
- $U\left(p_{\theta}\right)$ is nonnegative since $0 \leq H\left(p_{\theta}\right) \leq U\left(p_{\theta}\right)$


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U\left(p_{\theta}\right)=E_{\gamma}\left[\sum_{i=1}^{n} \gamma_{i}\left(y_{i}^{*}\right)\right] \quad y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
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- $U($ zero-one distribution $)=0$

$$
\begin{aligned}
& \theta(\hat{y})=0, \quad \forall y \neq \hat{y} \quad \theta(y)=-\infty \\
& E\left[\gamma_{i}\left(\hat{y}_{i}\right)\right]=0
\end{aligned}
$$

## Uncertainty

$$
U\left(p_{\theta}\right)=E_{\gamma}\left[\sum_{i=1}^{n} \gamma_{i}\left(y_{i}^{*}\right)\right] \quad y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
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- $U($ zero-one distribution $)=0$
- $U($ uniform distribution $)=$ maximal

$$
\theta(y) \equiv 0
$$

higher $\theta(y)$ favor lower $\gamma(y)$ at the expanse of higher $\gamma(\hat{y})$

## Uncertainty

- How does it compare to standard entropy bounds?
- Perturb-max entropy bound:

$$
H\left(p_{\theta}\right) \leq E\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]=\sum_{i} E\left[\gamma_{i}\left(y_{i}^{*}\right)\right]
$$

- Standard entropy independence bound:

$$
\begin{aligned}
& H\left(p_{\theta}\right) \leq \sum_{i} H\left(p_{\theta}\left(y_{i}\right)\right) \\
& p_{\theta}\left(y_{i}\right)=P_{\gamma}\left[y_{i}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}\right]
\end{aligned}
$$

- Perturb-max entropy bound requires less samples since sampled average tail decreases exponentially.


## Perturb-max entropy bounds

- Spin glass, $5 \times 5$ grid

$$
\begin{aligned}
& \sum_{i} \theta_{i}\left(y_{i}\right)+\sum_{i, j} \theta_{i, j}\left(y_{i}, y_{j}\right) \\
& y_{i} \in\{-1,1\} \\
& \theta_{i}\left(y_{i}\right)=w_{i} y_{i} \\
& w_{i} \sim N(0,1) \\
& \theta_{i, j}\left(y_{i}, y_{j}\right)=w_{i, j} y_{i} y_{j}
\end{aligned}
$$

- attractive $w_{i, j} \geq 0$. Graph-cuts.



## Uncertainty*

- Theorem: $H\left(p_{\theta}\right) \leq E\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]$

$$
y^{*}=\arg \max _{\hat{y}}^{i}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
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- Proof idea: conjugate duality

$$
H(p)=\min _{\hat{\theta}}\left\{\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\}
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H(p)=\min _{\hat{\theta}}\{ & \left.\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
& \log Z(\hat{\theta}) \leq E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]
\end{aligned}
$$

## The flashback slide

- Max stability:

$$
\log \left(\sum_{y} \exp (\theta(y))\right)=E_{\gamma \sim \text { Gumbel }}\left[\max _{y}\{\theta(y)+\gamma(y)\}\right]
$$

## Uncertainty*

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& H(p)= \min _{\hat{\theta}}\left\{\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
& \log Z(\hat{\theta}) \leq E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right] \\
& H(p) \leq \min _{\hat{\theta}}\left\{E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y} \hat{\theta}(y) p(y)\right\}
\end{aligned}
$$

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$$

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& H(p)=\min _{\hat{\theta}}\left\{\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
& \\
& \quad \log Z(\hat{\theta}) \leq E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right] \\
& p_{\theta} \\
& H(p) \leq \min _{\hat{\theta}}\left\{E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y} \hat{\theta}(y) p(y)\right\}
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$$

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$$

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y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
$$

- Proof idea: conjugate duality

$$
\begin{gathered}
H(p)=\min _{\hat{\theta}}\left\{\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
\log Z(\hat{\theta}) \leq E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right] \\
p_{\theta} \quad \hat{\theta}^{*}=\theta \\
H(p) \leq \min _{\hat{\theta}}\left\{E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
H\left(p_{\theta}\right) \leq E_{\gamma}\left[\max _{y}\left\{\theta(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y} \theta(y) p_{\theta}(y)
\end{gathered}
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y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
$$

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$$
\begin{aligned}
& H(p)=\min _{\hat{\theta}}\left\{\log Z(\hat{\theta})-\sum_{y} \hat{\theta}(y) p(y)\right\} \\
& \log Z(\hat{\theta}) \leq E_{\gamma}\left[\max _{y}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right] \\
& \begin{array}{c}
p_{\theta} \\
H(p) \leq \min _{\hat{\theta}}\left\{E_{\gamma}\left[\max _{y}^{*}\left\{\hat{\theta}(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y}^{\hat{\theta}^{*}}=\theta \quad p_{\theta}\right. \\
\hat{\theta}(y) p(y)\}
\end{array} \\
& \left.\left.H\left(p_{\theta}\right) \leq E_{\gamma}\left[\begin{array}{l}
y
\end{array}\right](y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]-\sum_{y} \theta(y)
\end{aligned}
$$

## Sample complexity*

- The upper bounds hold in expectation.

$$
\begin{aligned}
& H\left(p_{\theta}\right) \leq E\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right] \\
& y^{*}=\arg \max _{\hat{y}}\left\{\theta(\hat{y})+\sum_{i=1}^{n} \gamma_{i}\left(\hat{y}_{i}\right)\right\}
\end{aligned}
$$

- The distance between the sampled average and the true expectation decays exponentially


## Sample complexity*


$P[$ avg of M samples $\leq \operatorname{expectation}+r] \leq \exp \left(-\frac{M}{20} \min \left(r, \frac{r^{2}}{n}\right)\right)$
[Orabona, H., Sarwate, Jaakkola I4], [Nguyen I4]

## Sample complexity*

-Why is it hard to get exponential decay?

## Sample complexity*

-Why is it hard to get exponential decay?

$$
P\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)>r\right] \leq \frac{E\left[\exp \left(\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right)\right]}{\exp (r)}
$$

## Sample complexity*

-Why it is hard to get exponential decay?

$$
P\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)>r\right] \leq \frac{E\left[\exp \left(\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right)\right]}{\exp (r)}
$$

- Measure concentration requires to bound the moment generating function
- Hoeffding concentration requires bounded perturbations.
- McDiarmid concentration requires bounded differences.
- Our perturbations are unbounded with exponential tail.


## Sample complexity*

- The exponential tail of Gumbel distribution

$$
E\left[\exp \left(\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right)\right]=\int^{q\left(\gamma_{i}\left(y_{i}\right)\right) \rightarrow \exp \left(-\gamma_{i}\left(y_{i}\right)\right)} q(\gamma) \exp \left(\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right)
$$



## Sample complexity*

- A function concentrates around its expectation if it does not change too much.


## Sample complexity*

- A function concentrates around its expectation if it does not change too much.
- Use tensorization to deal with one dimension at a time


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- A function concentrates around its expectation if it does not change too much.
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$$
\operatorname{Var}\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]=\sum_{j, y_{j}} \operatorname{Var}_{\gamma_{j, y_{j}}}\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]
$$

## Sample complexity*

- A function concentrates around its expectation if it does not change too much.
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$$
\operatorname{Var}\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]=\sum_{j, y_{j}} \operatorname{Var}_{\gamma_{j, y_{j}}}\left[\sum_{i} \gamma_{i}\left(y_{i}^{*}\right)\right]
$$

- Bound any dimension's variance with its perturb-max probability (a Poincare inequality)

$$
\operatorname{Var}_{j, y_{j}}[\cdot] \leq P_{\gamma_{j}\left(y_{j}\right)}\left[y_{j}=\arg \max _{y}\left\{\theta(y)+\sum_{i} \gamma_{i}\left(y_{i}\right)\right\}\right]
$$

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## Open problems

- Perturb-max models:
- How do perturb-max models generalize - Follow the Perturbed Leader [Manfred Warmuth, Jacob Abernethy]
- Adversarial learning objective [lan Goodfellow]
- Perturb-max models stabilize the prediction. Do they connect computational and statistical stability [Yury Makarychev]?
- Perturb-max models in continuous space [Maddison et. al I4]
- When does fixing variables in the max-function amount to statistical conditioning?
- When do perturb-max models preserve the most likely assignment?
- How do the perturbations dimension affect the model properties?
- How to encourage diverse sampling?

Thank you


