# STA 4273H: Statistical Machine Learning 

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## Lecture 4

## Graphical Models

- Probabilistic graphical models provide a powerful framework for representing dependency structure between random variables.
- Graphical models offer several useful properties:
- They provide a simple way to visualize the structure of a probabilistic model and can be used to motivate new models.
- They provide various insights into the properties of the model, including conditional independence.
- Complex computations (e.g. inference and learning in sophisticated models) can be expressed in terms of graphical manipulations.


## Graphical Models

- A graph contains a set of nodes (vertices) connected by links (edges or arcs)

- In a probabilistic graphical model, each node represents a random variable, and links represent probabilistic dependencies between random variables.
- The graph specifies the way in which the joint distribution over all random variables decomposes into a product of factors, where each factor depends on a subset of the variables.
- Two types of graphical models:
- Bayesian networks, also known as Directed Graphical Models (the links have a particular directionality indicated by the arrows)
- Markov Random Fields, also known as Undirected Graphical Models (the links do not carry arrows and have no directional significance).
- Hybrid graphical models that combine directed and undirected graphical models, such as Deep Belief Networks.


## Bayesian Networks

- Directed Graphs are useful for expressing causal relationships between random variables.
- Let us consider an arbitrary joint distribution $p(a, b, c)$ over three random variables $a, b$, and $c$.
- Note that at this point, we do not need to specify anything else about these variables (e.g. whether they are discrete or continuous).
- By application of the product rule of probability (twice), we get

$$
p(a, b, c)=p(c \mid a, b) p(a, b)=p(c \mid a, b) p(b \mid a) p(a)
$$

- This decomposition holds for any choice of the joint distribution.


## Bayesian Networks

- By application of the product rule of probability (twice), we get

$$
p(a, b, c)=p(c \mid a, b) p(a, b)=p(c \mid a, b) p(b \mid a) p(a)
$$

- Represent the joint distribution in terms of a simple graphical model:

- Introduce a node for each of the random variables.
- Associate each node with the corresponding conditional distribution in above equation.
- For each conditional distribution we add directed links to the graph from the nodes corresponding to the variables on which the distribution is conditioned.
- Hence for the factor $p(c \mid a, b)$, there will be links from nodes $\mathbf{a}$ and $\mathbf{b}$ to node c.
- For the factor $p(a)$, there will be no incoming links.


## Bayesian Networks

- By application of the product rule of probability (twice), we get

$$
p(a, b, c)=p(c \mid a, b) p(a, b)=p(c \mid a, b) p(b \mid a) p(a)
$$

- If there is a link going from node a to node $b$, then we say that:

- node $a$ is a parent of node $b$.
- node $b$ is a child of node $a$.
- For the decomposition, we choose a specific ordering of the random variables: $a, b, c$.
- If we chose a different ordering, we would get a different graphical representation (we will come back to that point later).
- The joint distribution over K variables factorizes:

$$
p\left(x_{1}, \ldots, x_{K}\right)=p\left(x_{K} \mid x_{1}, \ldots, x_{K-1}\right) \ldots p\left(x_{2} \mid x_{1}\right) p\left(x_{1}\right)
$$

- If each node has incoming links from all lower numbered nodes, then the graph is fully connected; there is a link between all pairs of nodes.


## Bayesian Networks

- Absence of links conveys certain information about the properties of the class of distributions that the graph conveys.

- Note that this graph is not fully connected (e.g. there is no link from $x_{1}$ to $x_{2}$ ).
- The joint distribution over $\mathrm{x}_{1}, \ldots, \mathrm{x}_{7}$ can be written as a product of a set of conditional distributions.

$$
\begin{aligned}
p\left(x_{1}, \ldots, x_{7}\right)= & p\left(x_{1}\right) p\left(x_{2}\right) p\left(x_{3}\right) p\left(x_{4} \mid x_{1}, x_{2}, x_{3}\right) \\
& p\left(x_{5} \mid x_{1}, x_{3}\right) p\left(x_{6} \mid x_{4}\right) p\left(x_{7} \mid x_{4}, x_{5}\right)
\end{aligned}
$$

- Note that according to the graph, $\mathrm{x}_{5}$ will be conditioned only on $\mathrm{x}_{1}$ and $\mathrm{x}_{3}$.


## Factorization Property

- The joint distribution defined by the graph is given by the product of a conditional distribution for each node conditioned on its parents:


$$
p(\mathbf{x})=\prod_{k=1}^{K} p\left(x_{k} \mid \mathrm{pa}_{k}\right)
$$

where $p a_{k}$ denotes a set of parents for the node $\mathrm{X}_{\mathrm{k}}$.

- This equation expresses a key factorization property of the joint distribution for a directed graphical model.
- Important restriction: There must be no directed cycles!
- Such graphs are also called directed acyclic graphs (DAGs).


## Bayesian Curve Fitting

- As an example, remember Bayesian polynomial regression model:

$$
y(x, \mathbf{w})=\sum_{j=0}^{M} w_{j} x^{j}
$$

- We are given inputs $\mathbf{X}=\left\{x_{1}, x_{2}, \ldots, x_{N}\right\}$ and target values $\mathbf{t}=\left[t_{1}, t_{2}, \ldots, t_{N}\right]^{T}$.
- Given the prior over parameters, the joint distribution is given by:


$$
p(\mathbf{t}, \mathbf{w} \mid \mathbf{X})=p(\mathbf{w}) \underbrace{\prod_{i=1}^{N} p\left(t_{n} \mid y\left(\mathbf{w}, x_{n}\right)\right)}_{\text {Prior term }} \text {. }
$$

## Graphical Representation

$$
p(\mathbf{t}, \mathbf{w} \mid \mathbf{X})=p(\mathbf{w}) \prod_{i=1}^{N} p\left(t_{n} \mid y\left(\mathbf{w}, x_{n}\right)\right)
$$

- This distribution can be represented as a graphical model.

- Same representation using plate notation.

- Compact representation: we introduce a plate that represents N nodes of which only a single example $t_{n}$ is shown explicitly.
- Note that $\mathbf{w}$ and $\mathbf{t}=\left[t_{1}, t_{2}, \ldots, t_{N}\right]^{T}$ represent random variables.


## Graphical Representation

- It will often be useful to make the parameters of the model as well as random variables be explicit.

$$
p\left(\mathbf{t}, \mathbf{w} \mid \mathbf{x}, \alpha, \sigma^{2}\right)=p(\mathbf{w} \mid \alpha) \prod_{n=1}^{N} p\left(t_{n} \mid \mathbf{w}, x_{n}, \sigma^{2}\right) .
$$



- Random variables will be denoted by open circles and deterministic parameters will be denoted by smaller solid circles.


## Graphical Representation

- When we apply a graphical model to a problem in machine learning, we will set some of the variables to specific observed values (e.g. condition on the data).

$p(\mathbf{w} \mid \mathbf{t}) \propto p(\mathbf{w}) \prod_{n=1}^{N} p\left(t_{n} \mid \mathbf{w}\right)$
- For example, having observed the values of the targets $\left\{\mathrm{t}_{\mathrm{n}}\right\}$ on the training data, we wish to infer the posterior distribution over parameters w.
- In this example, we conditioned on observed data $\mathbf{t}=\left[t_{1}, t_{2}, \ldots, t_{N}\right]^{T}$ by shadowing the corresponding nodes.


## Predictive Distribution

- We may also be interested in making predictions for a new input value $\hat{x}$.

$$
p\left(\widehat{t} \widehat{x}, \mathbf{x}, \mathbf{t}, \alpha, \sigma^{2}\right) \propto \int p\left(\widehat{t}, \mathbf{t}, \mathbf{w} \mid \widehat{x}, \mathbf{x}, \alpha, \sigma^{2}\right) \mathrm{d} \mathbf{w}
$$

 variables is given by:

$$
p\left(\widehat{t}, \mathbf{t}, \mathbf{w} \mid \widehat{x}, \mathbf{x}, \alpha, \sigma^{2}\right)=
$$

$$
\left[\prod_{n=1}^{N} p\left(t_{n} \mid x_{n}, \mathbf{w}, \sigma^{2}\right)\right] p(\mathbf{w} \mid \alpha) p\left(\hat{t} \mid \widehat{x}, \mathbf{w}, \sigma^{2}\right)
$$

- Here we are setting the random variables in $t$ to the specific values observed in the data.


## Ancestral Sampling

- Consider a joint distribution over K random variables $p\left(x_{1}, x_{2}, \ldots, x_{K}\right)$ that factorizes as:

$$
p(\mathbf{x})=\prod_{k=1}^{K} p\left(x_{k} \mid \mathrm{pa}_{k}\right)
$$



- Our goal is draw a sample from this distribution.
- Start at the top and sample in order.

$$
\begin{aligned}
& \hat{x}_{1} \sim p\left(x_{1}\right) \\
& \hat{x}_{2} \sim p\left(x_{2}\right) \\
& \hat{x}_{3} \sim p\left(x_{3}\right) \\
& \hat{x}_{4} \sim p\left(x_{4} \mid \hat{x}_{1}, \hat{x}_{2}, \hat{x}_{3}\right) \\
& \hat{x}_{5} \sim p\left(x_{5} \mid \hat{x}_{1}, \hat{x}_{3}\right)
\end{aligned}
$$

- To obtain a sample from the marginal distribution, e.g. $p\left(x_{2}, x_{5}\right)$, we sample from the full joint distribution, retain $\hat{x}_{2}, \hat{x}_{5}$, and discard the remaining values.


## Generative Models

- Higher-level nodes will typically represent latent (hidden) random variables.
- The primary role of the latent variables is to allow a complicated distribution over observed variables to be constructed from simpler (typically exponential family) conditional distributions.

Generative Model of an Image


- Object identity, position, and orientation have independent prior probabilities.
- The image has a probability distribution that depends on the object identity, position, and orientation (likelihood function).

$$
P(\mathrm{Im}, \mathrm{Ob}, \mathrm{Po}, \mathrm{Or})=\underbrace{P(\mathrm{Im} \mid \mathrm{Ob}, \mathrm{Po}, \mathrm{Or}}_{\text {Likelihood }}) \underbrace{P(\mathrm{Ob}) P(\mathrm{Po}) P(\mathrm{O} r}_{\text {Prior }})
$$

- The graphical model captures the causal process, by which the observed data was generated (hence the name generative models).


## Discrete Variables

- We now examine the discrete random variables.
- Assume that we have two discrete random variables $x_{1}$ and $x_{2}$, each of which has K states.


$$
p\left(\mathbf{x}_{1}, \mathbf{x}_{2} \mid \boldsymbol{\mu}\right)=\prod_{k=1}^{K} \prod_{l=1}^{K} \mu_{k l}^{x_{1 k} x_{2 l}}
$$

- Using 1-of-K encoding, we denote the probability of observing both $\mathrm{x}_{1 \mathrm{k}}=1$, $\mathrm{x}_{21}=1$ by the parameter $\mu_{k l}$, where $\mathrm{x}_{1 \mathrm{k}}$ denotes the $\mathrm{k}^{\text {th }}$ component of $\mathrm{x}_{1}$ (similarly for $\mathrm{x}_{2}$ ).
- This distribution is governed by $\mathrm{K}^{2}-1$ parameters.
- The total number of parameters that must be specified for an arbitrary joint distribution over M random variables is $\mathrm{K}^{\mathrm{M}}-1$ (corresponds to a fully connected graph).
- Grows exponentially in the number of variables M !


## Discrete Variables

- General joint distribution: K²-1 parameters.


$$
p\left(\mathbf{x}_{1}, \mathbf{x}_{2} \mid \boldsymbol{\mu}\right)=\prod_{k=1}^{K} \prod_{l=1}^{K} \mu_{k l}^{x_{1 k} x_{2 l}}
$$

- Independent joint distribution: 2(K-1) parameters.


$$
\hat{p}\left(\mathbf{x}_{1}, \mathbf{x}_{2} \mid \boldsymbol{\mu}\right)=\prod_{k=1}^{K} \mu_{1 k}^{x_{1 k}} \prod_{l=1}^{K} \mu_{2 l}^{x_{2 l}}
$$

- We dropped the link between the nodes, so each variables is described by a separate multinomial distribution.


## Discrete Variables

- In general:
- Fully connected graphs have completely general distributions and have exponential $\mathrm{K}^{\mathrm{M}}$-1 number of parameters (too complex).
- If there are no links, the joint distribution fully factorizes into the product of the marginals, and has $\mathrm{M}(\mathrm{K}-1)$ parameters (too simple).
- Graphs that have an intermediate level of connectivity allow for more general distributions compared to the fully factorized one, while requiring fewer parameters than the general joint distribution.
- Let us look at the example of the chain graph.


## Chain Graph

- Consider an M-node Markov chain:

- The marginal distribution $p\left(\mathbf{x}_{1}\right)$ requires K-1 parameters.
- The remaining conditional distributions $p\left(\mathbf{x}_{i} \mid \mathbf{x}_{i-1}\right), i=2, \ldots, M$ require $\mathrm{K}(\mathrm{K}-1)$ parameters.
- Total number of parameters: $\mathrm{K}-1+(\mathrm{M}-1)(\mathrm{K}-1) \mathrm{K}$, which is quadratic in $K$ and linear in the length $M$ of the chain.
- This graphical model forms the basis of a simple Hidden Markov Model.


## Adding Priors

- We can turn a graph over discrete random variables into a Bayesian model by introducing Dirichlet priors for the parameters.
- From a graphical model point of view, each node acquires an additional parent representing the Dirichlet distribution over parameters.



## Shared Prior

- We can further share the common prior over the parameters governing the conditional distributions.



## Parameterized Models

- We can use parameterized models to control exponential growth in the number of parameters.

$$
\begin{aligned}
& \text { If } x_{1}, \ldots, x_{M} \text { are discrete, } \mathrm{K} \text {-state } \\
& \text { variables, } p\left(y=1 \mid x_{1}, \ldots, x_{M}\right) \\
& \text { in general has } \mathrm{O}\left(\mathrm{~K}^{\mathrm{M}}\right) \text { parameters. }
\end{aligned}
$$

- We can obtain a more parsimonious form of the conditional distribution by using a logistic function acting on a linear combination of the parent variables:

$$
p\left(y=1 \mid x_{1}, \ldots, x_{M}\right)=\sigma\left(w_{0}+\sum_{i=1}^{M} w_{i} x_{i}\right)=\sigma\left(\mathbf{w}^{\mathrm{T}} \mathbf{x}\right)
$$

- This is a more restricted form of conditional distribution, but it requires only M+1 parameters (linear growth in the number of parameters).


## Linear Gaussian Models

- So far we worked with joint probability distributions over a set of discrete random variables (expressed as nodes in directed acyclic graphs).
- We now show how a multivariate Gaussian distribution can be expressed as a directed graph corresponding to a linear Gaussian model.
- Consider an arbitrary acyclic graph over D random variables, in which each node represent a single continuous Gaussian distribution with its mean given by the linear function of the parents:

$$
p\left(x_{i} \mid \mathrm{pa}_{i}\right)=\mathcal{N}\left(x_{i} \mid \sum_{j \in \mathrm{pa}_{i}} w_{i j} x_{j}+b_{i}, v_{i}\right)
$$

where $\mathrm{w}_{\mathrm{ij}}$ and $\mathrm{b}_{\mathrm{i}}$ are parameters governing the mean, and $\mathrm{v}_{\mathrm{i}}$ is the variance.

## Linear Gaussian Models

- The log of the joint distribution takes form:

$$
\ln p(\mathbf{x})=\sum_{i=1}^{D} \ln p\left(x_{i} \mid \mathrm{pa}_{i}\right)=-\sum_{i=1}^{D} \frac{1}{2 v_{i}}\left(x_{i}-\sum_{j \in \mathrm{pa}_{i}} w_{i j} x_{j}-b_{i}\right)^{2}+\text { const },
$$

where 'const' denotes terms independent of $x$.

- This is a quadratic function of $x$, and hence the joint distribution $p(x)$ is a multivariate Gaussian.
- For example, consider a directed graph over three Gaussian variables with one missing link:



## Computing the Mean

- We can determine the mean and covariance of the joint distribution.

Remember:
hence

$$
p\left(x_{i} \mid \mathrm{pa}_{i}\right)=\mathcal{N}\left(x_{i} \mid \sum_{j \in \mathrm{pa}_{i}} w_{i j} x_{j}+b_{i}, v_{i}\right)
$$

$$
x_{i}=\sum_{j \in \operatorname{pa}_{i}} w_{i j} x_{j}+b_{i}+\sqrt{v_{i}} \epsilon_{i}, \quad \epsilon_{i} \sim \mathcal{N}(0,1),
$$

so its expected value:

$$
\mathbb{E}\left[x_{i}\right]=\sum_{j \in \mathrm{pa}_{i}} w_{i j} \mathbb{E}\left[x_{j}\right]+b_{i}
$$

- Hence we can find components: $\mathbb{E}[\mathbf{x}]=\left[\mathbb{E}\left[x_{1}\right], \ldots, \mathbb{E}\left[x_{D}\right]\right]$ by doing ancestral pass: start at the top and proceed in order (see example):



## Computing the Covariance

- We can obtain the $\mathrm{i}, \mathrm{j}$ element of the covariance matrix in the form of a recursion relation:

$$
\begin{aligned}
\operatorname{cov}\left[x_{i}, x_{j}\right] & =\mathbb{E}\left[\left(x_{i}-\mathbb{E}\left[x_{i}\right]\right)\left(x_{j}-\mathbb{E}\left[x_{j}\right]\right)\right] \\
& =\mathbb{E}\left[\left(x_{i}-\mathbb{E}\left[x_{i}\right]\right)\left(\sum_{k \in \operatorname{pa}_{j}} w_{j k}\left(x_{k}-\mathbb{E}\left[x_{k}\right]\right)+\sqrt{v_{i}} \epsilon_{j}\right)\right] \\
& =\sum_{k \in \mathrm{pa}_{j}} w_{j k} \operatorname{cov}\left[x_{i}, x_{k}\right]+I_{i j} v_{j} .
\end{aligned}
$$

- Consider two cases:
- There are no links in the graph (graph is fully factorized), so that $\mathrm{w}_{\mathrm{ij}}$ 's are zero. In this case: $\mathbb{E}[\mathbf{x}]=\left[b_{1}, \ldots, b_{D}\right]^{T}$, and the covariance is diagonal $\operatorname{diag}\left(v_{1}, \ldots, v_{D}\right)$. The joint distribution represents $D$ independent univariate Gaussian distributions.
- The graph is fully connected. The total number of parameters is $D+D(D-1) / 2$. The covariance corresponds to a general symmetric covariance matrix.


## Bilinear Gaussian Model

－Consider the following model：

$u \sim \mathcal{N}(0,1)$,
$v \sim \mathcal{N}(0,1)$,
$r \sim \mathcal{N}(u v, 1)$.


Gaussian terms

|  | d | d | J | dJ | d |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 |  | ？ | ？ | 市市的 | 市市 |
| 9 | ？ | 办为全 | ＊ | ？ | 亦交 |
| $\bullet$ | 市交方 | ？ | 亦交它 |  | ？ |



$$
\begin{aligned}
& u_{i} \sim \mathcal{N}(0,1), i=1, \ldots, N \\
& v_{j} \sim \mathcal{N}(0,1), j=1, \ldots, M \\
& r_{i j} \sim \mathcal{N}\left(u_{i} v_{j}, 1\right) .
\end{aligned}
$$

－The mean is given by the product of two Gaussians．

## Hierarchical Models



## Conditional Independence

- We now look at the concept of conditional independence.
- $a$ is independent of $b$ given $c$ :

$$
p(a \mid b, c)=p(a \mid c)
$$

- Equivalently:

$$
\begin{aligned}
p(a, b \mid c) & =p(a \mid b, c) p(b \mid c) \\
& =p(a \mid c) p(b \mid c)
\end{aligned}
$$

- We will use the notation:

$$
a \Perp b \mid c
$$

- An important feature of graphical models is that conditional independence properties of the joint distribution can be read directly from the graph without performing any analytical manipulations
- The general framework for achieving this is called d-separation, where d stands for 'directed' (Pearl 1988).


## Example 1: Tail-to-Tail Node

- The joint distribution over three variables can be written:


$$
p(a, b, c)=p(a \mid c) p(b \mid c) p(c)
$$

- If none of the variables are observed, we can examine whether $a$ and $b$ are independent:

$$
p(a, b)=\sum_{c} p(a \mid c) p(b \mid c) p(c)
$$

- In general, this does not factorize into the product $p(a, b)=p(a) p(b)$.

$$
a \not \perp b \mid \emptyset
$$

- $a$ and $b$ have a common cause.
- The node c is said to be tail-to-tail node with respect to this path (the node is connected to the tails of the two arrows).


## Example 1: Tail-to-Tail Node

- Suppose we condition on the variable c:


$$
\begin{aligned}
p(a, b \mid c) & =\frac{p(a, b, c)}{p(c)} \\
& =p(a \mid c) p(b \mid c)
\end{aligned}
$$

- We obtain conditional independence property:

$$
a \Perp b \mid c
$$

- Once chas been observed, $a$ and $b$ can no longer have any effect on each other. They become independent.


## Example 2: Head-to-Tail Node

- The joint distribution over three variables can be written:

- If none of the variables are observed, we can examine whether $a$ and $b$ are independent:

$$
\begin{gathered}
p(a, b)=p(a) \sum_{c} p(c \mid a) p(b \mid c)=p(a) p(b \mid a) \\
a \not \Perp b \mid \emptyset
\end{gathered}
$$

- If c is not observed, a can influence c , and c can influence b .
- The node c is said to be head-to-tail node with respect to the path from node a to node b.


## Example 2: Head-to-Tail Node

- Suppose we condition on the variable c :

- We obtain conditional independence property:

$$
a \Perp b \mid c
$$

- If $c$ is observed, the value of $a$ can no longer influence $b$.


## Example 3: Head-to-Head Node

- The joint distribution over three variables can be written:


$$
p(a, b, c)=p(a) p(b) p(c \mid a, b)
$$

- If none of the variables are observed, we can examine whether $a$ and $b$ are independent:

$$
\begin{gathered}
p(a, b)=p(a) p(b) \\
a \Perp b \mid \emptyset
\end{gathered}
$$

- Opposite to Example 1.
- An unobserved descendant has no effect.
- The node c is said to be head-to-head node with respect to the path from a to $b$ (because it connects to the heads of two arrows).


## Example 3: Head-to-Head Node

- Suppose we condition on the variable c:


$$
\begin{aligned}
p(a, b \mid c) & =\frac{p(a, b, c)}{p(c)} \\
& =\frac{p(a) p(b) p(c \mid a, b)}{p(c)}
\end{aligned}
$$

- In general, this does not factorize into the product.

$$
a \not \perp b \mid c
$$

- Opposite to Example 1.
- If the descendant (or any of its descendants) is observed, its value has implications for both a and b,


## Fuel Example

- Consider the following example over three binary random variables:


$$
\begin{gathered}
p(B=1)=0.9 \\
p(F=1)=0.9 \\
\text { and hence } \\
p(F=0)=0.1
\end{gathered}
$$

$B=$ Battery ( $0=$ dead, $1=$ fully charged)
F = Fuel Tank ( $0=$ empty, $1=$ full)
G = Fuel Gauge Reading
(0=empty, 1=full)

$$
\begin{gathered}
p(G=1 \mid B=1, F=1)=0.8 \\
p(G=1 \mid B=1, F=0)=0.2 \\
p(G=1 \mid B=0, F=1)=0.2 \\
p(G=1 \mid B=0, F=0)=0.1
\end{gathered}
$$

## Fuel Example

- Suppose that we observe that the Fuel Gauge Reading is empty $\mathrm{G}=0$.

$$
\begin{aligned}
p(F=0 \mid G=0) & =\frac{p(G=0 \mid F=0) p(F=0)}{p(G=0)} \\
& \simeq 0.257
\end{aligned}
$$



- Probability of an empty tank increased by observing $G=0$.
$B=$ Battery ( $0=$ dead, $1=$ fully charged)
F = Fuel Tank ( $0=$ empty, $1=$ full)
G = Fuel Gauge Reading (0=empty, 1=full)


## Explaining Away

- If we observe that the Fuel Gauge Reading is empty G = 0 and that the battery is dead $\mathrm{B}=0$.

$$
\begin{aligned}
p(F=0 \mid G=0, B=0) & =\frac{p(G=0 \mid B=0, F=0) p(F=0)}{\sum_{F \in\{0,1\}} p(G=0 \mid B=0, F) p(F)} \\
& \simeq 0.111
\end{aligned}
$$



- Probability of an empty tank $\mathrm{F}=0$ is reduced by observing that the battery is dead $\mathrm{B}=0$.
- If we observe that the fuel gauge reading is empty, you assume that one of the causes happen (either the battery
$B=$ Battery ( $0=$ dead, $1=$ fully charged)
F = Fuel Tank (0=empty, 1=full)
G = Fuel Gauge Reading (0=empty, 1=full)
is dead or the fuel tank is empty).
- One cause removes `explains away’ the need for the other cause.


## D-separation



- $a$ is independent of $b$ if and only if all paths connecting $a$ and $b$ are blocked.
- head-to-tail and tail-to-tail nodes are blocked when observed.
- head-to-head nodes are blocked when the node and all its descendants are unobserved.
- For example (on top), the path from a to $b$ is not blocked by f because it is tail-to-tail node and it is unobserved.
- But conditioned on f, a and b become independent.


## D-separation and i.i.d data



- Another example of conditional independence and d-separation is provided by the concept of independent and identically distributed data.
- Consider the problem of finding the posterior distribution over mean $\mu$ in Bayesian linear regression model.
- Suppose that we condition on $\mu$ and consider the joint over observed variables.
- Using d-separation, note that there is unique path from $x_{i}$ to any other $x_{j}$, and this path is head-to-head with respect to $\mu$.
- If we integrate out $\mu$, the observations
$p(\mathcal{D})=\int_{-\infty}^{\infty} p(\mathcal{D} \mid \mu) p(\mu) \mathrm{d} \mu \neq \prod_{n=1}^{N} p\left(x_{n}\right)^{\text {are no longer independent. }}$


## Markov Blanket in Directed Models

- The Markov blanket of a node is the minimal set of nodes that must be observed to make this node independent of all other nodes
- In a directed model, the Markov blanket includes parents, children and co-parents (i.e. all the parents of the node's children) due to explaining away.


Factors independent of $x_{i}$ cancel between numerator and denominator.

## Directed Graphs as Distribution Filters

- We can view the graphical model as a filter.

- The joint probability distribution $p(x)$ is allowed through the filter if and only if it satisfies the factorization property.
- Note: The fully connected graph exhibits no conditional independence properties at all.
- The fully disconnected graph (no links) corresponds to a joint distribution that factorizes into the product of marginal distributions.


## Popular Models

Latent Dirichlet Allocation


- One of the popular models for modeling word count vectors. We will see this model later.

Bayesian Probabilistic Matrix Factorization


- One of the popular models for collaborative filtering applications.


## Undirected Graphical Models

Directed graphs are useful for expressing causal relationships between random variables, whereas undirected graphs are useful for expressing soft constraints between random variables

- The joint distribution defined by the graph is given by
 the product of non-negative potential functions over the maximal cliques (connected subset of nodes).

$$
p(\mathbf{x})=\frac{1}{\mathcal{Z}} \prod_{C} \phi_{C}\left(x_{C}\right) \quad \mathcal{Z}=\sum_{\mathbf{x}} \prod_{C} \phi_{C}\left(x_{C}\right)
$$

where the normalizing constant $\mathcal{Z}$ is called a partition function.

- For example, the joint distribution factorizes:

$$
p(A, B, C, D)=\frac{1}{\mathcal{Z}} \phi(A, C) \phi(C, B) \phi(B, D) \phi(A, D)
$$

- Let us look at the definition of cliques.


## Cliques

- The subsets that are used to define the potential functions are represented by maximal cliques in the undirected graph.
- Clique: a subset of nodes such that there exists a link between all pairs of nodes in a subset.
- Maximal Clique: a clique such that it is not possible to include any other nodes in the set without it ceasing to be a clique.
- This graph has 5 cliques:

$$
\begin{aligned}
& \left\{x_{1}, x_{2}\right\},\left\{x_{2}, x_{3}\right\},\left\{x_{3}, x_{4}\right\} \\
& \left\{x_{4}, x_{2}\right\},\left\{x_{1}, x_{3}\right\}
\end{aligned}
$$



- Two maximal cliques:

$$
\left\{x_{1}, x_{2}, x_{3}\right\},\left\{x_{2}, x_{3}, x_{4}\right\}
$$

## Using Cliques to Represent Subsets

- If the potential functions only involve two nodes, an undirected graph has a nice representation.
- If the potential functions involve more than two nodes, using a different factor graph representation is much more useful.
- For now, let us consider only potential functions that are defined over two nodes.



## Markov Random Fields (MRFs)



$$
p(\mathbf{x})=\frac{1}{\mathcal{Z}} \prod_{C} \phi_{C}\left(x_{C}\right)
$$

- Each potential function is a mapping from the joint configurations of random variables in a clique to non-negative real numbers.
- The choice of potential functions is not restricted to having specific probabilistic interpretations.

Potential functions are often represented as exponentials:

$$
p(\mathbf{x})=\frac{1}{\mathcal{Z}} \prod_{C} \phi_{C}\left(x_{C}\right)=\frac{1}{\mathcal{Z}} \exp \left(-\sum_{C} E\left(x_{c}\right)\right)=\underbrace{\frac{1}{\mathcal{Z}} \exp (-E(\mathbf{x}))}
$$

where $E(x)$ is called an energy function.
Boltzmann distribution

## MRFs with Hidden Variables

For many interesting real-world problems, we need to introduce hidden or latent variables.


- Our random variables will contain both visible and hidden variables $\mathrm{x}=(\mathrm{v}, \mathrm{h})$.

$$
p(\mathbf{v})=\frac{1}{\mathcal{Z}} \sum_{\mathbf{h}} \exp (-E(\mathbf{v}, \mathbf{h}))
$$

- In general, computing both partition function and summation over hidden variables will be intractable, except for special cases.
- Parameter learning becomes a very challenging task.


## Conditional Independence

- Conditional Independence is easier compared to directed models:

- Observation blocks a node.
- Two sets of nodes are conditionally independent if the observations block all paths between them.


## Markov Blanket

- The Markov blanket of a node is simply all of the directly connected nodes.

Markov Blanket


- This is simpler than in directed models, since there is no explaining away.
- The conditional distribution of $x_{i}$ conditioned on all the variables in the graph is dependent only on the variables in the Markov blanket.


## Conditional Independence and Factorization

- Consider two sets of distributions:
- The set of distributions consistent with the conditional independence relationships defined by the undirected graph.
- The set of distributions consistent with the factorization defined by potential functions on maximal cliques of the graph.
- The Hammersley-Clifford theorem states that these two sets of distributions are the same.


$$
p(\mathbf{x})=\frac{1}{\mathcal{Z}} \prod_{C} \phi_{C}\left(x_{C}\right)
$$

## Interpreting Potentials

- In contrast to directed graphs, the potential functions do not have a specific probabilistic interpretation.


$$
p(\mathbf{x})=\frac{1}{\mathcal{Z}} \prod_{C} \phi_{C}\left(x_{C}\right)=\frac{1}{\mathcal{Z}} \exp \left(-\sum_{C} E\left(x_{c}\right)\right)
$$

- This gives us greater flexibility in choosing the potential functions.
- We can view the potential function as expressing which configuration of the local variables are preferred to others.
- Global configurations with relatively high probabilities are those that find a good balance in satisfying the (possibly conflicting) influences of the clique potentials.
- So far we did not specify the nature of random variables, discrete or continuous.


## Discrete MRFs

- MRFs with all discrete variables are widely used in many applications.
- MRFs with binary variables are sometimes called Ising models in statistical mechanics, and Boltzmann machines in machine learning

- Denoting the binary valued variable at node $j$ by $x_{j} \in\{0,1\}$, the Ising model for the joint probabilities is given by:

$$
P_{\theta}(\mathbf{x})=\frac{1}{\mathcal{Z}(\theta)} \exp \left(\sum_{i j \in E} x_{i} x_{j} \theta_{i j}+\sum_{i \in V} x_{i} \theta_{i}\right)
$$

- The conditional distribution is given by logistic:

$$
P_{\theta}\left(x_{i}=1 \mid \mathbf{x}_{-i}\right)=\frac{1}{1+\exp \left(-\theta_{i}-\sum_{i j \in E} x_{j} \theta_{i j}\right)}
$$

where $x_{-i}$ denotes all nodes except for $i$.

Hence the parameter $\theta_{\mathrm{ij}}$ measures the dependence of $\mathrm{x}_{\mathrm{i}}$ on $\mathrm{x}_{\mathrm{j}}$, conditional on the other nodes.

## Example: Image Denoising

- Let us look at the example of noise removal from a binary image.
- Let the observed noisy image be described by an array of binary pixel values: $y_{j} \in\{-1,+1\}, \mathrm{i}=1, \ldots, \mathrm{D}$.
- We take a noise-free image $x_{j} \in\{-1,+1\}$,
 and randomly flip the sign of pixels with some small probability.

Neighboring pixels

Bias term are likely to have the


$$
p(\mathbf{x}, \mathbf{y})=\frac{1}{Z} \exp \{-E(\mathbf{x}, \mathbf{y})\}
$$

## Iterated Conditional Modes

- Iterated conditional modes: coordinate-wise gradient descent.
- Visit the unobserved nodes sequentially and set each $x$ to whichever of its two values has the lowest energy.
- This only requires us to look at the Markov blanket, i.e. the connected nodes.
- Markov blanket of a node is simply all of the directly connected nodes.


Original Image


Noisy Image


ICM

## Gaussian MRFs

- We assume that the observations have a multivariate Gaussian distribution with mean $\mu$ and covariance matrix $\Sigma$.

$$
\mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}, \boldsymbol{\Sigma})=\frac{1}{(2 \pi)^{D / 2}} \frac{1}{|\boldsymbol{\Sigma}|^{1 / 2}} \exp \left\{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right\}
$$

- Since the Gaussian distribution represents at most second-order relationships, it automatically encodes a pairwise MRF. We rewrite:


$$
P(\mathbf{x})=\frac{1}{\mathcal{Z}} \exp \left(-\frac{1}{2} \mathbf{x}^{T} J \mathbf{x}+\mathbf{g}^{T} \mathbf{x}\right)
$$

where

$$
J=\Sigma^{-1}, \quad \mu=J^{-1} \mathbf{g}
$$

- The positive definite matrix $J$ is known as the information matrix and is sparse with respect to the given graph: $\mathbf{x}^{T} J \mathbf{x}=\sum_{i} J_{i i} x_{i}^{2}+2 \sum_{i j \in E} J_{i j} x_{i} x_{j}$,
if $(i, j) \neq E$, then $J_{i j}=0$.
- The information matrix is sparse, but the covariance matrix is not sparse.


## Restricted Boltzmann Machines

- For many real-world problems, we need to introduce hidden variables.
- Our random variables will contain visible and hidden variables $\mathrm{x}=(\mathrm{v}, \mathrm{h})$.


Stochastic binary visible variables $\mathbf{v} \in\{0,1\}^{D}$ are connected to stochastic binary hidden variables $\mathbf{h} \in\{0,1\}^{F}$.

The energy of the joint configuration:

$$
E(\mathbf{v}, \mathbf{h} ; \theta)=-\sum_{i j} W_{i j} v_{i} h_{j}-\sum_{i} b_{i} v_{i}-\sum_{j} a_{j} h_{j}
$$

$\theta=\{W, a, b\}$ model parameters.
Probability of the joint configuration is given by the Boltzmann distribution:

$$
\begin{aligned}
& P_{\theta}(\mathbf{v}, \mathbf{h})=\frac{1}{\mathcal{Z}(\theta)} \exp (-E(\mathbf{v}, \mathbf{h} ; \theta))=\underbrace{\frac{1}{\mathcal{Z}(\theta)}}_{\text {partition function }} \prod_{\text {ij }} \underbrace{e^{W_{i j} v_{i} h_{j}}}_{\text {potential functions }} \prod_{i} e^{b_{i} v_{i}} \prod_{j} e^{a_{j} h_{j}} \\
& \mathcal{Z}(\theta)=\sum_{\mathbf{h}, \mathbf{v}} \exp (-E(\mathbf{v}, \mathbf{h} ; \theta))
\end{aligned}
$$

## Restricted Boltzmann Machines



Restricted: No interaction between hidden variables

Inferring the distribution over the hidden variables is easy:

$$
P(\mathbf{h} \mid \mathbf{v})=\underbrace{\prod_{j} P\left(h_{j} \mid \mathbf{v}\right)}_{\text {Similarly: }} P\left(h_{j}=1 \mid \mathbf{v}\right)=\frac{1}{1+\exp \left(-\sum_{i} W_{i j} v_{i}-a_{j}\right)}
$$

$$
P(\mathbf{v} \mid \mathbf{h})=\prod_{i} P\left(v_{i} \mid \mathbf{h}\right) \quad P\left(v_{i}=1 \mid \mathbf{h}\right)=\frac{1}{1+\exp \left(-\sum_{j} W_{i j} h_{j}-b_{i}\right)}
$$

Markov random fields, Boltzmann machines, log-linear models.

## Restricted Boltzmann Machines

Observed Data
Subset of 25,000 characters


New Image: $\quad p\left(h_{7}=1 \mid v\right)$

$$
\begin{array}{r}
\leftrightarrows=\sigma(0.99 \times \\
\\
\sigma(x)=\frac{1}{1+\exp (-x)}
\end{array}
$$

Learned W: "edges"
Subset of 1000 features

$$
p\left(h_{29}=1 \mid v\right)
$$

Logistic Function: Suitable for modeling binary images

$$
\text { as } \quad P(\mathbf{h} \mid \mathbf{v})=[0,0,0.82,0,0,0.99,0,0 \ldots]
$$

## Gaussian-Bernoulli RBMs

## Gaussian-Bernoulli RBM:



$$
P_{\theta}(\mathbf{v}, \mathbf{h})=\frac{1}{\mathcal{Z}(\theta)} \exp (-E(\mathbf{v}, \mathbf{h} ; \theta))
$$

Define energy functions for various data modalities:

$$
\begin{array}{ll}
P\left(v_{i}=x \mid \mathbf{h}\right)=\frac{1}{\sqrt{2 \pi} \sigma_{i}} \exp \left(-\frac{\left(x-b_{i}-\sigma_{i} \sum_{j} W_{i j} h_{j}\right)^{2}}{2 \sigma_{i}^{2}}\right) & \text { Gaussian } \\
P\left(h_{j}=1 \mid \mathbf{v}\right)=\frac{1}{1+\exp \left(-\sum_{i} W_{i j} \frac{v_{i}}{\sigma_{i}}-a_{j}\right)} & \text { Bernoulli }
\end{array}
$$

## Gaussian-Bernoulli RBMs

## Images: Gaussian-Bernoulli RBM

4 million unlabelled images
Learned features (out of 10,000 )


Text: Multinomial-Bernoulli RBM

REUTERS: :
AP Associated Press

stock
wall street point dow

## Relation to Directed Graphs

- Let us try to convert directed graph into an undirected graph:



## Directed vs. Undirected

- Directed Graphs can be more precise about independencies than undirected graphs.

- All the parents of $x_{4}$ can interact to determine the distribution over $\mathrm{x}_{4}$.
- The directed graph represents independencies that the undirected graph cannot model.
- To represent the high-order interaction in the directed graph, the undirected graph needs a fourth-order clique.
- This fully connected graph exhibits no conditional independence properties


## Undirected vs. Directed

- Undirected Graphs can be more precise about independencies than directed graphs
- There is no directed graph over four variables that represents the same set of conditional independence properties.


$$
\begin{gathered}
A \not \Perp B \mid \emptyset \\
A \Perp B \mid C \cup D \\
C \Perp D \mid A \cup B
\end{gathered}
$$

## Directed vs. Undirected

- If every conditional independence property of the distribution is reflected in the graph and vice versa, then the graph is a perfect map for that distribution.

- Venn diagram:
- The set of all distributions P over a given set of random variables.
- The set of distributions D that can be represented as a perfect map using directed graph.
- The set of distributions $U$ that can be represented as a perfect map using undirected graph.
- We can extend the framework to graphs that include both directed and undirected graphs.

