CSC321 Lecture 18: Mixture Modeling

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Overview

• Some examples of situations where you'd use unupservised learning

- You want to understand how a scientific field has changed over time. You want to take a large database of papers and model how the distribution of topics changes from year to year. But what are the topics?
- You're a biologist studying animal behavior, so you want to infer a high-level description of their behavior from video. You don't know the set of behaviors ahead of time.
- You want to reduce your energy consumption, so you take a time series of your energy consumption over time, and try to break it down into separate components (refrigerator, washing machine, etc.).
- Common theme: you have some data, and you want to infer the causal structure underlying the data.
- This structure is latent, which means it's never observed.

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Overview

- In last lecture, we looked at density modeling where all the random variables were fully observed.
- The more interesting case is when some of the variables are latent, or never observed. These are called latent variable models.
 - Today's lecture: mixture models, where the latent variable comes from a small discrete set
 - Next week: latent variable models which have distributed representations these are much more powerful

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Clustering

• Sometimes the data form clusters, where examples within a cluster are similar to each other, and examples in different clusters are dissimilar:



- Such a distribution is multimodal, since it has multiple modes, or regions of high probability mass.
- Grouping data points into clusters, with no labels, is called clustering
- E.g. clustering machine learning papers based on topic (deep learning, Bayesian models, etc.)
 - This is an overly simplistic model more on that later

K-Means

- First, let's look at a simple clustering algorithm, called k-means.
- This is an iterative algorithm. In each iteration, we keep track of:
 - An assignment of data points to clusters
 - The center of each cluster
- Start with random cluster locations, then alternate between:
 - Assignment step: assign each data point to the nearest cluster
 - Refitting step: move each cluster center to the average of its data points



K-Means

- Each iteration can be shown to decrease a particular cost function: the sum of squared distances from data points to their corresponding cluster centers.
 - More on this in CSC411.
- Problem: what if the clusters aren't spherical?



- Let's instead treat clustering as a distribution modeling problem.
 - Last lecture, we fit Gaussian distributions to data.
 - To model multimodal distributions, let's fit a mixture model, where each data point belongs to a different component.
 - E.g., in a mixture of Gaussians, each data point comes from one of several different Gaussian distributions.
 - We don't need to use Gaussians we can pick whatever distribution best represents our data.

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Mixture of Gaussians

 In a mixture model, we define a generative process where we first sample the latent variable z, and then sample the observations x from a distribution which depends on z.

$$p(z,\mathbf{x}) = p(z) p(\mathbf{x} | z).$$

• E.g. mixture of Gaussians:

 $z \sim \text{Multinomial}(0.7, 0.3) \tag{1}$ $x \mid z = 1 \sim \text{Gaussian}(0, 1) \tag{2}$ $x \mid z = 2 \sim \text{Gaussian}(6, 2) \tag{3}$

• The probabilities used to sample z are called the mixing proportions.

Mixture of Gaussians





• The probability density function over x is defined by marginalizing, or summing out, z:

$$p(\mathbf{x}) = \sum_{k=1}^{K} \Pr(z=k) \, p(\mathbf{x} \mid z=k)$$

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- Suppose we know the model parameters (mixture probabilities and component means and variances)
- In posterior inference, we infer the posterior over z using Bayes' Rule:

$$p(z \mid \mathbf{x}) \propto p(z) p(\mathbf{x} \mid z).$$

ullet For a univariate Gaussian mixture with mixing proportions π

$$p(z=1 \mid x) = \frac{\pi_1 \cdot \mathcal{N}(x; \mu_1, \sigma_1)}{\pi_1 \cdot \mathcal{N}(x; \mu_1, \sigma_1) + \pi_2 \cdot \mathcal{N}(x; \mu_2, \sigma_2)}$$

Posterior Inference

Example:



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Posterior Inference

- Sometimes the observables aren't actually observed then we say they're missing
- One use of probabilistic models is to make predictions about missing data
 - E.g. image completion, which you'll do in Assignment 4
- Analogously to Bayesian parameter estimation, we use the posterior predictive distribution:

$$p(x_2 \mid x_1) = \sum_{z} \underbrace{p(z \mid x_1)}_{\text{posterior}} p(x_2 \mid z, x_1).$$

• If the dimensions of **x** are conditionally independent given *z*, this is just a reweighting of the original mixture model, where we use the posterior rather than the prior.

$$p(x_2 \mid x_1) = \sum_{z} \underbrace{p(z \mid x_1)}_{\text{posterior component PDF}} \underbrace{p(x_2 \mid z)}_{\text{component PDF}}.$$

Posterior Inference

Example:



Fully worked-through example in the lecture notes.

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Parameter Learning

- Now let's talk about learning. We need to fit two sets of paramters:
 - The mixture probabilities $\pi_k = \Pr(z = k)$
 - The mean μ_k and standard deviation σ_k for each component
- If someone hands us the values of all the latent variables, it's easy to fit the parameters using maximum likelihood.

$$\begin{aligned} \ell &= \log \prod_{i=1}^{N} p(z^{(i)}, x^{(i)}) \\ &= \log \prod_{i=1}^{N} p(z^{(i)}) p(x^{(i)} \mid z^{(i)}) \\ &= \sum_{i=1}^{N} \underbrace{\log p(z^{(i)})}_{\pi} + \underbrace{\log p(x^{(i)} \mid z^{(i)})}_{\mu_{k}, \sigma_{k}} \end{aligned}$$

Parameter Learning

- Let $r_k^{(i)}$ be the indicator variable for $z^{(i)} = k$. This is called the responsibility
- Solving for the mixing probabilities:

$$\ell = \sum_{i=1}^{N} \log p(z^{(i)}) + \log p(x^{(i)} | z^{(i)})$$
$$= \text{const} + \sum_{i=1}^{N} \log p(z^{(i)})$$

• This is just the maximum likelihood problem for the multinomial distirbution. The solution is just the empirical proabilities, which we can write as:

$$\pi_k \leftarrow \frac{1}{N} \sum_{i=1}^N r_k^{(i)}$$

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Parameter Learning

• Solving for the mean parameter μ_k for component k:

$$\ell = \sum_{i=1}^{N} \log p(z^{(i)}) + \log p(x^{(i)} | z^{(i)})$$

= const + $\sum_{i=1}^{N} \log p(x^{(i)} | z^{(i)})$
= const + $\sum_{i=1}^{N} r_k^{(i)} \log \mathcal{N}(x^{(i)}; \mu_k, \sigma_k)$

- This is just maximum likelihood for the parameters of a Gaussian distribution, where only certain data points count.
- Solution:

$$\mu_k \leftarrow \frac{\sum_{i=1}^N r_k^{(i)} x^{(i)}}{\sum_{i=1}^N r_k^{(i)}}$$

- We've seen how to do two things:
 - Given the model parameters, compute the posterior over the latent variables
 - Given the latent variables, find the maximum likelihood parameters
- But we don't know the parameters or latent variables, so we have a chicken-and-egg problem.
- Remember k-means? We iterated between an assignment step and a refitting step.
- Expectation-Maximization (E-M) is an analogous procedure which alternates bewteen two steps:
 - Expectation step (E-step): Compute the posterior expectations of the latent variables z
 - **Maximization step (M-step):** Solve for the maximum likelihood parameters given the full set of x's and z's

E-step:

• This is like the assignment step in k-means, except that we assign fractional responsibilities.

$$egin{aligned} r_k^{(i)} &\leftarrow \Pr(z^{(i)} = k \,|\, x^{(i)}) \ &\propto \pi_k \cdot \mathcal{N}(x^{(i)}; \mu_k, \sigma_k) \end{aligned}$$

• This is just posterior inference, which we've already talked about.

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M-step:

• Maximum likelihood with fractional counts:

$$\boldsymbol{\theta} \leftarrow \arg \max_{\boldsymbol{\theta}} \sum_{i=1}^{N} \sum_{k=1}^{K} r_k^{(i)} \left[\log \Pr(z^{(i)} = k) + \log p(\mathbf{x}^{(i)} \mid z^{(i)} = k) \right]$$

• The maximum likelihood formulas we already saw don't depend on the responsibilities being 0 or 1. They also work with fractional responsibilities. E.g.,

$$\pi_k \leftarrow \frac{1}{N} \sum_{i=1}^N r_k^{(i)}$$
$$\mu_k \leftarrow \frac{\sum_{i=1}^N r_k^{(i)} x^{(i)}}{\sum_{i=1}^N r_k^{(i)}}$$

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- We initialize the model parameters randomly and then repeatedly apply the E-step and M-step.
- Each step can be shown to increase the log-likelihood, but this is beyond the scope of the class.
 - Optional mathematical justification in the lecture notes, in case you're interested.
 - Also, there's a full explanation in CSC 411.
 - Next lecture, we'll fit a different latent variable model by doing gradient descent on the parameters. This will turn out to have an EM-like flavor.

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 Suppose we recorded a bunch of temperatures in March for Toronto and Miami, but forgot to record which was which, and they're all jumbled together.



Let's try to separate them out using a mixture of Gaussians and E-M.

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Step 1:





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Step 2:







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Step 3:





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Step 10:







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- We used univariate Gaussian components for simplicity, but other distributions are possible.
- Multivariate Gaussians:



 In Programming Assignment 4, you will fit a mixture of Bernoullis model.

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Odds and Ends

- E-M can get stuck in local optima.
 - Initialize from k-means, which can be more robust in practice
 - Use multiple random restarts and pick the one with the best objective function
- Mixture models are a localist representation: the latent variables take values in a small discrete set.
 - We can use more complex distributions over latent variables to get a distributed representation.
 - The difficulty is posterior inference: while this is easy to do exactly for mixture models, it's intractable in general, and we'll need to make approximations.

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