

Inductive Principles for Restricted Boltzmann Machine Learning

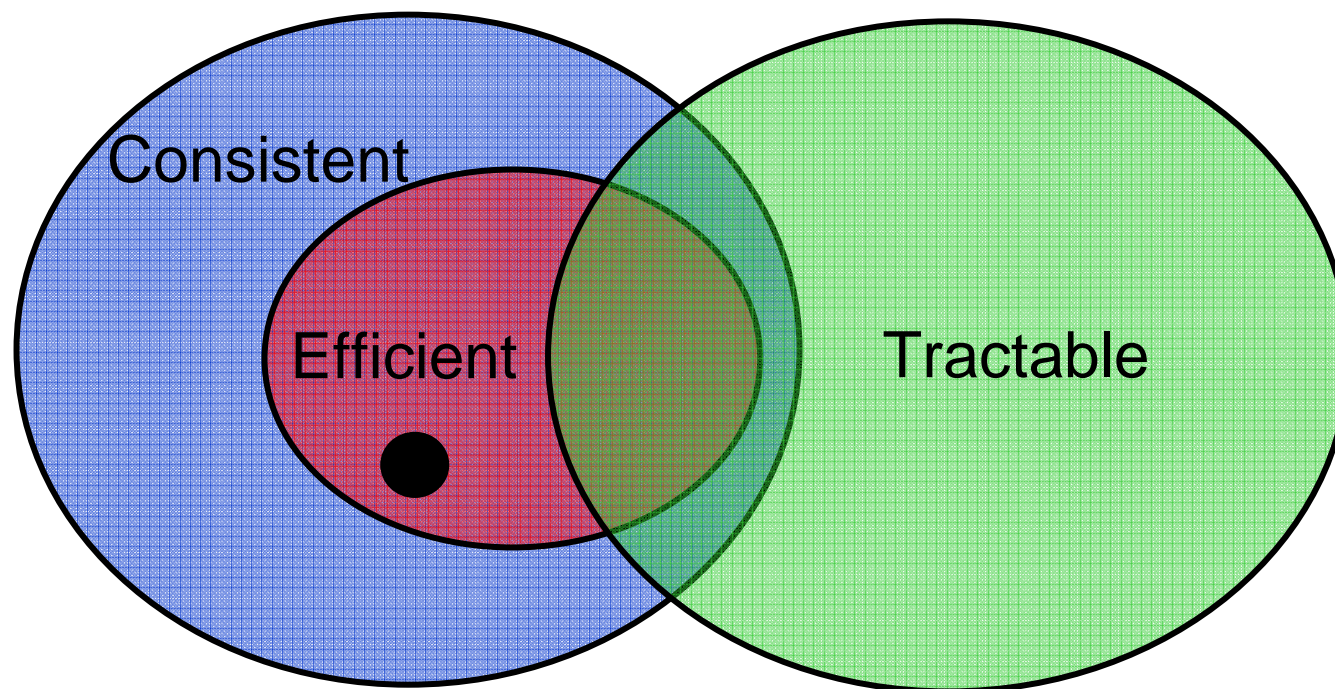
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Joint work with Kevin Swersky, Bo Chen and Nando de Freitas

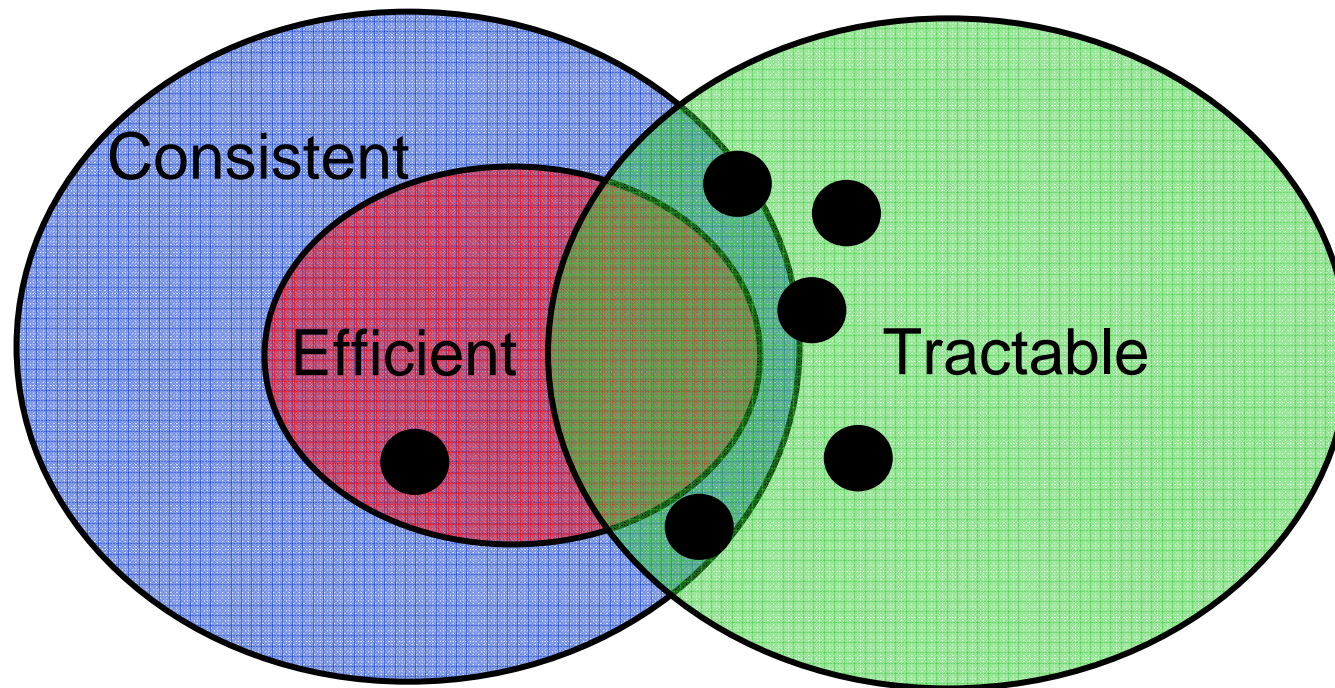
Introduction: Maximum Likelihood

- Maximum Likelihood Estimation is statistically consistent and efficient but is not computationally tractable for many models of interest like RBM's, MRF's, CRF's due to the partition function.



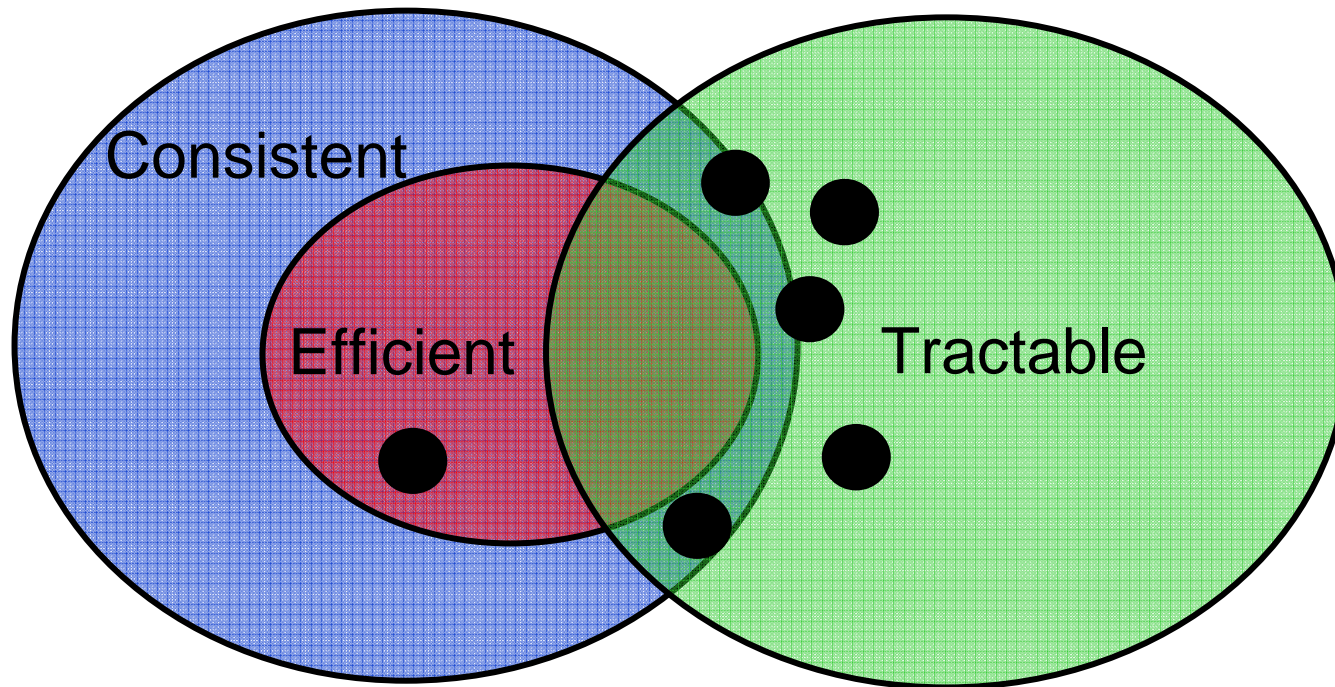
Introduction: Alternative Estimators

- Recent work has seen the proposal of many new estimators that trade consistency/efficiency for computational tractability including RM, SM, GSM, MPF, NCE, NLCE.



Introduction: Alternative Estimators

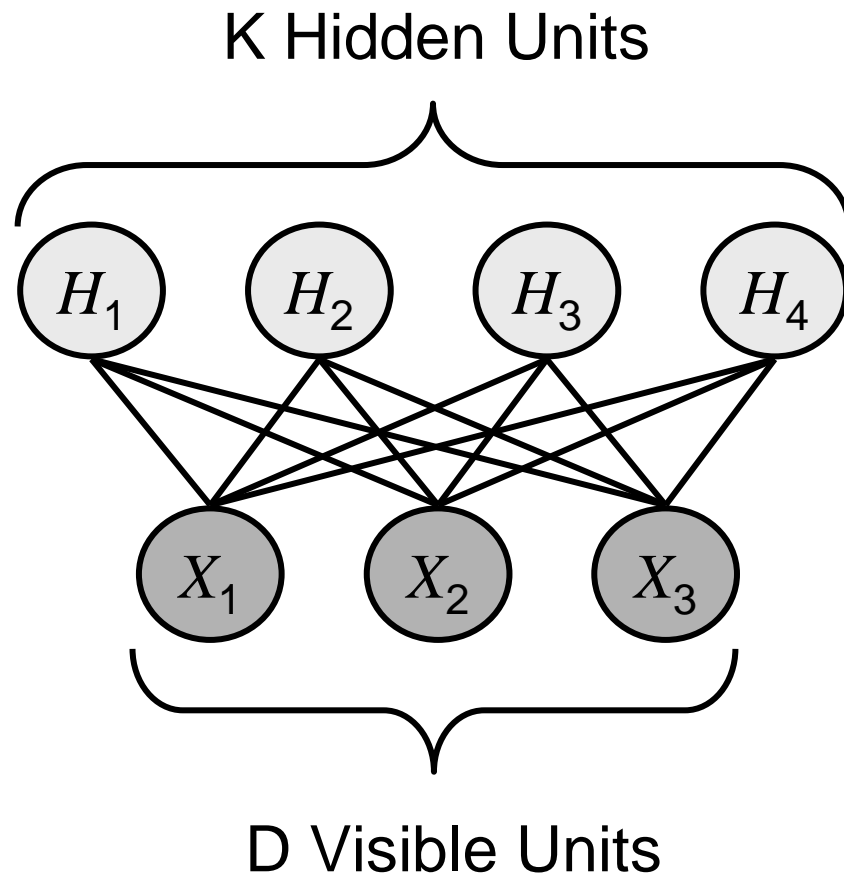
- Our main interest is uncovering the relationships between these estimators and studying their theoretical and empirical properties.



Outline:

- Boltzmann Machines and RBMs
- Inductive Principles
 - Maximum Likelihood
 - Contrastive Divergence
 - Pseudo-Likelihood
 - Ratio Matching
 - Generalized Score Matching
 - Minimum Probability Flow
- Experiments
- Demo

Introduction: Restricted Boltzmann Machines



- A Restricted Boltzmann Machine (RBM) is a Boltzmann Machine with a bipartite graph structure.
- Typically one layer of nodes are fully observed variables (the visible layer), while the other consists of latent variables (the hidden layer).

Introduction: Restricted Boltzmann Machines

- The joint probability of the visible and hidden variables is defined through a bilinear energy function.

$$E_{\theta}(\mathbf{x}, \mathbf{h}) = -(\mathbf{x}^T W \mathbf{h} + \mathbf{x}^T \mathbf{b} + \mathbf{h}^T \mathbf{c})$$

$$P_{\theta}(\mathbf{x}, \mathbf{h}) = \frac{1}{Z} \exp(-E_{\theta}(\mathbf{x}, \mathbf{h}))$$

$$Z = \sum_{\mathbf{x}' \in \mathcal{X}} \sum_{\mathbf{h}' \in \mathcal{H}} \exp(-E_{\theta}(\mathbf{x}', \mathbf{h}'))$$

Introduction: Restricted Boltzmann Machines

- The bipartite graph structure gives the RBM a special property: the visible variables are conditionally independent given the hidden variables and vice versa.

$$P_{\theta}(x_d = 1 | \mathbf{h}) = \frac{1}{1 + \exp(-(\sum_{k=1}^K W_{dk} h_k + x_d b_d))}$$
$$P_{\theta}(h_k = 1 | \mathbf{x}) = \frac{1}{1 + \exp(-(\sum_{d=1}^D W_{dk} x_d + h_k c_k))}$$

Introduction: Restricted Boltzmann Machines

- The marginal probability of the visible vector is obtained by summing out over all joint states of the hidden variables.

$$P_{\theta}(\mathbf{x}) = \frac{1}{\mathcal{Z}} \sum_{\mathbf{h} \in \mathcal{H}} \exp(-E_{\theta}(\mathbf{x}, \mathbf{h}))$$

$$P_{\theta}(\mathbf{x}) = \frac{1}{\mathcal{Z}} \exp(-F_{\theta}(\mathbf{x}))$$

$$F_{\theta}(\mathbf{x}) = - \left(\mathbf{x}^T \mathbf{b} + \sum_{k=1}^K \log \left(1 + \exp(\mathbf{x}^T \mathbf{W}_k + c_k) \right) \right)$$

Introduction: Restricted Boltzmann Machines

- This construction eliminates the latent, hidden variables, leaving a distribution defined in terms of the visible variables.
- However, computing the normalizing constant (partition function) still has exponential complexity in D .

$$Z = \sum_{\mathbf{x}' \in \mathcal{X}} \exp \left(-F_{\theta}(\mathbf{x}') \right)$$

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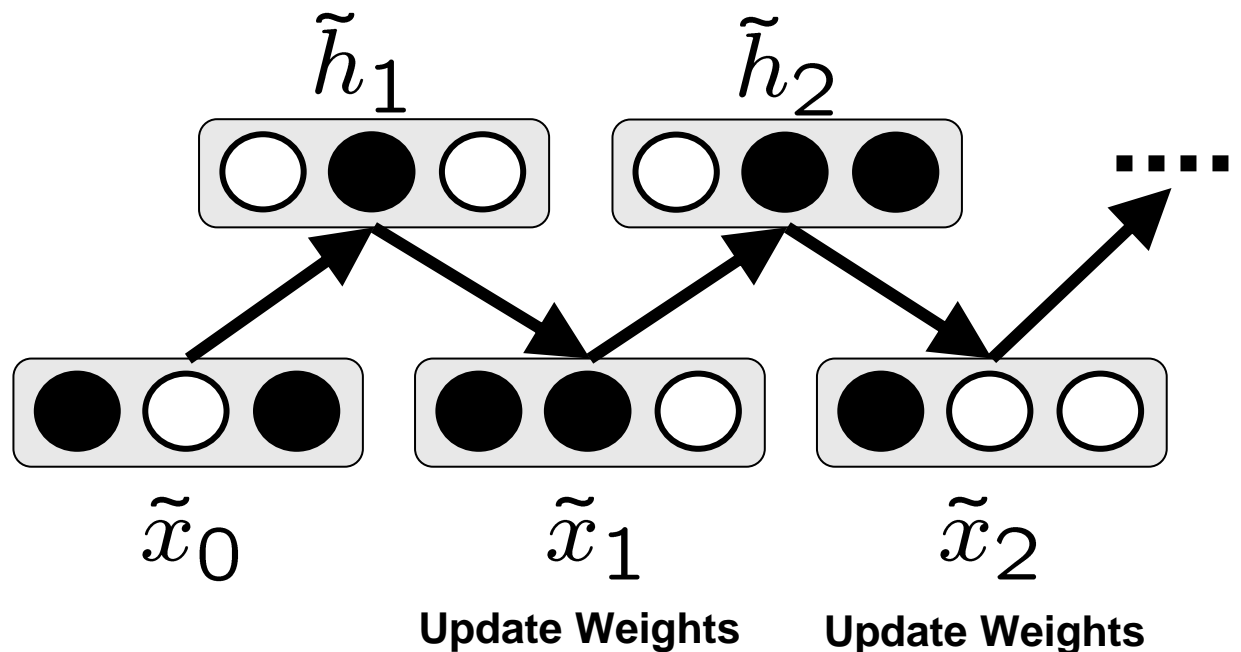
Stochastic Maximum Likelihood

- Exact maximum likelihood learning is intractable in an RBM. Stochastic ML estimation can instead be applied, usually using a simple block Gibbs sampler.

$$f^{ML}(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} P_e(\mathbf{x}) \log P_\theta(\mathbf{x})$$

Stochastic Maximum Likelihood

- Exact maximum likelihood learning is intractable in an RBM. Stochastic ML estimation can instead be applied, usually using a simple block Gibbs sampler.



- L. Younes. Parametric inference for imperfectly observed Gibbsian fields. *Prob. Th. and Related Fields*, 82(4):625–645, 1989.
- T. Tieleman. Training restricted Boltzmann machines using approximations to the likelihood gradient. *ICML 25*, 2008.

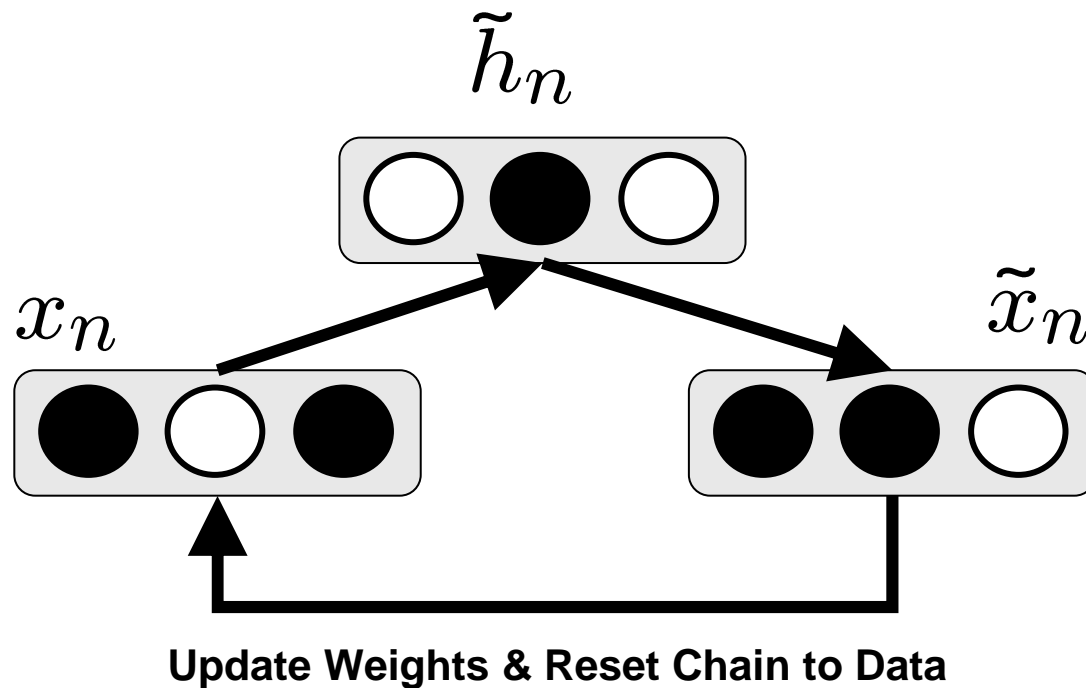
Contrastive Divergence

- The contrastive divergence principle results in a gradient that looks identical to stochastic maximum likelihood. The difference is that CD samples from the T-step Gibbs distribution.

$$f^{CD}(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} P_e(\mathbf{x}) \log \left(\frac{P_e(\mathbf{x})}{P_\theta(\mathbf{x})} \right) - Q_\theta^t(\mathbf{x}) \log \left(\frac{Q_\theta^t(\mathbf{x})}{P_\theta(\mathbf{x})} \right)$$

Contrastive Divergence

- The contrastive divergence principle results in a gradient that looks identical to stochastic maximum likelihood. The difference is that CD samples from the T-step Gibbs distribution.



Pseudo-Likelihood

- The principle of maximum pseudo-likelihood is based on optimizing a product of one-dimensional conditional densities under a log loss.

$$f^{PL}(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} \sum_{d=1}^D P_e(\mathbf{x}) \log P_{\theta}(x_d | \mathbf{x}_{-d})$$

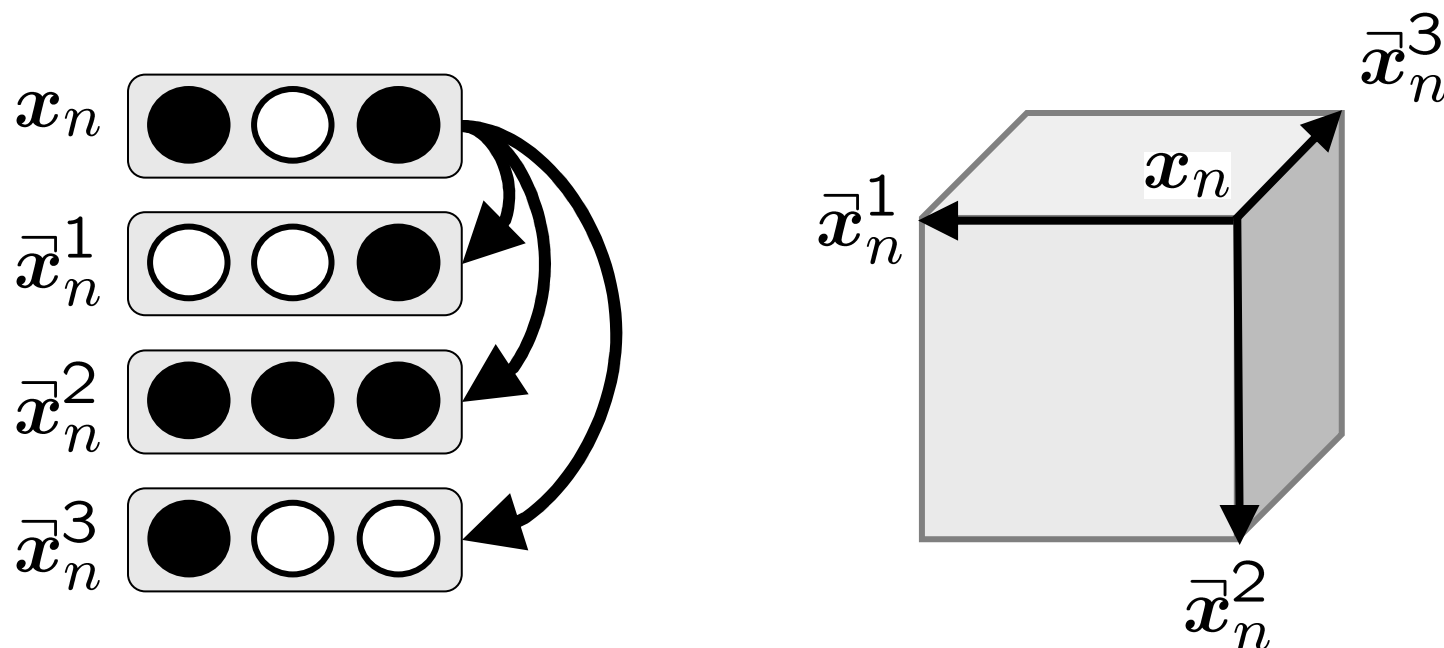
$$= \frac{1}{N} \sum_{n,d} g_{PL}(r_{dn})$$

$$g_{PL}(r) = -\log(1 + r^{-1})$$

$$r_{dn} = P_{\theta}(\mathbf{x}_n) / P_{\theta}(\bar{\mathbf{x}}_n^d)$$

Pseudo-Likelihood

- The principle of maximum pseudo-likelihood is based on optimizing a product of one-dimensional conditional densities under a log loss.



Ratio Matching

- The ratio matching principle is very similar to pseudo-likelihood, but is based on minimizing a squared difference between one dimensional conditional distributions.

$$\begin{aligned} f^{RM}(\theta) &= \sum_{\mathbf{x} \in \mathcal{X}} \sum_{d=1}^D \sum_{\xi \in \{0,1\}} P_e(\mathbf{x}) \left(P_\theta(X_d = \xi | \mathbf{x}_{-d}) - P_e(X_d = \xi | \mathbf{x}_{-d}) \right)^2 \\ &= \frac{1}{N} \sum_{n=1}^N \sum_{d=1}^D g_{RM}(r_{dn}) \\ g_{RM}(r) &= (1 + r)^{-2} \end{aligned}$$

Aapo Hyvarinen. Some extensions of score matching. Computational Statistics & Data Analysis, 51(5):2499–2512, 2007.

Generalized Score Matching

- The generalized score matching principle is similar to ratio matching, except that the difference between inverse one dimensional conditional distributions is minimized.

$$\begin{aligned} f^{GSM}(\theta) &= \sum_{\mathbf{x} \in \mathcal{X}} \sum_{d=1}^D P_e(\mathbf{x}) \left(\frac{1}{P_\theta(x_d | \mathbf{x}_{-d})} - \frac{1}{P_e(x_d | \mathbf{x}_{-d})} \right)^2 \\ &= \frac{1}{N} \sum_{n=1}^N \sum_{d=1}^D g_{GSM}(r_{dn}) \\ g_{GSM}(r) &= r^{-2} - 2r \end{aligned}$$

Siwei Lyu. Interpretation and generalization of score matching. In *Uncertainty in Artificial Intelligence 25*, 2009.

Minimum Probability Flow

- Minimize the flow of probability from data states to non-data states (as we've just seen!).

$$\begin{aligned} f^{MPF}(\theta) &= \sum_{\mathbf{x} \in \mathcal{X}} \sum_{d=1}^D P_e(\mathbf{x}) \log P_{\theta}^{(\epsilon)}(\mathbf{x}) \\ &\approx \frac{1}{N} \sum_{n=1}^N \sum_{d=1}^D I[P_e(\bar{\mathbf{x}}_n^d) = 0] g_{MPF}(r_{dn}) \\ g_{MPF}(r) &= r^{-1/2} \end{aligned}$$

[Jascha Sohl-Dickstein](#), [Peter Battaglino](#), [Michael R. DeWeese](#). Minimum Probability Flow Learning.

Comparison: Gradients

$$\nabla f^{ML} \approx - \left(\frac{1}{N} \sum_{n=1}^N \nabla F_{\theta}(\mathbf{x}_n) - \frac{1}{S} \sum_{s=1}^S \nabla F_{\theta}(\tilde{\mathbf{x}}_s) \right)$$

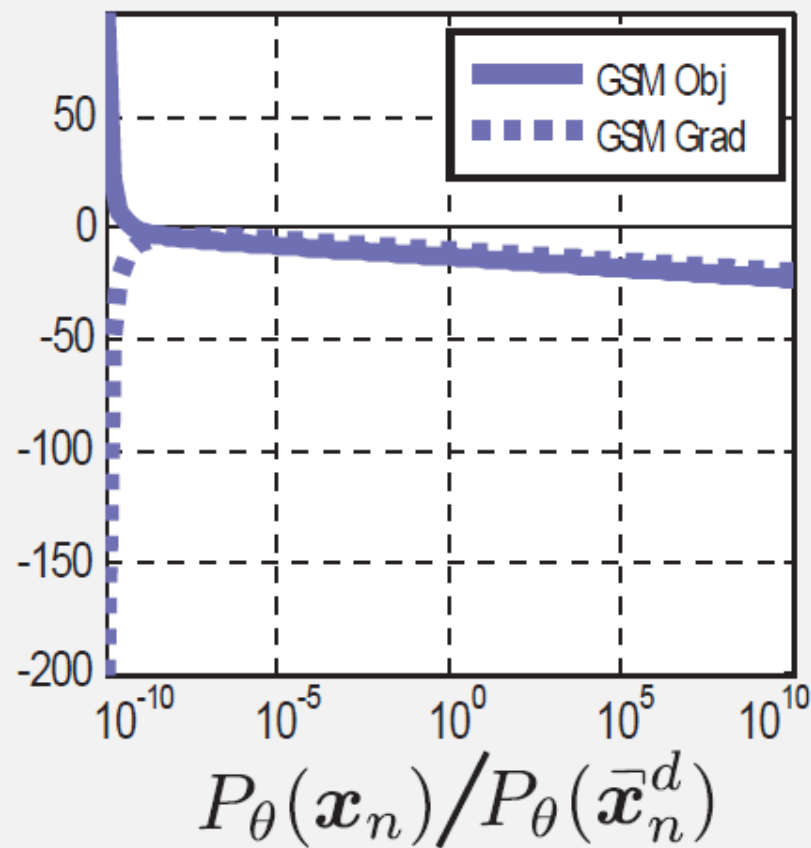
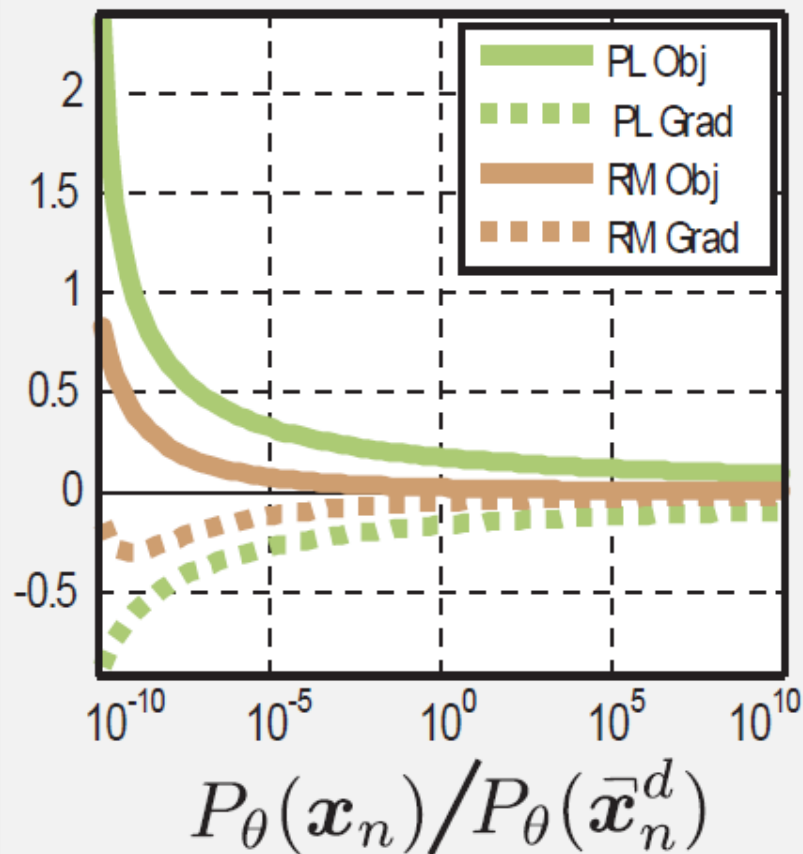
$$\nabla f^{CD} \approx - \frac{1}{N} \sum_{n=1}^N (\nabla F_{\theta}(\mathbf{x}_n) - \nabla F_{\theta}(\tilde{\mathbf{x}}_n))$$

$$\nabla f^{PL} = \frac{-1}{N} \sum_{n,d} g'_{PL}(r_{dn}) r_{dn} (\nabla F_{\theta}(\mathbf{x}_n) - \nabla F_{\theta}(\bar{\mathbf{x}}_n^d))$$

$$\nabla f^{RM} = \frac{2}{N} \sum_{n=1}^N \sum_{d=1}^D g'_{RM}(r_{dn}) r_{dn} (\nabla F_{\theta}(\mathbf{x}_n) - \nabla F_{\theta}(\bar{\mathbf{x}}_n^d))$$

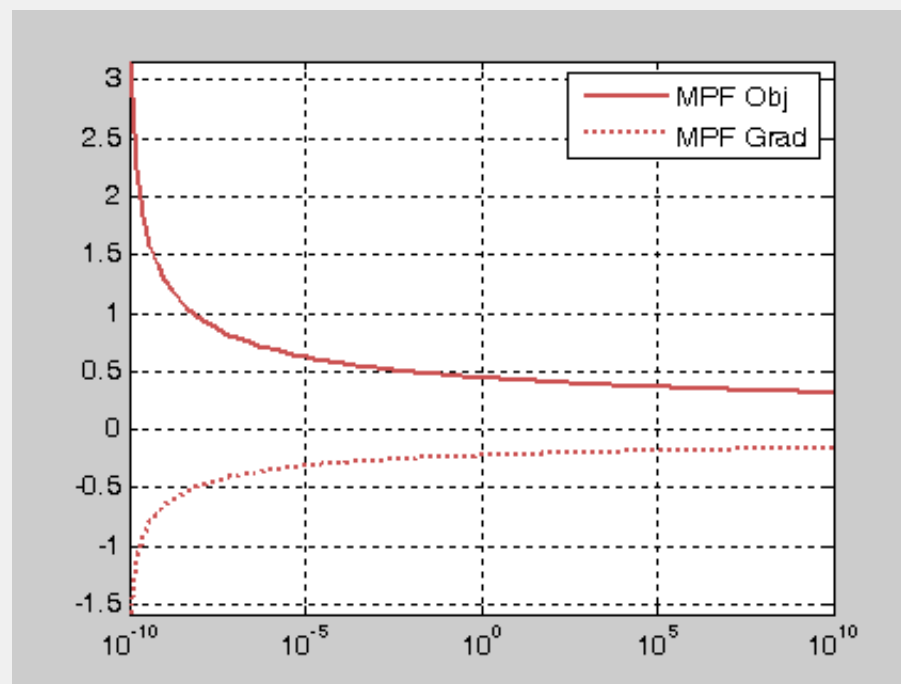
$$\nabla f^{GSM} = \frac{2}{N} \sum_{n=1}^N \sum_{d=1}^D g'_{GSM}(r_{dn}) r_{dn} (\nabla F_{\theta}(\mathbf{x}_n) - \nabla F_{\theta}(\bar{\mathbf{x}}_n^d))$$

Comparison: Weighting Functions



Comparison: Weighting Functions

What about MPF?



$$P_\theta(\mathbf{x}_n) / P_\theta(\bar{\mathbf{x}}_n^d)$$

Comparison: A Manifold of Estimators?

Dimensions:

1. Neighborhood structure around data configurations.
2. Form of loss function on the probability ratio.
 - Smooth
 - Monotonically decreasing
 - Bounded below
 - Others?

Covers: PL, GLP, NLCE, RM, GSM, MPF

Limitations: No good for missing data/explicit latent variables.

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Experiments:

Data Sets:

- MNIST handwritten digits
- 20 News Groups
- CalTech 101 Silhouettes

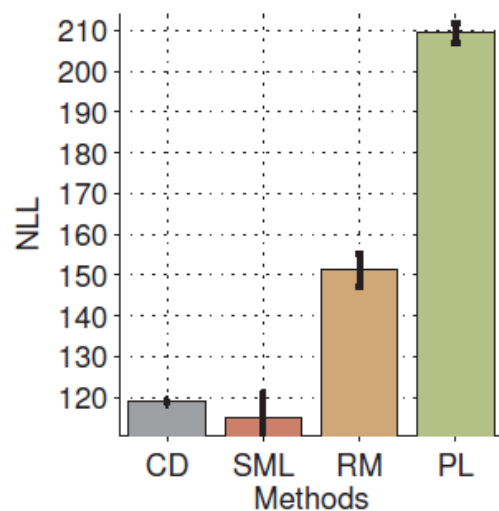
Evaluation Criteria:

- Log likelihood (using AIS estimator)
- Classification error
- Reconstruction error
- De-noising
- Novelty detection

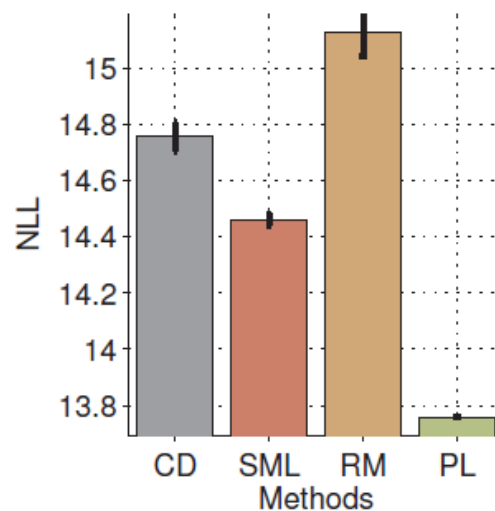
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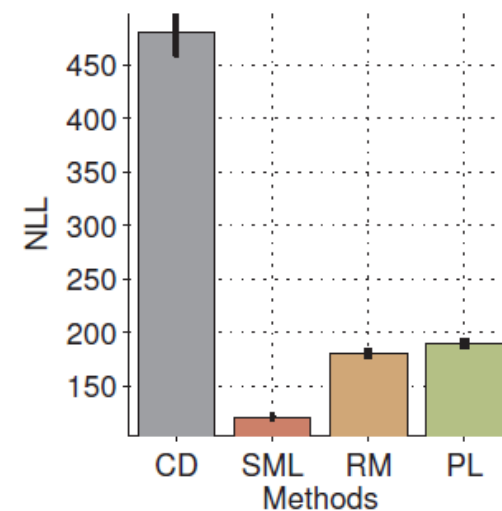
Experiments: Log Likelihood



(a) MNIST



(b) 20News

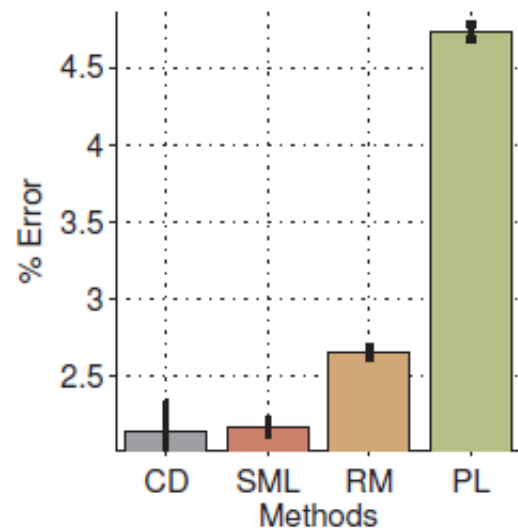


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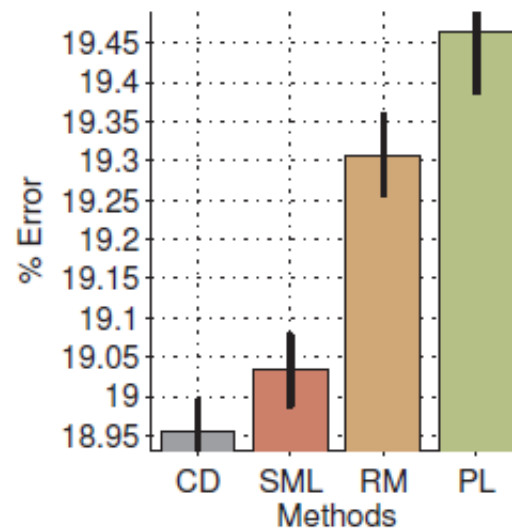
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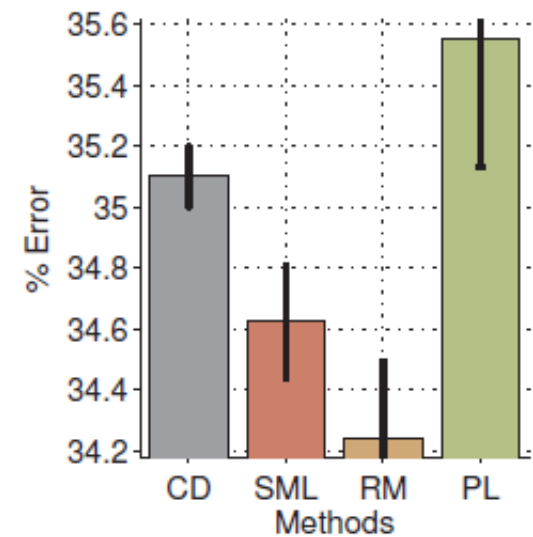
Experiments: Classification Error



(a) MNIST



(b) 20News

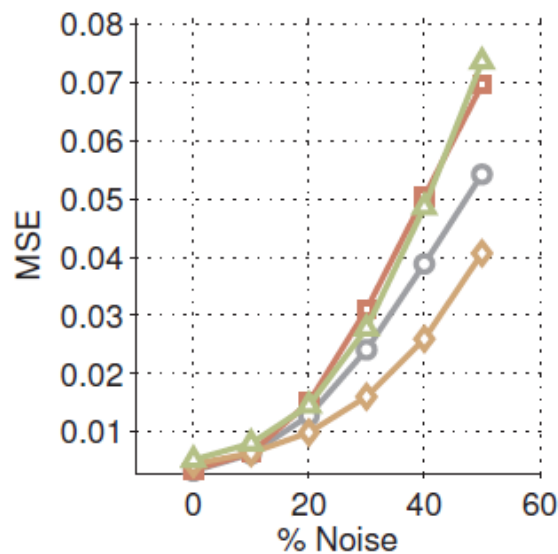


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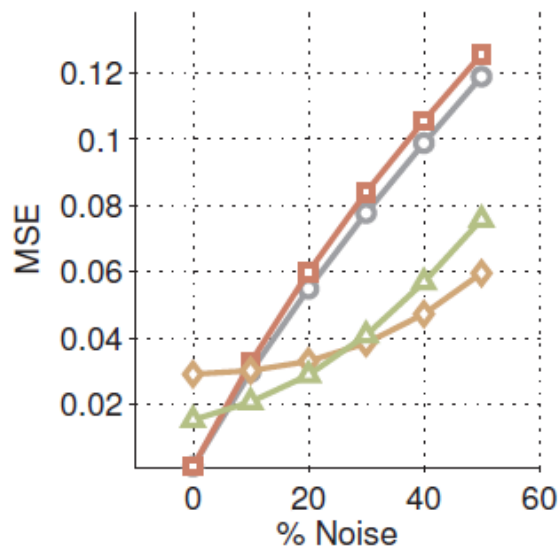
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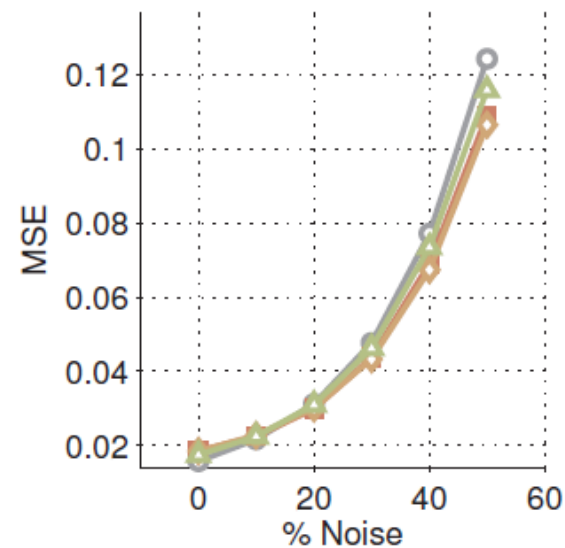
Experiments: De-noising



(a) MNIST



(b) 20News



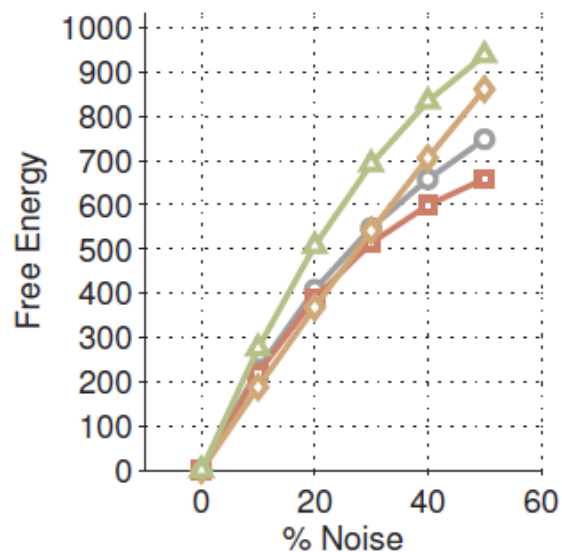
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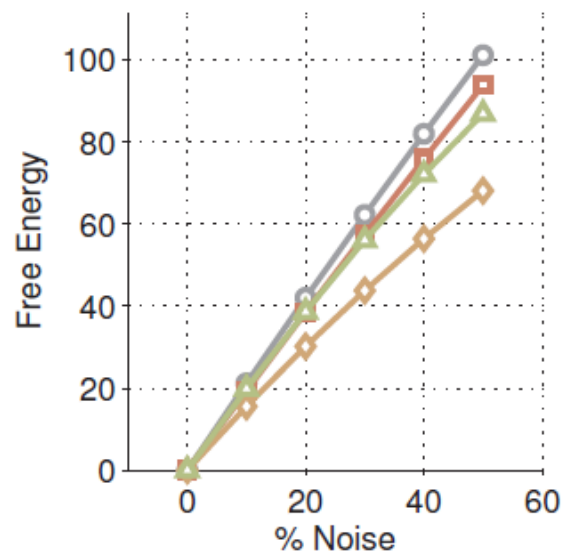
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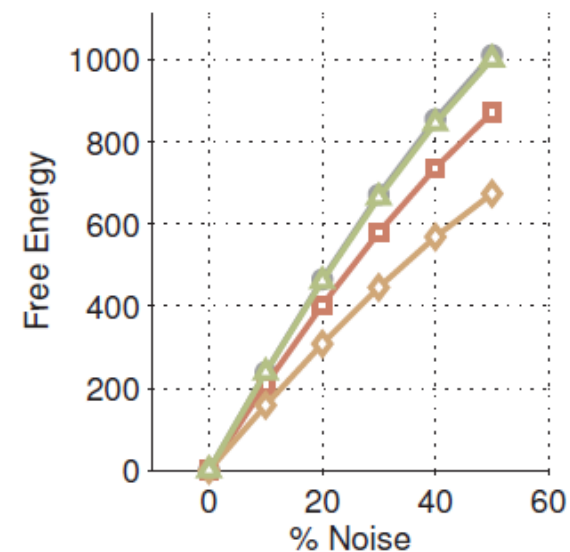
Experiments: Novelty Detection



(a) MNIST



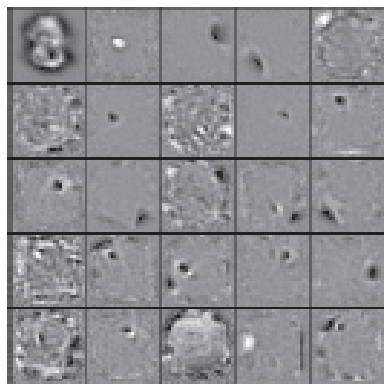
(b) 20News



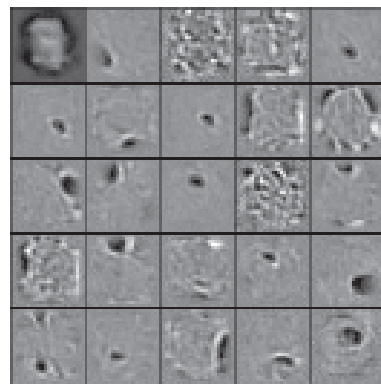
(c) CalTech



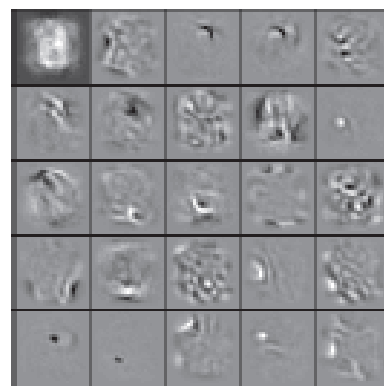
Experiments: Learned Weights on MNIST



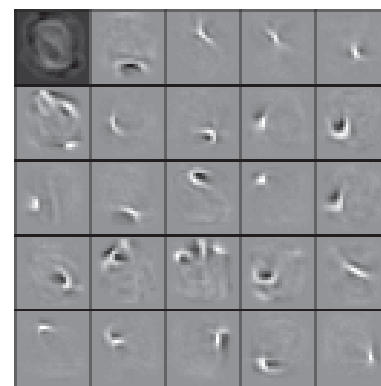
(a) CD



(b) SML



(c) PL



(d) RM